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Comparing Global Hydrological Models and Combining them with GRACE by Dynamic Model Data Averaging (DMDA)

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Abstract

Historically, hydrological models have been developed to represent land-atmosphere interactions by simulating water storage and water fluxes. These models, however, have their own unique characteristics (strength and weakness) in capturing different aspects of the water cycle, and their results are typically compared to or calibrated against in-situ observations such as river runoff measurements. As a result, there may be gross inaccuracies in the estimation of water storage states produced by these models. In this study, we present the novel approach of Dynamic Model Data Averaging (DMDA), which can be used to compare and merge multi-model water storage simulations with monthly Terrestrial Water Storage (TWS, a vertical summation of surface and sub-surface water storage) estimates from the Gravity Recovery And Climate Experiment (GRACE) satellite mission. Here, the main hypothesis is that merging GRACE data with multi-model outputs likely provides more skillful hydrological estimations compared to a single model or data set. Theoretically, the proposed DMDA combines the benefits of the Kalman Filter (KF) and Bayesian Model Averaging (BMA) techniques and has the capability to deal with various observations and models with different error structures. Based on the Bayes theory, DMDA provides time-variable weights for hydrological models to compute an average of their outputs that are best fitted to GRACE TWS estimates. Numerically, the DMDA method is implemented by integrating the output of six hydrological and land surface models (PCR-GLOBWB, SURFEX-TRIP, LISFLOOD, HBV-SIMREG, W3RA, and ORCHIDEE) and monthly GRACE TWS estimates (2002–2012) within the world’s 33 largest river basins, while considering the inherent uncertainties of

all inputs. Our results indicate that DMDA correctly separates GRACE TWS estimates into surface water, soil moisture and groundwater compartments. Linear trends fitted to the DMDA-derived groundwater compartment are found to be different from those of original models. This means that anthropogenic influences within the GRACE data, which are not well reflected by models, are introduced by DMDA. We also find that temporal correlation coefficients between the DMDA-derived individual water storage estimations (surface water, soil moisture, and groundwater) and the El Niño Southern Oscillation (ENSO) index are considerably increased compared to those derived between individual model simulations and ENSO (e.g., an increase from -0.2 to 0.6 in the Murray River Basin). For the Nile River Basin, they changed from 0.1 to 0.4 for the soil moisture, and from 0.3 to 0.7 for the surface water compartment. Comparisons between the DMDA-derived surface water and those from independent satellite altimetry observations indicate that after implementing DMDA, temporal correlation coefficients within major lakes are increased. Based on these results, we have gained confidence in the DMDA water storage estimates to be used for improving the characterization of water storage over broad regions of the globe.

Keywords: GRACE, Terrestrial Water Storage (TWS), Dynamic Model Data Averaging (DMDA), Kalman Filter (KF), Bayesian Model Averaging (BMA), Multi-Hydrological Models, Satellite Altimetry

1. Introduction

Studying global water storage changes and their relationships with climate variability and exploring their trends are important to understand the interactions between the Earth’s water, energy, and carbon cycles. It is also essential for managing water resources and understanding floods and food risks in a changing climate. In-situ and/or remote sensing observations provide estimates of different aspects of the Earth system, but they do not provide water cycle closure due to sampling and retrieval errors. In practice, hydrological models are used to quantify hydro-meteorological processes such as interactions between the global climate system and the water cycle ([Sheffield et al., 2012](#)), the contribution of land hydrology to global sea level rise ([Boening et al., 2012](#)), as well as to support applications related to water resources planning and management ([Hanington et al., 2017](#)). However, model simulations are prone to errors due to imperfect model structure, as well as errors in inputs and forcing data that are used to run model simulations. As a result, available models operating at regional to global scales have limited skills to reflect human

70 impacts on water storage and runoff changes (*Wada et al., 2012; Scanlon et al., 2018;*
71 *Singer et al., 2018*).

72 Among available remote sensing techniques, the Gravity Recovery And Climate
73 Experiment (GRACE, 2002–2017) satellite mission (*Tapley et al., 2004*) and its
74 Follow-On mission (GRACE-FO, 2018–onward) provide an opportunity to assess
75 the global water cycle by monitoring time-variable gravity fields. Global GRACE-
76 derived time-variable gravity field data can be used to estimate changes in Terrestrial
77 Water Storage (TWS), which is a vertical summation of canopy, surface water (lakes,
78 rivers, and wetlands), as well as soil moisture and groundwater storage. Changes in
79 TWS provide a critical measure of regional and global water balances, which cannot
80 be measured by any other satellite mission. A review of GRACE applications in
81 hydrology, and particularly for groundwater monitoring, can be found in *Frappart*
82 *and Ramillien (2018)*.

83 GRACE data can be used in conjunction with hydrological models to maximize
84 information gained from modelling with rationalisation and separation of GRACE
85 TWS. Thus, the gravimetric data from GRACE can inject realism into regional hy-
86 drological predictions, which are often poorly constrained in terms of TWS. Generally
87 speaking, integrating GRACE data with hydrological models is important from two
88 perspectives: (1) it can update (modify) water storage simulation within hydrologi-
89 cal models and (2) it vertically separates GRACE TWS into storage compartments.
90 The first point is of interest for hydrologists since most global models are not usually
91 combined with water storage observations (*Bai et al., 2018*). Therefore, such updates
92 may lead to more realistic water storage simulations, which makes these models more
93 useful for water resource applications (see e.g., *Werth et al., 2009; Mostafaie et al.,*
94 *2018*). Regarding the second point, it is important to state that any attempt to
95 vertically separate GRACE-derived TWS into its individual components requires a
96 priori information from other sources, such as, hydrological models, satellite altime-
97 try observations to estimate surface water storage, and soil moisture remote sensing
98 data to estimate shallow depth soil moisture storage changes (*Forootan et al., 2014*).

99 Various studies have developed techniques to merge multi-resources and achieve
100 vertical separation of surface and sub-surface water storage compartments by several
101 methods outlined below.

102 (a) Forward modeling techniques are used to evaluate different compartments of
103 mass variations through a simple reduction process, relying on model and/or observa-
104 tion data for other compartments, e.g., surface water and soil moisture, if groundwa-
105 ter should be estimated (e.g., *Tiwari et al., 2009; Rodell et al., 2009; Strassberg et al.,*
106 *2009; Feng et al., 2013; Khandu et al., 2016*). This method is relatively straightfor-
107 ward, but it is not necessarily the most accurate way to separate GRACE signals,

108 due to the reflection of modeling error and/or observation errors on the final estima-
 109 tion of mass changes. Also, the spatial and temporal resolution of the observations
 110 (from satellites or in-situ) and model outputs, as well as their signal content are not
 111 necessarily consistent (see the discussions in, e.g., [Forootan et al., 2014](#)). Most of
 112 these limitations are taken into account by the methods described in what follows.

113 (b) Statistical inversion techniques, which are formulated based on statistical sig-
 114 nal decomposition techniques, such as Principal Component Analysis (PCA, [Lorenz,](#)
 115 [1956](#)) and its alternatives, e.g., Independent Component Analysis (ICA, [Forootan](#)
 116 [and Kusche, 2012, 2013](#)), have been used in previous studies to separate GRACE
 117 TWS into individual water storage estimates. For example, [Schmeer et al. \(2012\)](#)
 118 used PCA to generate a priori information about mass changes from global ocean,
 119 atmosphere, and land hydrology models. Then, they applied a least squares tech-
 120 nique to use GRACE TWS to modify their priori estimates. A statistical inversion,
 121 which works based on both PCA and ICA, was proposed in [Forootan et al. \(2014,](#)
 122 [2017\)](#) and [Awange et al. \(2014\)](#) to separate GRACE TWS using auxiliary data of sur-
 123 face water from satellite altimetry and individual sub-surface water storage estimate
 124 from a land surface model (Global Land Data Assimilation System (GLDAS, [Rodell](#)
 125 [et al., 2004](#))). This inversion harmonizes the use of all available data sets within a
 126 single least squares framework. As a result, a more consistent mass estimate (than
 127 that of the forward modeling in (a)) for individual water storage components can be
 128 achieved.

129 (c) Data Assimilation (DA) as well as simultaneous Calibration/Data Assimila-
 130 tion (C/DA) have been used in recent years to merge GRACE data with hydrological
 131 model outputs or other types of observations. These techniques rely on the model
 132 equations to relate water and energy fluxes to water storage changes. Therefore,
 133 unlike the inversion approach (b), combining information from observations (e.g.,
 134 GRACE TWS estimates) and a model is performed in a physically justifiable way.
 135 DA or C/DA can potentially increase physical understanding of the model and im-
 136 prove the model states by decreasing the simulation errors. For example, DA is used
 137 in [Zaitchik et al. \(2008\)](#); [Giroto et al. \(2016, 2017\)](#); [Tian et al. \(2017\)](#); [Khaki et al.](#)
 138 [\(2018d,e\)](#), while C/DA is applied in [Schumacher et al. \(2016, 2018\)](#) to improve global
 139 models such as GLDAS ([Rodell et al., 2004](#)), World-Wide Water Resources Assess-
 140 ment (W3RA, [Van Dijk, 2010](#)), WaterGap Global Hydrological Model (WGHM, [Döll](#)
 141 [et al., 2003](#)), and NOAH Multi Parameterization Land Surface Model (NOAH-MP
 142 LSM, [Niu et al., 2011](#)). Most of the previous DA and C/DA are implemented region-
 143 ally (except [Van Dijk et al. \(2014\)](#); [Khaki et al. \(2017a, 2018a\)](#)) for example over the
 144 Mississippi River Basin ([Zaitchik et al., 2008](#); [Schumacher et al., 2016](#)), Bangladesh
 145 ([Khaki et al., 2018d](#)), the Middle East ([Khaki et al., 2018e](#)), and the Murray-Darling

146 River Basin (*Tian et al.*, 2017; *Schumacher et al.*, 2018). In addition, these studies
 147 rely on simulation from (only) one selected hydrological model, which could contain
 148 errors in the model structure such as biases in the model’s internal parameters and
 149 boundary conditions. In each of these studies, multiple realisations of the model-
 150 derived water storage simulations were generated by perturbing the input forcing
 151 data and/or model parameters. A sequential integration techniques such as the En-
 152 samble Kalman Filtering (EnKF, *Evensen*, 1994) or its extensions was then used to
 153 merge GRACE data with the (ensemble) outputs of a single model (e.g., *Schumacher*
 154 *et al.*, 2016, 2018; *Khaki et al.*, 2017b). *Van Dijk et al.* (2014) used EnKF to merge
 155 GRACE data with a priori data from models and other remote sensing techniques.
 156 Their study covered the period of 2003-2012 and focused on updating the individual
 157 water storage estimates rather than interpreting the water storage estimates in terms
 158 of trends or addressing the suitability of models used to perform the analyses.

159 (d) In recent years, Bayesian-based techniques have been used to combine differ-
 160 ent observations with models and update their outputs. For example, *Long et al.*
 161 (2017) applied the Bayesian Model Averaging (BMA, *Hsu et al.*, 2009) technique to
 162 average multiple GRACE TWS products and global hydrological models to analyse
 163 spatial and temporal variability of global TWS. However, their study did not as-
 164 sess the update of individual surface and sub-surface water storage estimates. *Sha*
 165 *et al.* (2018) used a model-data synthesis framework based on Bayesian Hierarchical
 166 Modelling (BHM, see e.g., *Banerjee et al.*, 2004) to use GRACE TWS estimates to
 167 update land surface deformations derived from Glacial Isostatic Adjustment (GIA)
 168 models. Their study did not, however, address global hydrological mass changes.

169 It is worth mentioning here that the Ensemble Kalman Filter used for DA and
 170 C/DA can also be classified as a Bayesian-based technique because the cost function
 171 for updating unknown state parameters condition on the measurement data, is for-
 172 mulated based on the Bayes theory (see e.g., *Evensen*, 2003; *Schumacher*, 2016; *Fang*
 173 *et al.*, 2018). Methods, such as Particle Filter (PF) and Particle Smoother (PS) are
 174 also Bayesian (*Särkkä*, 2013), and have already been applied in a wide range of geo-
 175 physical and hydrological applications. For example, *Weerts and El Serafy* (2006)
 176 compared the capability of EnKF and PF to update a conceptual rainfall-runoff
 177 model using discharge and rainfall data. *Plaza Guingla et al.* (2013) also used the
 178 standard PF to assimilate a densely sampled discharge records into a conceptual
 179 rainfall-runoff model. However, *Bain and Crisan* (2008) and *Del Moral and Miclo*
 180 (2000) show that the rate of convergence of the approximate probability distribu-
 181 tion until attainment of the true posterior is inversely proportional to the number
 182 of particles used in the filter. This means that the filter perfectly approximates the
 183 posterior distribution when the number of particles tends to infinity. However, since

the computational cost of PF grows with the number of particles, choosing a specific number of particles in the design of filters is a key parameter for these methods. The rationale for introducing a new Bayesian data-model merging algorithm in this study is described in (e).

(e) In this study, we present the Dynamic Model Data Averaging method (DMDA, i.e., a modified version of Dynamic Model Averaging (DMA) approach presented by [Raftery et al., 2010](#)) to merge multi-model derived water storage simulations with GRACE TWS estimates, as an alternative technique to that described in (d). Our main goal is to evaluate available model outputs against GRACE TWS and merge them in a sensible way to gain more realistic insights about global surface and sub-surface water storage changes. The main hypothesis behind the presented approach is that each global hydrological model has its own unique characteristics and strengths in capturing different aspects of the water cycle. Therefore, relying on a single model often leads to predictions that represent some phenomena or events well at the expenses of others. [Scanlon et al. \(2018\)](#) recently compared GRACE TWS with the outputs of global models, whose results indicated inconsistencies in long-term trends and cyclic (e.g., seasonal) components. Besides, many studies have concluded that effective combination of multiple models may provide more skillful hydrological simulations compared to a single model ([Duan et al., 2007](#)). Therefore, a multi-model choice is considered in this study.

Our motivation to formulate the DMDA is based on its capability to deal with various observations and models with different structures. In summary, DMDA is based on the Bayes theory and provides time-variable weights to compute an average of hydrological model outputs, yielding the best fit to GRACE TWS estimates, while considering their errors (see section 3). These time-variable weights indicate which of the available models at a given point in time fits better to GRACE TWS estimates. These weights can then be used to separate the components of TWS and modify the estimation of water storage in these individual components. Therefore, the DMDA-derived ensemble is expected to yield more skillful (realistic) hydrological simulations compared to any individual model (see similar arguments in [Duan et al., 2007](#)). Here, we promote the use of DMDA over the previously introduced EnKF, PF, and PS methods because it is computationally more efficient in handling large dimensional problems such as the global integration implemented in this study. In addition, the DMDA’s time-variable weights can be used to assess the performance of hydrological models, whereas this aspect is missing in other merging techniques. More details about the computational aspects of DMDA are provided in section 3.

To implement the DMDA method, surface and sub-surface water storage simulations of the six published global hydrological and land surface models ([Schellekens](#)

222 *et al.*, 2017) are used. These models are structurally different but they are all forced
 223 by the same reanalysis data set (WATCH-Forcing-Data-ERA-Interim, WFDEI *Wee-*
 224 *don et al.*, 2014) as inputs. GRACE-derived TWS estimates are then used in the
 225 DMDA method to compare their outputs and merge them. A challenging problem in
 226 merging GRACE TWS with the outputs from multiple hydrological models is related
 227 to their different spatial and temporal resolutions. To overcome the computational
 228 problem caused by the spatial and temporal mismatch, *Schumacher et al.* (2016)
 229 introduced spatial and temporal matching functions, which are able to avoid compu-
 230 tational problems. In this study, we did not implement the spatial/temporal operator
 231 because both model outputs and GRACE data were set at monthly (temporal) and
 232 basin-averaged (spatial). Handling the differences in spectral domain is described
 233 in section 2.2. A realistic synthetic example is presented in section 4.1 to test the
 234 performance of the DMDA method, where the true merged values are known and the
 235 method can be evaluated to provide the confidence that it can be applied to a real
 236 case study. Our numerical results cover the world’s 33 largest river basins (see Figure
 237 ESM.1 in Electronic Supporting Material, ESM) for the period of 2002–2012, during
 238 which both GRACE data and model simulations are available. Global hydrological
 239 model outputs are compared against GRACE TWS, using DMDA-derived temporal
 240 weights, within the largest river basins for the period of this study (see section 4.2).
 241 The DMDA-derived updates, which are assigned to the long-term trend of surface
 242 and sub-surface water storage components, are explored and interpreted (see section
 243 4.3).

244 Among many climatic factors that influence inter-annual to decadal TWS changes,
 245 the El Niño Southern Oscillation (ENSO, *Barnston and Livezey*, 1987) events rep-
 246 resent a dominant impact on global precipitation and TWS changes (see, e.g., *Hurk-*
 247 *mans et al.*, 2009; *Chen et al.*, 2010; *Zhang et al.*, 2015; *Forootan et al.*, 2016; *Ni et al.*,
 248 2018; *Anyah et al.*, 2018; *Forootan et al.*, 2019). In this study, temporal correlation
 249 coefficients between model-derived storage outputs and the ENSO index are used as
 250 a measure to determine whether implementing the DMDA helps to derive realistic
 251 storage simulations (see section 4.3.1). In addition, independent surface water level
 252 observations from satellite altimetry within 14 major lakes, located in different river
 253 basins around the world, are used to validate our results (see section 4.4). This paper
 254 contains an Electronic Supporting Material (ESM) document that provide auxiliary
 255 information to improve understanding of the performed investigations.

256 2. Data sources

257 The data used in this paper include the monthly GRACE data to compute Terres-
 258 trial Water Storage (TWS) and individual water storage estimates from global models

259 to provide a priori estimates to perform a Bayesian signal separation. GRACE TWS
260 estimates are used in the DMDA to modify the multi-model water storage outputs.

261 2.1. GRACE Data

262 The latest release of the monthly GRACE level-2 (L2) product (RL06), expressed
263 as dimensionless spherical harmonic coefficients up to degree and order 90, are down-
264 loaded for the period of April 2002 to December 2012 from the Center for Space Re-
265 search (CSR, <http://www2.csr.utexas.edu/grace/RL06.html>). A limited length
266 of the GRACE data is used here since the global hydrological model outputs of
267 *Schellekens et al. (2017)* were available until 2012.

268 Recommended corrections are applied to generate monthly TWS fields from the
269 GRACE product, i.e., degree 1 coefficients are replaced by those from *Swenson et al.*
270 (2008) to account for the movement of the Earth’s center of mass. The zonal degree
271 2 spherical harmonic coefficients (C20) are replaced by more stable ones derived from
272 Satellite Laser Ranging (SLR) data (*Chen et al., 2007*). Surface deformations known
273 as the Glacial Isostatic Adjustment (GIA) are reduced using the output of the model
274 provided by *Wahr and Zhong (2012)*. GRACE level-2’s correlated errors are reduced
275 by applying the DDK2 an-isotropic de-correlation filter (*Kusche et al., 2009*). The
276 application of smoothing filters causes a spatial leakage problem, which is evaluated
277 in terms of TWS errors following the approach in *Wahr et al. (1998)*; *Khaki et al.*
278 (2018c) over the world’s 33 largest river basins as shown in Fig. ESM.1. An overview
279 of the TWS’s strength and our error estimates is shown in ESM-section 2 (see Figure
280 ESM.2).

281 2.2. Global Hydrological Model (GHM) Outputs

282 Monthly water balance components from six large-scale Global Hydrological Mod-
283 els (GHMs) including PCR-GLOBWB (*Van Beek et al., 2011*; *Wada et al., 2014*),
284 SURFEX-TRIP (*Decharme et al., 2013*), LISFLOOD (*Van Der Knijff et al., 2010*),
285 HBV-SIMREG (*Lindström et al., 1997*), W3RA (*Van Dijk, 2010*), and ORCHIDEE
286 (*Polcher et al., 2011*) are used in this study to provide a priori information about
287 groundwater, soil moisture, surface water, canopy, and snow water storage com-
288 ponents. The output of these models are published by the earth2Observe Tier-1
289 (*Schellekens et al., 2017*), and are available at 0.5° spatial resolution covering the pe-
290 riod of 1979–2012 which can be downloaded from <http://earth2observe.github.io/water-resource-reanalysis-v1>.
291

292 Although, these models are structurally different, i.e., they use different method-
293 ology to simulate water changes, they are driven by the same reanalysis-based forcing
294 data set, WFDEI (WATCH Forcing Data methodology applied to ERA-Interim re-
295 analysis *Weedon et al., 2014*). In other words, all hydrological models that are used

in this study may represent the TWS, but their respective approaches for simulating TWS and its corresponding storage compartments are not identical. For example, [Schellekens et al. \(2017\)](#) state that PCR-GLOBWB and SURFEX-TRIP contain all surface and sub-surface water storage components in their TWS estimation. In contrast, TWS derived from LISFLOOD, HBV-SIMREG, and W3RA are equal to the summation of groundwater, soil moisture, and snow, while that of ORCHIDEE is the summation of soil moisture, surface water, and snow storage components.

An overview of the model outputs used in this study is provided in Table 1, and the linear trend (as a representative of monotonic long-term storage changes) fitted to the model outputs are shown in [ESM-section 3](#).

TABLE 1

To ensure that the TWS estimates from GRACE L2 data and model outputs have the same spectral content, 0.5° resolution hydrological model outputs are transformed into the spectral domain and truncated to the maximum degree and order 90. The conversion follows an ordinary integration while considering the Gibbs effect along the coast lines (for more details please see, e.g., [Wang et al., 2006](#); [Forootan et al., 2013](#)). Basin averages of each model components and their errors in terms of water storage are obtained from the same procedure used to process GRACE L2 data, i.e., implemented here following [Wahr et al. \(1998\)](#); [Khaki et al. \(2018c\)](#).

2.3. El Niño Southern Oscillation (ENSO) Index

The El Niño Southern Oscillation (ENSO, [Barnston and Livezey, 1987](#)) is a large-scale inter-annual climate variability phenomenon in the Tropical Pacific Ocean, which affects the climate of many regions of the Earth due to its ability to change the global atmospheric circulation, which influences temperature and precipitation across the globe ([Trenberth, 1990](#); [Forootan et al., 2016](#)). The positive phase on ENSO is known as El Niño, and its opposite phase is known as La Nina. The ENSO index used in this study is derived from sea surface temperature in the Niño 3.4 region ($5^\circ N - 5^\circ S, 170^\circ E - 120^\circ W$). Monthly ENSO index (Niño 3.4 index), which is provided by the NOAA National Center for Environmental Information (NCEI) covering 1948 onward, is downloaded from <https://www.esrl.noaa.gov/psd/data/correlation/nina34.data>. This index will be used later in this study to demonstrate whether the DMDA-derived surface and sub-surface water storage estimates are closer to the reality than those from individual models.

2.4. Satellite Altimetry of Major Lakes

Water level measurement by satellite altimetry has been developed and optimised for open oceans, yet improved post-processing techniques can be used to obtain reli-

able satellite altimetry-derived height measurements within inland water bodies such as lakes, rivers, floodplains and wetlands (e.g., *Moore and Williams, 2014; Uebbing et al., 2015*). In this study, satellite altimetry-derived surface water observations are used to validate TWS changes of GRACE and models as well as surface water derived from GHMs and the DMDA method. Satellite altimetry time series of major global lakes are available from the U.S. Department of Agriculture (USDA) (<https://ipad.fas.usda.gov/>). Repeated observations of the TOPEX/Poseidon (T/P), Jason-1, and Jason2/OSTM altimetry missions are included in this database. USDA provides time series of lake water level variations from 1992 to the present-day within 81 lakes, and from 2008 to present-day within more than 280 lakes around the world. An assessment over 14 lakes located within 8 river basins of this study is presented in section 4.4 for the period of 2002–2012. Details of these lakes are reported in Table 2.

TABLE 2

3. Dynamic Model Data Averaging (DMDA) Method

In this section, we present the mathematical formulation of Dynamic Model Data Averaging (DMDA), which follows the method of Dynamic Model Averaging (DMA, *Raftery et al., 2010*) but with some modifications to achieve a recursive update of hydrological model outputs using GRACE TWS data (Fig. 1 summarises the DMDA method). It will also be shown that the implementation of DMDA combines the benefits of state-space merging techniques, such as Kalman Filtering (KF, *Evensen, 1994*) or Particle Filtering (PF, *Gordon et al., 1993*), Markov Chain (MC, *Metropolis et al., 1953; Chan and Geyer, 1994; Kuczera and Parent, 1998*), and Bayesian Model Averaging (BMA, *Hsu et al., 2009*). DMDA can be applied in data assimilation applications that work with only one model, e.g., (*Giroto et al., 2016; Khaki et al., 2017c,b; Schumacher et al., 2018*), as well as in handling multi-model outputs as in *Van Dijk et al. (2014)*.

DMDA is formulated based on the representation of a state-space equation, which dynamically relates the GRACE TWS estimates and hydrological model outputs as:

$$y_t = z_t \theta_t + \epsilon_t, \quad (1)$$

$$\theta_t = \theta_{t-1} + \delta_t, \quad (2)$$

Equation (1) is known as ‘observation equation’ and represents a linear regression between the observation y_t (GRACE TWS estimates) and the vector of predictors

361 z_t (model-derived water storage simulations). The unknown regression parameter
 362 θ_t , commonly known as the ‘state vector’ ([Bernstein, 2005](#)), is allowed to evolve in
 363 time, according to equation (2), and is known as the ‘state equation’. In equations
 364 (1) and (2), ϵ_t and δ_t can be interpreted as the residual of output vector and state
 365 parameters, respectively. They are usually defined using a normal distribution with
 366 the mean value of zero and a standard deviation, which will be computed during the
 367 DMDA procedure.

368 It is worth mentioning here that the EnKF ([Evensen, 1994](#)) and PF are among
 369 popular algorithms that can be used to recursively update an estimate of the model
 370 states and produce corresponding innovation values given a sequence of observations
 371 in the state-space equation (similar to what introduced above). In theory, EnKF
 372 accomplishes this goal by linear projections, and the estimations in PF are performed
 373 through a Sequential Monte Carlo sampling. Comparing EnKF and PF, the latter
 374 includes a random element so it converges to the true posterior probability function
 375 if the number of samples is very large. While the strength of PF is in its ability to
 376 account for both Gaussian and non-Gaussian error distributions, it suffers from the
 377 curse of dimensionality, which means that the sample size increases exponentially
 378 with the dimension of the state-space in order to achieve a certain performance.
 379 This fact precludes the use of PF in high-dimensional data-model fusion problems
 380 ([Bengtsson et al., 2008](#); [Daum and Huang, 2003](#); [Snyder et al., 2008](#)). For linear and
 381 Gaussian-type state-space models, as presented in this study, the PF method will
 382 yield the same likelihood as EnKF when the number of simulations is large enough
 383 (this has been tested but the results are not shown to keep the focus of this study on
 384 presenting the DMDA). Therefore, the DMDA, which combines the benefits of the
 385 EnKF and it is mathematically rigorous like PF, is adopted for the global data-model
 386 integration of this study.

387 Equations (1) and (2) are formulated with the main assumption that there is little
 388 physical knowledge about how the defined regression model and its parameters are
 389 likely to evolve in time. However, we will show that, by introducing two parameters
 390 of λ and α , which are referred to as ‘forgetting factors’, one can control the temporal
 391 dependency of the DMDA solutions. These two parameters provide the opportunity
 392 to treat model simulations and observations of each step temporally dependent on,
 393 or independent from, previous steps. Since changes in water storage depend on the
 394 history of hydrological processes, accounting for temporal dependency between water
 395 states sounds logical.

396 Formulating DMDA to Update Multi-Model Outputs using GRACE TWS

397 Here the DMDA method is formulated to update the outputs of multi-hydrological
 398 models, M_k , (for six models: $k = 1, \dots, 6$). It is worth mentioning that since available
 399 models have different storage definitions, the length of the state vector can change
 400 from one model to another. Additionally, the structure of each individual storage
 401 components can also be defined differently in different models (e.g., the number of soil
 402 layers does not remain constant in different hydrological models). These differences
 403 can be handled by DMDA.

404 In the following, $Y_t = [y_1, \dots, y_t]$ represents the vector of observations (i.e., GRACE
 405 TWS estimates in our study) up to the time step t . To use this vector to update the
 406 water storage simulation of a single-model, one can estimate the unknown (linear)
 407 regression parameters (θ_t) as

$$\theta_{t-1}|Y_{t-1} \sim N(\hat{\theta}_{t-1}, \hat{\Sigma}_{t-1}). \quad (3)$$

408 The distribution of each parameter can be assumed to be normal with unknown
 409 mean $\hat{\theta}_{t-1}$ and the variance $\hat{\Sigma}_{t-1}$. The regression coefficients at time t (θ_t) can then
 410 be obtained using θ_{t-1} from equation (3) and by introducing $\delta_t \sim \mathcal{N}(0, W_t)$ to the
 411 state equation (equation (2)). Therefore, the desired parameters at time t are defined
 412 by

$$\theta_t|Y_{t-1} \sim N(\hat{\theta}_{t-1}, R_t), \quad (4)$$

413 where

$$R_t = \hat{\Sigma}_{t-1} + W_t. \quad (5)$$

414 In equation (5), W_t is the covariance matrix of the state innovation vector (δ_t
 415 in equation (2)) and it shows the dependency of the regression parameters at each
 416 time point to the previous time. However, in practice, there is no information about
 417 the temporal relationship between GRACE TWS estimates and hydrological model
 418 outputs to be used to define W_t . Therefore, to mathematically define a temporal
 419 dependency, R_t in equation (4) can be replaced by

$$R_t = \lambda^{-1} \hat{\Sigma}_{t-1}, \quad (6)$$

420 where λ ($0 < \lambda \leq 1$) controls the influence of previous observations on the regression
 421 value at time t , and is known as ‘forgetting factor’ in the DMDA method (see, e.g.,
 422 [Fagin, 1964](#); [Jazwinski, 2007](#)).

423 [Hannan et al. \(1989\)](#) indicated that in the recursive estimation of auto-regressive
 424 models, the covariance of previous steps is derived as a weighted product of the

current step (i.e., weighted by λ^{-1} in equation (6)). By this assumption, the effective window size of temporal dependency is estimated by $1/(1 - \lambda)$. In our case, we choose λ to be 0.95, which means that for monthly data, the effective window size is equivalent to 18 months. This value is chosen experimentally because it minimized the Root Mean Square (RMS) of differences between TWS derived from DMDA and GRACE.

To apply DMDA and update water storage simulated by K different models, the parameter prediction of equation (4) is extended as

$$\theta_t^{(k)} | M_t = k, Y_{t-1} \sim N(\hat{\theta}_{t-1}^{(k)}, \lambda^{-1} \hat{\Sigma}_{t-1}^{(k)}), \quad k = 1, \dots, K, \quad (7)$$

where $M_t = k$ denotes which model (from the $k = 1, 2, \dots, K$ available models) applies at time t , and the solution $\theta_t^{(k)}$ and $\hat{\Sigma}_{t-1}^{(k)}$ can be obtained using a Kalman Filter (KF)-type update conditional on $M_t = k$ for each sample. This (KF-type) update at time t is derived as

$$\theta_t^{(k)} | Y_t \sim N(\hat{\theta}_t^{(k)}, \hat{\Sigma}_t^{(k)}). \quad (8)$$

Regression parameters to update multi-model storage simulations can be estimated as

$$\hat{\theta}_t^{(k)} = \hat{\theta}_{t-1}^{(k)} + R_t^{(k)} z_t^{(k)} (V_t + z_t^{(k)} (R_t^{(k)} + Q_t^{(k)}) z_t^{(k)T})^{-1} (y_t^{(k)} - z_t^{(k)} \hat{\theta}_{t-1}^{(k)}), \quad (9)$$

where V_t is the covariance matrix of GRACE TWS estimates (our observation), and Q_t is the covariance matrix of predictor z_t (see equation (1)). In this study, the leakage errors of model-derived TWS are estimated for the world's 33 river basins (similar to those of GRACE). These errors are used to generate Q_t , which is therefore a diagonal matrix in the DMDA implementation of this study. For a grid based implementation of DMDA, one can use the full covariance matrix of GRACE TWS similar to [Schumacher et al. \(2016\)](#). The covariance matrix $\hat{\Sigma}_t$ in equation (8) can be estimated from

$$\hat{\Sigma}_t^{(k)} = R_t^{(k)} - R_t^{(k)} z_t^{(k)T} (V_t + z_t^{(k)} (R_t^{(k)} + Q_t^{(k)}) z_t^{(k)T})^{-1} z_t^{(k)} R_t^{(k)}. \quad (10)$$

It is evident from equations (9) and (10) that the estimation of regression parameter $\hat{\theta}_t$ is conditional on a particular model. Therefore, the DMDA solution to obtain unconditional results and update multi-model simulations involves calculating the posterior model probability $P(M_t = k | Y_t)$ as a weight for each model, which changes at each time step. In the following, we show that time-variable weights need to be computed for each model k by choosing a forgetting factor α in a recursive method,

where $k = 1, \dots, K$. These weights are then used to average the models, which leads to the best fit to the GRACE TWS estimates. This justifies the term ‘Dynamic’ in the DMDA and makes the method different from other averaging techniques such as the Bayesian Model Averaging (BMA).

Let us assume that $P(M_t = k|Y_t) = \pi_{t|t,k}$, then the posterior model probability for each model k at time t can be estimated as

$$\pi_{t|t,k} = \frac{\pi_{t|t-1,k} P(y_t|M_t = k, Y_{t-1})}{\sum_{l=1}^K \pi_{t|t-1,l} P(y_t|M_t = l, Y_{t-1})}, \quad (11)$$

where, $P(y_t|M_t = k, Y_{t-1})$ is the density of the observation at time t , conditional on model k , as well as $Y_{t-1} = [y_1, y_2, \dots, y_{t-1}]$, which is estimated by a normal distribution as

$$y_t|M_t = k, Y_{t-1} \sim N(z_t^{(k)} \hat{\theta}_{t-1}^{(k)}, V_t + z_t^{(k)} (R_t^{(k)} + Q_t^{(k)}) z_t^{(k)T}), \quad (12)$$

and, $\pi_{t|t-1,k}$ is the model prediction equation, which is defined by

$$\pi_{t|t-1,k} = \sum_{l=1}^K \pi_{t-1|t-1,k} a_{kl}. \quad (13)$$

In equation (12), $\hat{\theta}_{t-1}^{(k)}$ is estimated using the KF-type update as formulated in equations (9) and (10), while $R_t^{(k)}$ is obtained from equation (6) by choosing a forgetting factor λ , i.e., between 0 and 1.

In equation (13) $a_{kl} = P(M_t = l|M_{t-1} = k)$ is the element of the $K \times K$ transition matrix $A(a_{kl})$ between models, which can be onerous when the number of models is large, e.g., for K models and τ time steps, the number of combinations of models will be $K^{2\tau}$. In our study, we have 6 hydrological models, and 122 time steps over the entire period of the study (2002–2012), which leads to 6^{244} combinations of models. To specify the transition matrix A , one way is to use the Markov Chain Monte Carlo method (MCMC, [Geyer, 2011](#)), which will typically be computationally expensive. Therefore, in this study, we avoid the implicit specification of the transition matrix using the forgetting factor of $0 < \alpha < 1$, which has the same role as λ in equation (6). As a result, the model prediction equation (13) can be rewritten as

$$\pi_{t|t-1,k} = \frac{\pi_{t-1|t-1,k}^\alpha}{\sum_{l=1}^K \pi_{t-1|t-1,l}^\alpha}. \quad (14)$$

The posterior model probability, or weights, for each model at time t is estimated in a recursive solution between equations (11), (12), and (14). This process is initialized by setting $\pi_{0|0,k} = \frac{1}{K}$ for $k = 1, \dots, K$, and assigning a prior values to the initial

condition of the states $\theta_0^{(k)} \sim N(0, \Sigma_0^{(k)})$ and $\Sigma_0^{(k)} = \text{Variance}(y_t^{(k)}) / \text{Variance}(z_t^{(k)})$. The reason of choosing this prior value is that in a linear regression, a regression coefficient for a predictor z_t is likely to be less than the standard deviation of the observations y_t divided by the standard deviation of predictors z_t (for more information see e.g., [Raftery, 1993](#)). In our numerical evaluation of DMDA with six hydrological models, the optimum regression estimates are found when $0.85 < \alpha < 0.9$, because the RMS of differences between the DMDA-derived TWS and those of GRACE were at a minimum here. By choosing a forgetting factor $\alpha = 0.9$, we assume a temporal smoothing window with 36 month time steps between 6 hydrological model ensembles to predict posterior probability values of each model k at time t . It means that the contribution of hydrological models at time $t - 37$ in to the posterior model probability of each model k at time t is negligible. The length of this smoothing window is reduced e.g., to 8 months if we choose $\alpha = 0.2$.

The multi-model predictions of y_t is a weighted average of model specific prediction \hat{y}_t , using the posterior model probabilities, $\pi_{t|t,k} = \text{Pr}(M_t = k|Y_t)$, as its weights, i.e.,

$$\hat{y}_t^{DMDA} = \sum_{l=1}^K \pi_{t|t,l} \hat{y}_t^{(l)}, \quad (15)$$

where $\hat{y}_t^{(k)} = z_t^{(k)} \hat{\theta}_t^{(k)}$.

The posterior model probability for each model at time t , along with the estimated time-variable regression parameter $\theta_t^{(k)}$ from KF-type updating equation (9) are used to estimate the multi-model prediction of water storage components as

$$\hat{z}_{j,t}^{DMDA} = \sum_{l=1}^K \pi_{t|t,l} z_{j,t}^{(l)} \hat{\theta}_{j,t}^{(l)}, \quad (16)$$

where j represents each of the water storage components, i.e. groundwater, soil moisture, surface water, canopy, and snow. To update the water storage simulations of a single-model using the GRACE TWS estimates and the DMDA approach, K needs to be set to 1, and the prediction step is limited to the conditional estimation of the parameter $\theta_t^{(k)} | M_t^{(k)}$ using equation (9).

The posterior model probability can also be used to estimate unconditional probability distribution of regression parameters $\Theta_t = (\theta_t^{(1)}, \dots, \theta_t^{(K)})$ given by observation Y_t following

$$p(\Theta_t | Y_t) = \sum_{l=1}^K p(\theta_t^{(l)} | M_t = l, Y_t) P(M_t = l | Y_t), \quad (17)$$

507 where $p(\theta_t^{(k)}|M_t^{(k)}, Y_t)$ shows the conditional distribution of $\theta_t^{(k)}$ which is approxi-
 508 mated by a normal distribution as:

$$\theta_t^{(k)}|M_t^{(k)}, Y_t \sim N(\hat{\theta}_t^{(k)}, \hat{\Sigma}_t^{(k)}). \quad (18)$$

509 The DMDA approach can be recovered to a standard Bayesian Model Averaging
 510 (BMA, [Hoeting et al. \(1999\)](#)) when $\alpha = \lambda = 1$. Then the posterior model probability
 511 of model k is given by

$$P(M_t = k|Y_t) = \frac{p(Y_t|M_t = k)}{\sum_{l=1}^K p(Y_t|M_t = l)}, \quad (19)$$

512 where $p(Y_t|M_t = k)$ is the marginal likelihood, obtained by integrating the product of
 513 the likelihood, $P(Y_t|\theta^{(k)}, M_t = k)$, and the prior, $P(\theta^{(k)}|M_t = k)$, over the parameter
 514 space (see also [Hsu et al., 2009](#)). Figure 1 summarises the work-flow of the DMDA
 515 approach.

FIGURE 1

516 4. Results

517 4.1. Setup a Simulation to Test the Performance of DMDA

518 Before applying the DMDA method on real data, its performance is tested in a
 519 controlled synthetic simulation, where the results of the Bayesian update are known
 520 by definition. In the first step of our simulation, we aim to compare DMDA and BMA
 521 in terms of updating hydrological model outputs with respect to the observations (i.e.,
 522 GRACE TWS estimates in this study). In the second step, it will be shown that the
 523 DMDA-derived time-variable weights are the same as the expected values.

524 To make the synthetic study simple, we assumed that TWS is defined as the
 525 summation of just groundwater and soil moisture components. By this definition,
 526 the time series of groundwater and soil moisture of two hydrological models, i.e., here
 527 selected as LISFLOOD (M_1) and SURFEX-TRIP (M_2), are introduced as predictors
 528 to the DMDA, and TWS derived from a third model, here selected to be PCR-
 529 GLOBWB, is considered as the observation (here standing in for GRACE derived
 530 TWS). By this choice, after applying DMDA to merge M_1 and M_2 with simulated
 531 observed TWS, we expect that the updated (DMDA-derived) groundwater and soil
 532 moisture storage estimates will be fitted to those of simulated observation. Here, we
 533 selected results within the Niger River Basin (id:20 in Fig. [ESM.1](#)), covering the pe-
 534 riod of 2002–2012. Figure 2 (A) shows the PCR-GLOBWB TWS as our observation,

Fig. 2 (B) represents the time series of groundwater and soil moisture derived from M_1 (B1, B3, blue curves) and M_2 (B2, B4, green curves), while the expected value of DMDA-derived groundwater and soil moisture (simulated observation) are shown with the red color curves in these figures.

The magnitude of minimum (Min), maximum (Max) and the Root Mean Square (RMS) of the signal for all simulated data sets can be found in Table 3. The uncertainty of these data sets are computed following a least squares error propagation, while considering the leakage error of GRACE TWS in the Niger River Basin. It is worth mentioning that the final results of the simulation do not depend on the selection of models and the adopted simplification. The RMS of differences between the simulated TWS and two selected models (reported in Table 3) indicates that M_2 (RMS of $\Delta_{TWS} = 14.1$ mm) had a better agreement with the observations compared to M_1 (RMS of $\Delta_{TWS} = 18.6$ mm). Figure 2 (C1) shows the estimated weights for the first model (W_1 , Mean= 0.47) and second model (W_2 , Mean= 0.53) obtained using DMDA (equation (11)). These results show that the model which had a better agreement with observations gained higher weights.

To compare DMDA and BMA methods to average hydrological components, we apply both of these methods on simulated data sets. The final results are shown in Fig. 2 (D1: groundwater) and (D2: soil moisture). Groundwater, soil moisture, and consequently TWS derived from DMDA shows better agreement with the expected values in comparison to the BMA results. The RMS of errors for both methods are reported in Table 3, which indicates that although TWS derived from BMA follow the expected value (RMS of error= 8.4 mm), the obtained individual components from this method are not close to the simulated values (RMS of errors of 20.4 mm and 18.6 mm are found for groundwater and soil moisture, respectively). A considerable decrease in the differences between hydrological components and the expected values of DMDA shows that the method is suitable to update multi-model water storage estimates. Details of the numerical comparisons can be found in Table 3.

In the second step of our simulation, we use the weights of the first step (W_1 , W_2 , Fig. 2 (C1)) plus a temporal white noise with standard deviation of 0.02 m (equal to the standard deviation of GRACE TWS error within the Niger River Basin) to simulate GRACE like TWS estimates. Reconstructed weights after applying the DMDA for the second time, using the new synthetic TWS observations, are shown in Fig. 2 (C2). The correlation coefficient between W_1 and W_2 with their reconstructed values is found to be 0.73 and the RMS of the reconstruction's errors is found to be 0.18. This indicates that the DMDA-derived weights are close to reality and further motivates us to apply it on real data sets.

FIGURE 2

TABLE 3

4.2. DMDA Weights to Compare Global Hydrological Models

TWS derived from DMDA is a weighted average of selected models by estimating time varying weights based on the Bayes rule as in equation (15). Figure 3 shows the estimated weights for ten basins with the largest RMS of differences between TWS derived from individual models and GRACE TWS. Time-variable weights derived from DMDA allow us (1) to quantify the quality and compare individual water storage simulations derived from each global hydrological model against GRACE TWS for different periods of time, and (2) to separate GRACE TWS in a Bayesian framework, while considering different model structures and errors within and between model simulations and GRACE data. The average of weights during 2002–2012 is considered as the basis to select the best model in DMDA results over 33 river basins which is shown in the middle of Fig. 3. From our numerical results, PCR-GLOBWB is found to gain the largest weights during this period, thus, it contributed the most in the DMDA-derived TWS in North Asia, Central Africa, and North America. The weights computed for SURFEX-TRIP are found to be larger than other models within the snow-dominated regions, such as, the Yukon and Mackenzie in the north part of America and the Lena in the Northeast Asia. Our results confirm the investigations by *Schellekens et al.* (2017), who compared the mentioned models against the Interactive Multi-sensor snow and Ice Mapping System (IMS, *Ramsay, 1998*). Apparently, multiple snow layers of SURFEX-TRIP helps it to better simulate snow dynamics during the cold seasons.

We also find that SURFEX-TRIP received the highest averaged weights (compared to other models) within the Amazon and Brahmaputra River Basins during 2002–2012. The explanation is that SURFEX-TRIP likely better accounts for (1) the snow coverage of the Brahmaputra River Basin, (2) the considerable contribution of surface water storage components in the TWS changes within the Amazon River Basin, and (3) the overall dry period within both basins (*Chen et al., 2009; Khandu et al., 2016*), specially the extreme hydrological droughts of 2005 and 2010 (*Forootan et al., 2019*). In the Amazon River Basin, we also find the highest performance for SURFEX-TRIP between 2009–2011. *Chen et al. (2009)* reported that in 2009 the Amazon River Basin experienced an extreme flood, which increased the magnitude of inter-annual TWS in this basin. TWS changes within the Amazon are also closely connected to the ENSO events in the tropical Pacific (*Kousky et al., 1984; Ropelewski*

606 *and Halpert, 1987*). Later we will show that surface water derived from SURFEX-
 607 TRIP shows the highest correlation with ENSO index in comparison with the other
 608 models of this study. This could be another reason that we derive the highest weights
 609 for SURFEX-TRIP between 2009-2011 within the Amazon River Basin.

610 Our results (Fig. 3) indicate that within the river basins with considerable irriga-
 611 tion (such as the Indus, Euphrates, and Orange River Basins), the relatively highest
 612 weights are assigned to the LISFLOOD and ORCHIDEE, where both account for
 613 human water-use (*Schellekens et al., 2017*). ORCHIDEE is also found to perform
 614 well within the Brahmaputra, Ganges, and Murray River Basins, each of which expe-
 615 rienced a strong decline in rainfall over the entire period of our study (e.g., 9.0 ± 4.0
 616 mm/decade between 1994–2014 over Ganges and Brahmaputra *Khandu et al., 2016*).
 617 Specifically, ORCHIDEE contains 14 soil layers (see Table 1) that help it to better
 618 resolve vertical water exchange within the irrigated regions. In *ESM-section 2*, it is
 619 shown that GRACE TWS changes within the Murray River Basin are considerably
 620 influenced by ENSO events (see also *Forootan et al., 2012, 2016*), and the simulated
 621 outputs of ORCHIDEE reflects these changes better than the other tested models
 622 justifying the higher weights that are assigned to this model within the DMDA pro-
 623 cedure. In *ESM-section 5*, we show that after applying the DMDA, model-derived
 624 TWS simulations are tuned to GRACE TWS.

FIGURE 3

625 4.3. DMDA-Derived Individual Water Storage Estimates

626 The estimated weights for the six models of section 4.2 along with the computed
 627 regression coefficients $\hat{\theta}_t$ (see the flowchart of Fig. 1), are used to compute the
 628 DMDA-derived groundwater, soil moisture, and surface water. In order to interpret
 629 the monotonic changes of water storage changes within the river basins, a long-term
 630 linear trend is fitted to the DMDA results that are shown in Figure 4, and the
 631 numerical values are reported in Table 4.

FIGURE 4

TABLE 4

632
 633 Figure 4 (a1) and (a2) show the linear trend fitted to the DMDA-derived ground-
 634 water and its uncertainty. The results indicate a decrease in groundwater in 42% of
 635 the assessed river basis (i.e., 14 of 33). The largest decreasing trends are found in
 636 basins with large-scale irrigation such as the Ganges (-14.77 ± 0.25 mm/yr), Indus
 637 (-8.26 ± 0.16 mm/yr) and Euphrates (-5.36 ± 0.23 mm/yr). The results confirm

the findings by [Khandu et al. \(2016\)](#), [Forootan et al. \(2019\)](#), and [Voss et al. \(2013\)](#), respectively. The strongest increasing trends in groundwater are seen in the Tocantins basin (South America) at the rate of 2.41 ± 0.47 mm/yr, the Okavango (South Africa) with a rate of 1.74 ± 1.31 mm/yr, and the Lena (Northeast Asia) with 1.74 ± 0.11 mm/yr. However, all of these trends are not statistically significant. The positive trends in groundwater storage in these last two basins are associated to the heavy rainfalls, seasonal floods and the geographical location of the Okavango Delta ([McCarthy et al., 1998](#)), and underground ice melting caused by global warming ([Dzhamalov et al., 2012](#)), respectively. Comparisons between the DMDA-derived groundwater and those of hydrological models indicate that after merging GRACE TWS with output from multiple hydrological models, the linear trend has changed considerably. This means that introducing GRACE data can successfully modify the anthropogenic effects, which are not well simulated by models (linear trends of the modelled groundwater are shown in [ESM-section 3](#)).

The linear trend fitted to the DMDA-derived soil moisture and its uncertainty are shown in Fig. 4 (b1) and (b2). We find strongest increasing trends in soil moisture estimates within the Murray (Australia), Okavango, and Orinoco (South America) River Basins with rates of 6.66 ± 0.15 , 3.92 ± 0.55 , and 3.45 ± 0.26 mm/yr respectively, and largest decreasing trends in the Brahmaputra and Euphrates with rates of -7.00 ± 0.69 and -5.75 ± 0.39 mm/yr.

Figure 4 (c1) and (c2) show the linear trends and their uncertainty fitted to the surface water storage estimated through the DMDA method. Linear trends of surface water within the 28 out of the 33 river basins are found to be statistically insignificant (values between -1 and +1 mm/yr). The strongest negative trends are found in the Euphrates, Murray, and Okavango River Basins with rates of -2.09 ± 0.09 , -1.47 ± 0.04 , and -1.42 ± 0.37 mm/yr respectively. In contrast, the largest positive trends are found within the Amazon and Colorado, at the rate of 1.43 ± 0.06 and 1.04 ± 0.04 mm/yr, respectively. The heavy flood during the summer of 2008–2009 ([Marengo et al., 2011](#); [Chen et al., 2010](#)), which was considerably bigger than the temporal mean, likely caused these positive trend in the Amazon River Basin. Negative trends in all three water storage compartments of the Euphrates River Basin (groundwater -5.36 ± 0.23 mm/yr, soil moisture -5.75 ± 0.39 mm/yr, and surface water -2.09 ± 0.09 mm/yr) can be associated to both irrigation and long-term drought as shown by [Forootan et al. \(2017\)](#).

4.3.1. Contribution of ENSO in DMDA-Derived Water Storage Components

To demonstrate that the DMDA-derived surface and sub-surface water storage estimates are closer to the reality than those from any individual model, we extract

the dominant ENSO mode from the DMDA estimates and compare them with climate indices (see e.g., [Anyah et al., 2018](#)) in terms of temporal correlation coefficients with the ENSO index (-Niño 3.4 index, Fig. 5, 6, and 7). The reason for this comparison is that GRACE captures considerable variability due to the ENSO events ([Phillips et al., 2012](#); [Forootan et al., 2018](#)). Therefore, by merging multi-model outputs with GRACE data, their skill in representing water storage changes due to large-scale teleconnections would be improved.

In order to extract the ENSO modes from the DMDA-derived water storage estimates and the original outputs of the six models (PCRGLOB-WB, SURFEX-TRIP, LISFLOOD, HBV-SIMREG, W3RA, and ORCHIDEE) Principal Component Analysis (PCA, [Lorenz, 1956](#)) method is applied after removing the long-term linear trend and seasonality from hydrological components. More details about PCA results and extracting ENSO modes from DMDA water storage components are reported in [ESM-section 6](#).

Figure 5 shows temporal correlations between the ENSO mode of groundwater (from DMDA and original models) and the ENSO index. Maximum and minimum correlation of 0.75 and 0.53 corresponding to a maximum lag of up to 2 months are found globally between the DMDA groundwater and the ENSO index, respectively. Smaller correlations are found between the original models and the ENSO index. Among these models, W3RA and HBV-SIMREG indicate stronger correlations (~ 0.6 and ~ 0.4 respectively) with the ENSO index with a maximum lag of 2 months. Other models such as LISFLOOD and SURFEX-TRIP indicate notably different correlations (compared to HBV-SIMREG and W3RA as well as that of DMDA) with ENSO in various basins. We find small positive correlations with a maximum value of 0.3 between original PCR-GLOBWB’s groundwater and the ENSO index. Although the maximum lag of 3 month is estimated in most of the 33 basins, a lag of 15 months is estimated for the Nile, Okavango, and Zambezi (Africa), Colorado and Nelson (North America), Ob, Lena, and Yellow (Asia) River Basins, which are likely not realistic (see, e.g., [Awange et al., 2014](#); [Anyah et al., 2018](#)).

FIGURE 5

Similar assessments are performed between the soil moisture and surface water storage changes with the ENSO index and the results are shown in Figs. 6 and 7. Correlation coefficients of up to 0.8 are computed from the DMDA estimates with a maximum lag of up to 2 months. Among the six models, correlation in soil moisture of the SURFEX-TRIP and LISFLOOD models is found to be the highest, i.e., correlations of 0.6 to 0.8 within the 33 river basins examined here. PCR-GLOBWB and W3RA show a correlation of ~ 0.5 , while those from HBV-

711 SIMREG and ORCHIDEE are different from our other estimations, for example,
 712 less than 0.1 in the Niger and Nile River Basins, and greater than 0.75 in North
 713 Asia. *Khaki et al. (2018b)* indicate that over the Nile River Basin, all the three
 714 hydrological components, (i.e., groundwater, surface water, and soil moisture) are
 715 strongly influenced by ENSO. Therefore, the obtained correlation of 0.1 in the Nile
 716 River Basin from HBV-SIMREG is likely not realistic.

FIGURE 6

717 The DMDA-derived surface water storage is compared with those of PCR-GLOBWB,
 718 SURFEX-TRIP, and ORCHIDEE, which contain the surface water storage compart-
 719 ment. The correlation coefficients are found to be generally smaller than those of soil
 720 moisture and groundwater components (with a maximum of 0.5), which likely shows
 721 that the modelling of surface water needs improvement because in reality surface wa-
 722 ter in lakes and rivers within regions like East Africa shows an immediate response to
 723 ENSO (e.g., *Becker et al., 2010; Khaki et al., 2018b*). Figure 7 shows that the surface
 724 water storage output of SURFEX-TRIP had the highest correlations with the ENSO
 725 index in all basins of America (values between 0.33 and 0.51) and Africa (values
 726 between 0.23 and 0.48), while ORCHIDEE shows the highest correlations (values
 727 between 0.32 and 0.58) in most parts of Asia. The correlations for PCR-GLOBWB
 728 are found to be relatively smaller, i.e., between 0.1 and 0.2 with lags of between 5-12
 729 months. Comparisons between the DMDA and original model outputs indicate that
 730 combining models with GRACE data improve the correlations with the ENSO index
 731 and the correlation lags are considerably reduced globally. It is worth mentioning
 732 that the DMDA results that are presented here are derived by setting the α value
 733 in equation (14) to 0.9. This means that we assume a 36 month temporal correla-
 734 tions between water storage simulations of the six models. This value guarantee an
 735 extraction of the ENSO modes within two PCA modes after merging GRACE and
 736 model outputs.

FIGURE 7

737 4.4. Evaluating the DMDA Results with satellite altimetry observation

738 To validate our results, TWS and surface water derived from DMDA and six
 739 hydrological models are compared with independent surface water observations from
 740 satellite altimetry. The results are shown for various regions with reliable satellite
 741 altimetry measurements such as the Nile, Niger, and Zambezi River Basins in Africa,
 742 Ob and Euphrates in Asia, St' Lawrence and Nelson in North America, and Orinoco
 743 in South Africa. Here, we assessed 14 lakes located in the 8 mentioned river basins.

744 Comparisons are performed in terms of correlation coefficients between TWS and
745 surface water estimates (within the river basins), and water mass variations within
746 the lakes (i.e., lake level heights from satellite altimetry data are converted to mass
747 variations following [Moore and Williams \(2014\)](#)). The numerical results are sum-
748 marized in Table 5, which indicates that after implementing the DMDA method,
749 correlation coefficients are increased in most of the lakes. High values are found in
750 the Nile River Basin, e.g., Tana Lake (0.718), Euphrates (Tharthar Lake, 0.569), and
751 Niger (Chad Lake, 0.558), while low values are found in the Kainiji Lake of the Niger
752 River Basin (0.102) and Winnipegosis of the Nelson River Basins (0.249). It should
753 be noted here that although low correlations are found for some lakes, the values are
754 increased when compared with the original model simulations. More details can be
755 found in [ESM-section 7](#).

TABLE 5

756 5. Summary and Conclusion

757 In this study, the method of Dynamic Model Data Averaging (DMDA) is intro-
758 duced, which can be used (1) to compare multi-model (individual) water storage
759 simulations with GRACE-derived Terrestrial Water Storage (TWS) estimates; and
760 (2) to separate GRACE TWS into horological water storage compartments. DMDA
761 combines the property of Kalman Filter (equations (9), (10)) and a Bayesian weight-
762 ing (equation (11)) to fit multi-model water storage changes to GRACE TWS esti-
763 mates. The method is flexible in accounting for errors in observations and a priori
764 information (equation 9 and equation 10), and can deal with state vectors of different
765 length.

766 The benefit of the DMDA method over the commonly used PF or PS methods
767 are twofold: 1) these methods might not be efficient for high-dimensional fusion
768 tasks (e.g., [Snyder et al., 2008](#); [Van Leeuwen, 2009](#)) such as the global hydrological
769 application presented here, but the DMDA’s computational load is lower than these
770 techniques; 2) DMDA provides time-variable weights that can be used to under-
771 stand the behavior of a priori information (here the output of hydrological models)
772 against GRACE TWS estimates, while considering their errors. The advantage of
773 the DMDA over the Ensemble Kalman Filter-based of techniques is that the poste-
774 rior distributions are computed through a Bayesian rule that result in more reliable
775 estimations of states and their errors, while avoiding the high computational loads
776 of the PF techniques.

777 A realistic synthetic example was defined to evaluate the performance of DMDA
778 (Fig. 2), which showed that the method is able to correctly separate GRACE TWS

estimates into its individual hydrological components. We also showed that the DMDA’s estimation of temporal weights (for each model) was close to the reality, and can be used to assess the performance of available models. Based on the real data, we showed that the representation of linear trends and seasonality within global hydrological models, as well as their water storage changes due to the El Niño Southern Oscillation (ENSO) can be improved using DMDA, while considering the uncertainties of models and observations (see Fig. 1). Our results also showed that how the DMDA method is able to deal with models with different structures, and how it updates their water storage simulations while considering their errors. Considering these arguments, we believe that the new water storage estimates, i.e., models combined with GRACE, are of great values and can be used for further hydrological and climate research investigations compared to model or GRACE only estimates. Therefore, the presented results can be considered as one step forward to improve model deficiencies following the insights of *Scanlon et al. (2018)*. In what follows, the main conclusions and remarks of this study are summarized.

- Estimated weights (Fig. 3) showed that the PCR-GLOBWB model gained the largest weights, thus, it contributed the most in the DMDA-derived TWS in North Asia, North America, and the center of Africa. SURFEX-TRIP performed best within basins with dominant surface water storage changes, as well as in snow-dominant regions. The LISFLOOD and ORCHIDEE models were found to perform well within irrigated basins, and those affected by ENSO events.
- DMDA results in Fig. 4 (a1) showed that considerable trends exist in groundwater storage changes within the Ganges, Indus, and Euphrates basins during 2002–2012. These changes are dominantly influenced by anthropogenic modifications. Trends in soil moisture (Fig. 4 (b1)) were found to be mostly related to meteorological prolonged drought events such as those in the Brahmaputra and Euphrates River Basins.
- DMDA was able to modify the ENSO mode of water storage variability in most of the world’s 33 largest river basins (see Fig. 5, Fig. 6, and Fig. 7). DMDA assigned the biggest corrections of ENSO mode in groundwater to the Nile, Murray, Tocantins, Ob, Okavango and Orange River Basins. The highest corrections of the ENSO mode in soil moisture were found for the Nile, Niger, Zambezi, and Amur River Basins, and in surface water to Nile, Niger, Congo, Tocantins, and Murray River Basin. For example, the correlation coefficient between groundwater storage and ENSO in the Murray River Basin changed

from -0.2 to 0.6. For the Nile River Basin, they changed from 0.1 to 0.4 for soil moisture, and from 0.3 to 0.7 for the surface water compartment.

- Comparison between TWS and surface water derived from DMDA with independent surface water observations from satellite altimetry (Fig. ESM.15 and Fig. ESM.16 in ESM-section 7) showed that, DMDA was able to correctly detect the best performing model and maximize its contribution in the dynamic averaging process which enhanced the reality of water storage estimates.
- To implement the DMDA in this study a forgetting factor of 0.95 was considered in equation (6), which is equivalent to the temporal dependency in estimating time variable regression parameters in equation (2). In section 3, it was shown that this selection is equivalent to 18 months temporal dependency between GRACE TWS observations and model simulations. This value is selected because the DMDA results were closest to that of GRACE. After selecting this value, we also obtained a distinguishable ENSO mode from the DMDA-derived TWS and individual water storage estimates. Therefore, we conclude that this temporal lag might be considered in other works that attempt to apply sequential mergers or smoothers to assimilate observed water storage data into models.
- In order to reduce the computational load of this work, instead of implementing a Markov Chain Monte Carlo (MCMC) technique to estimate the transition matrix between models in equation (13), a forgetting factor of 0.9 was considered in equations (14). This might be replaced with an efficient MCMC implementation in future.

The DMDA method, introduced in this study, has the potential to be used in different climate and hydrological applications to compare available models (which can be of various types of hydrological or climate models) against reliable observations. It can also be used to generate ensembles from multi-model outputs such as climate projections. The application of this study can also be extended by incorporating other types of remote sensing observations such as satellite based soil moisture or water level data beside those of GRACE. A secondary application of the DMDA can also be devoted to its application for predicting (or extrapolating) water storage estimates. To achieve this purpose, however, the DMDA's formulation needs to be extended. For example, one approach can be to use the DMDA weights, which are computed for the period of study, to identify best models in different river basins covering different seasons. By analysing this information and knowing the TWS in the

850 future, one can use a combination of different model runs (weighted by the DMDA
851 outputs) and extrapolate the surface and sub-surface water storage estimates.

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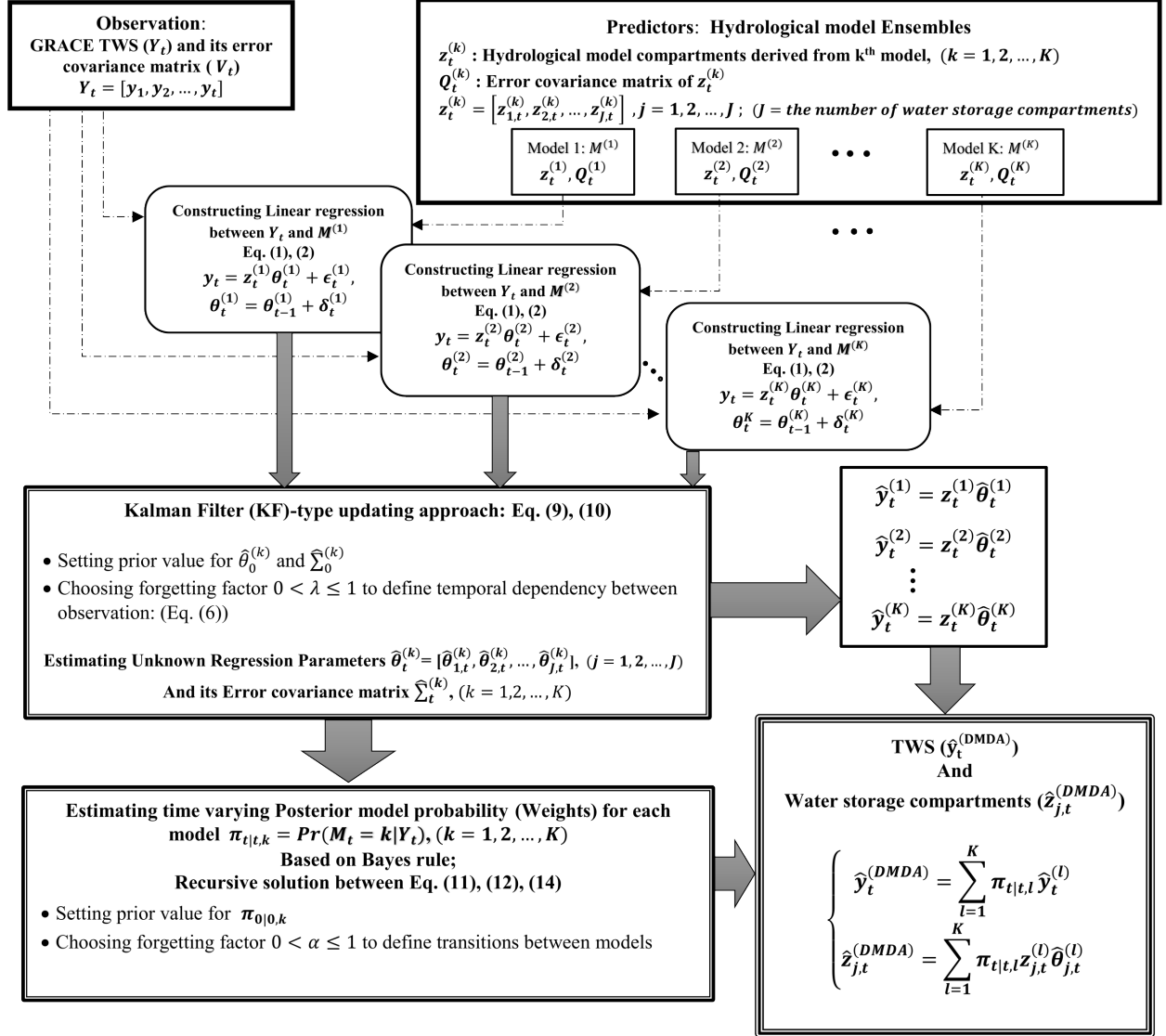
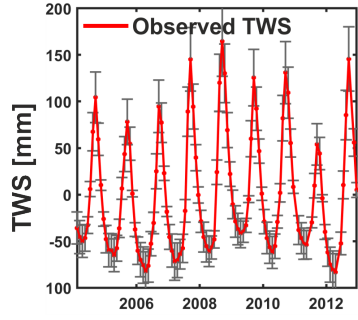
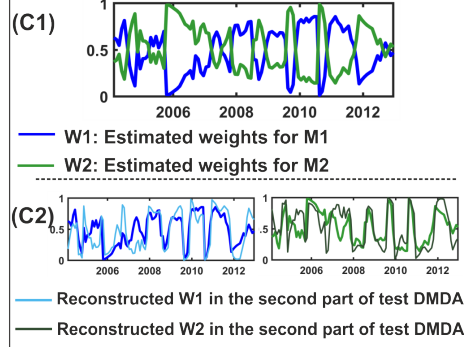


Figure 1: Flowchart of the Dynamic Model Data Averaging (DMDA) method. The framework can accept an arbitrary number of models and it can be extended to accept various type of observations.

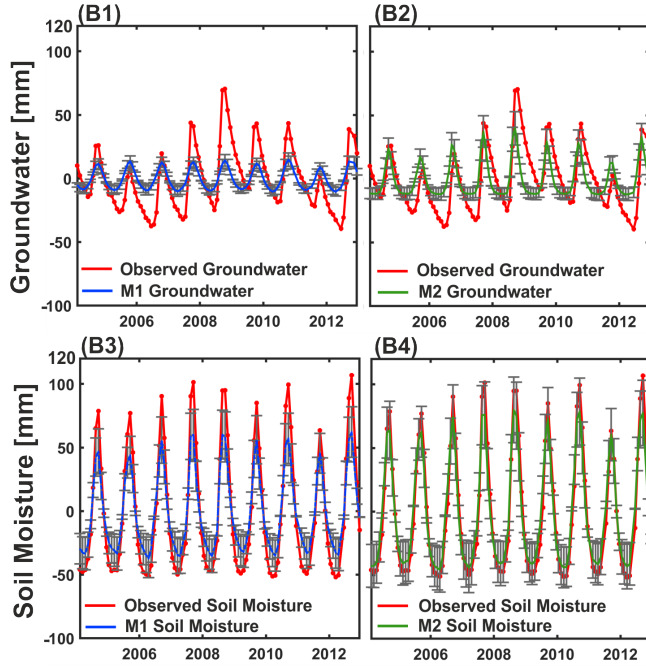
(A): Observation (Input)



(C): Estimated weights for models



(B): Predictors (Input)



(D): DMDA and BMA updated components

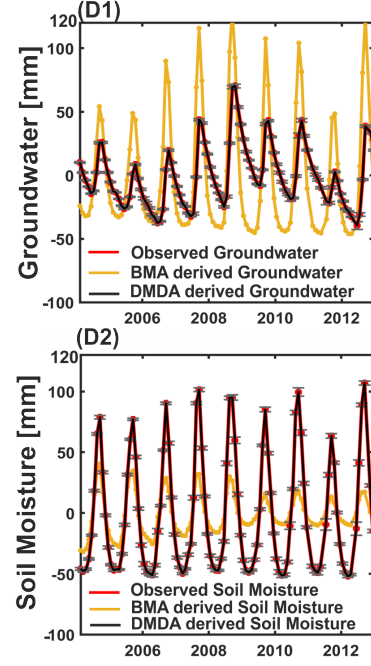


Figure 2: A synthetic example, where DMDA is applied in a controlled set up, to integrate 2 hydrological models (here selected as SURFEX-TRIP and LISFLOOD) with simulated observed TWS to separate its compartments (i.e., groundwater and soil moisture). All data sets in this simulation is related to the Niger River Basin and covering the period between 2002–2012; Figure 2 (A) shows TWS simulated from PCR-GLOBWB (here standing in for observed TWS); Figure 2 (B) shows the time series of groundwater and soil moisture derived from model 1 (B1, B3) and model 2 (B2, B4), which are considered as the input predictors in DMDA; Figure 2 (C1) presents the time varying weights estimated for two selected model, and Figure 2 (C2) shows the reconstructed of weights in the second step of our simulation. Figure 2 (D1) and (D2) show the updated hydrological components obtained from the DMDA and BMA method and comparison between the obtained results and the expected values derived from simulated observation data.

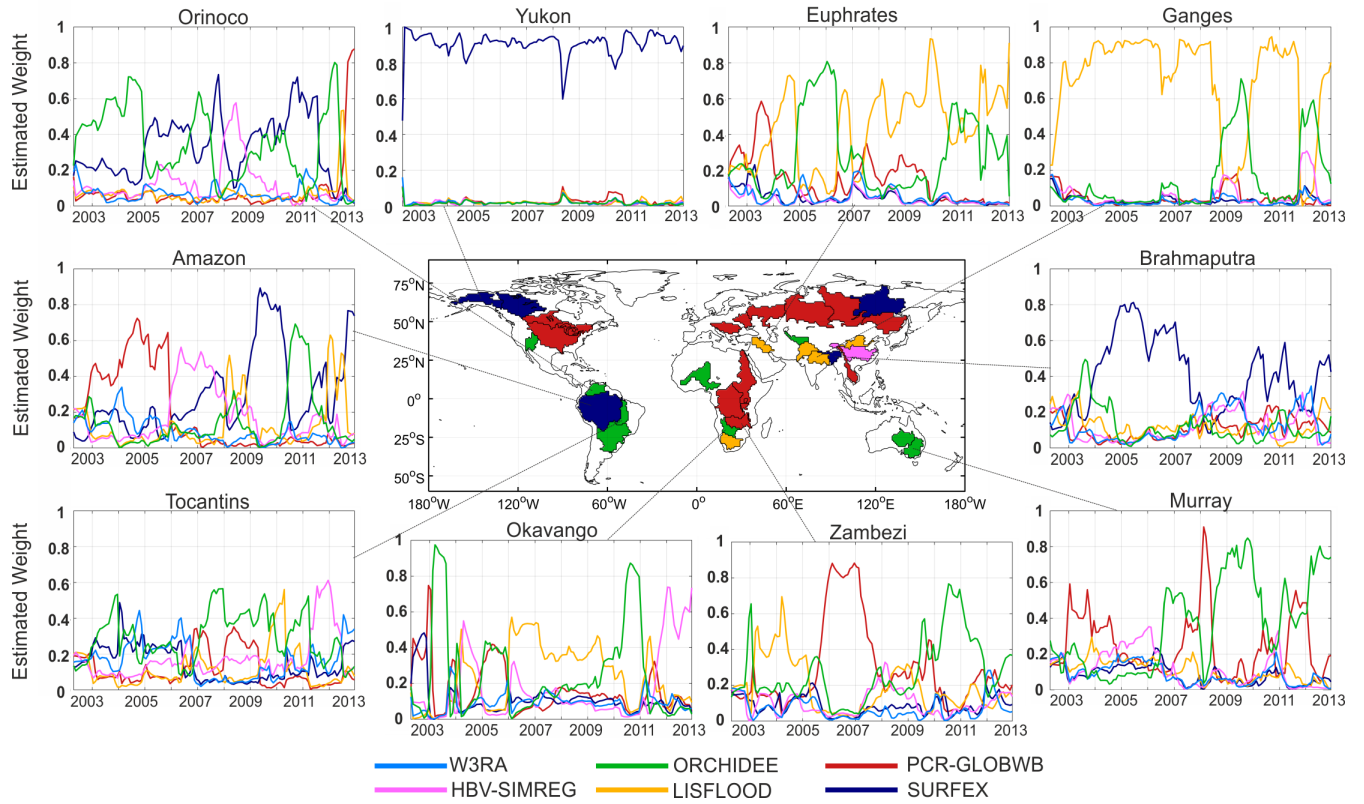


Figure 3: Posterior model probabilities for the six initially considered models, over 10 selected river basins with the biggest RMSEs computed using GRACE and models-derived TWS. In the middle of Fig 3 the most contributed models in the DMDA-derived TWS are shown over the world's 33 largest river basins, covering the period of 2002–2012.

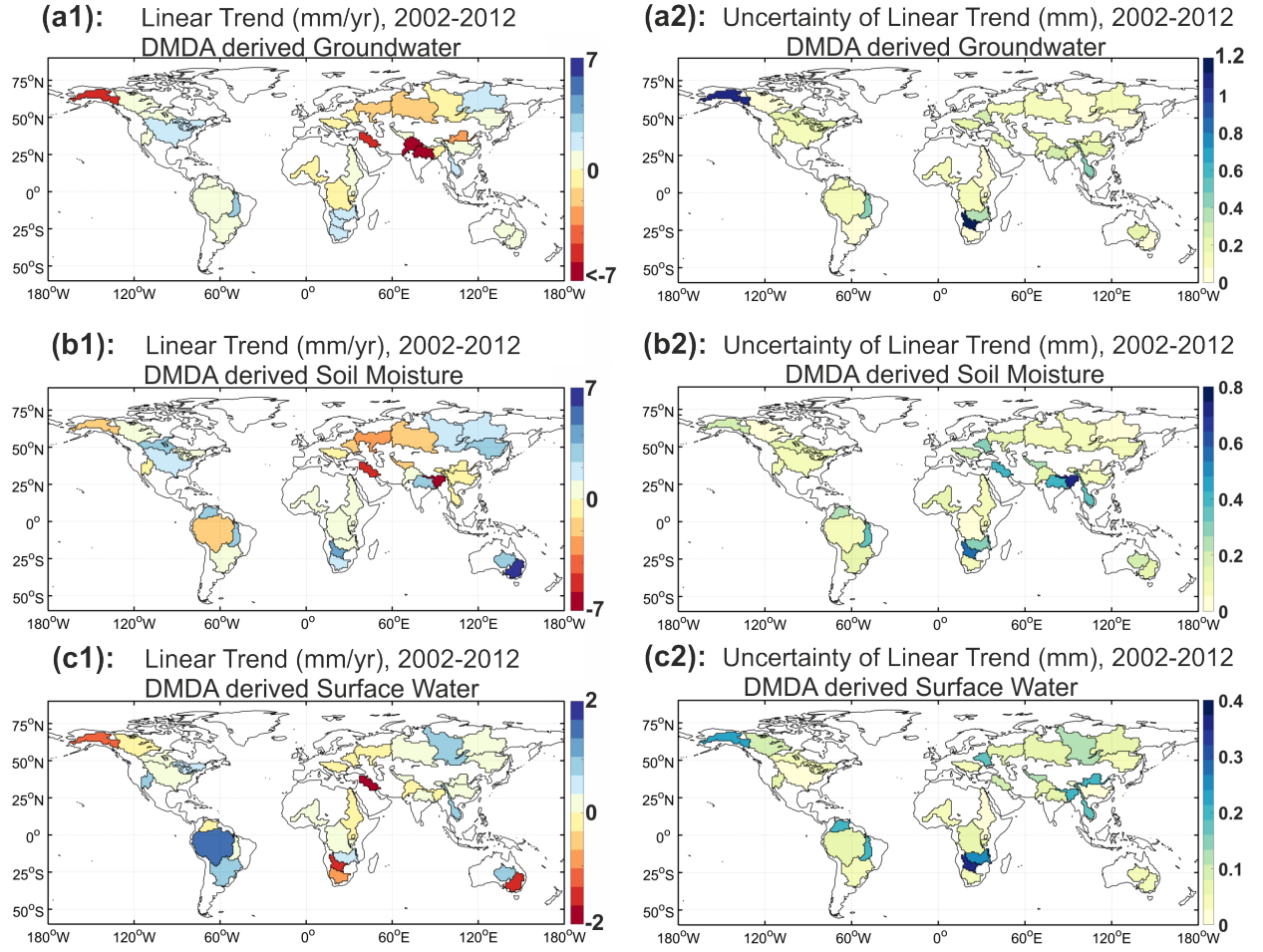


Figure 4: Long-term (2002–2012) linear trend in the DMDA-derived groundwater (a1), soil moisture (b1), and surface water (c1) components, expressed in mm/yr. The uncertainty of these fitted linear trends are shown in (a2), (b2), (c2) respectively.

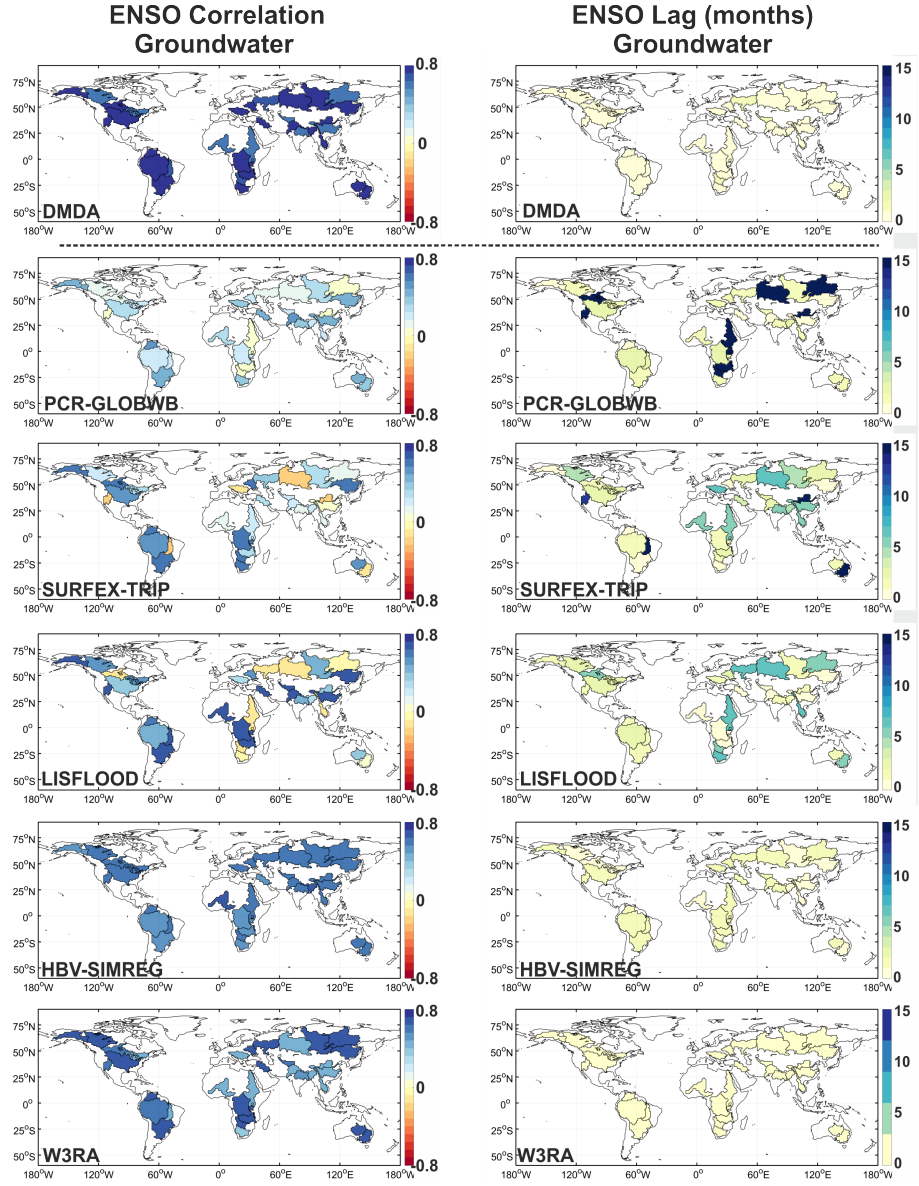


Figure 5: Correlation coefficients and their lags between the ENSO (-Niño 3.4 index) and groundwater estimates derived from the DMDA method and hydrological models used in this study for the period of 2002–2012.

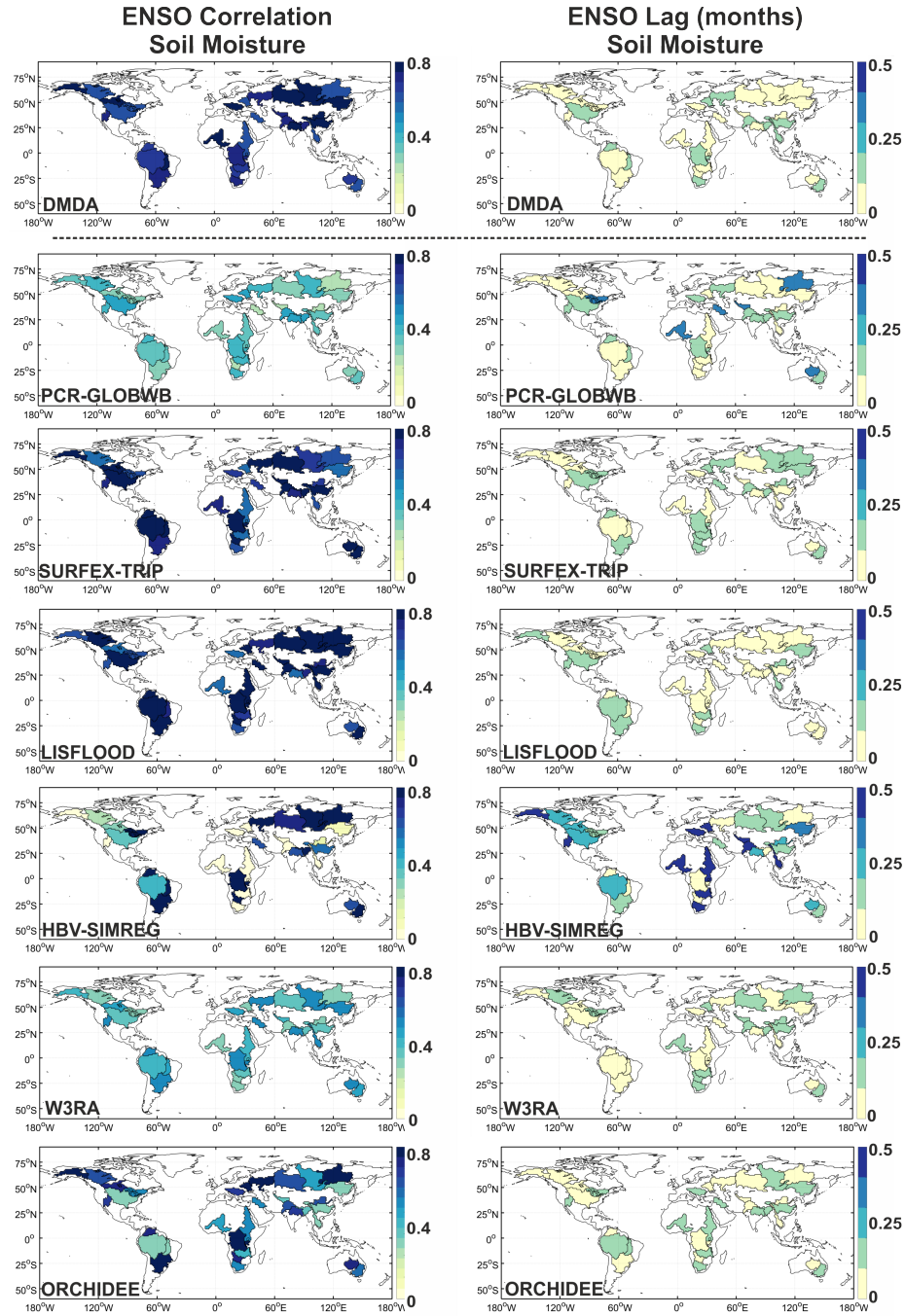


Figure 6: Correlation coefficients and their lags between the ENSO (-Niño 3.4 index) and soil moisture estimates derived from the DMDA method and hydrological models used in this study for the period of 2002–2012.

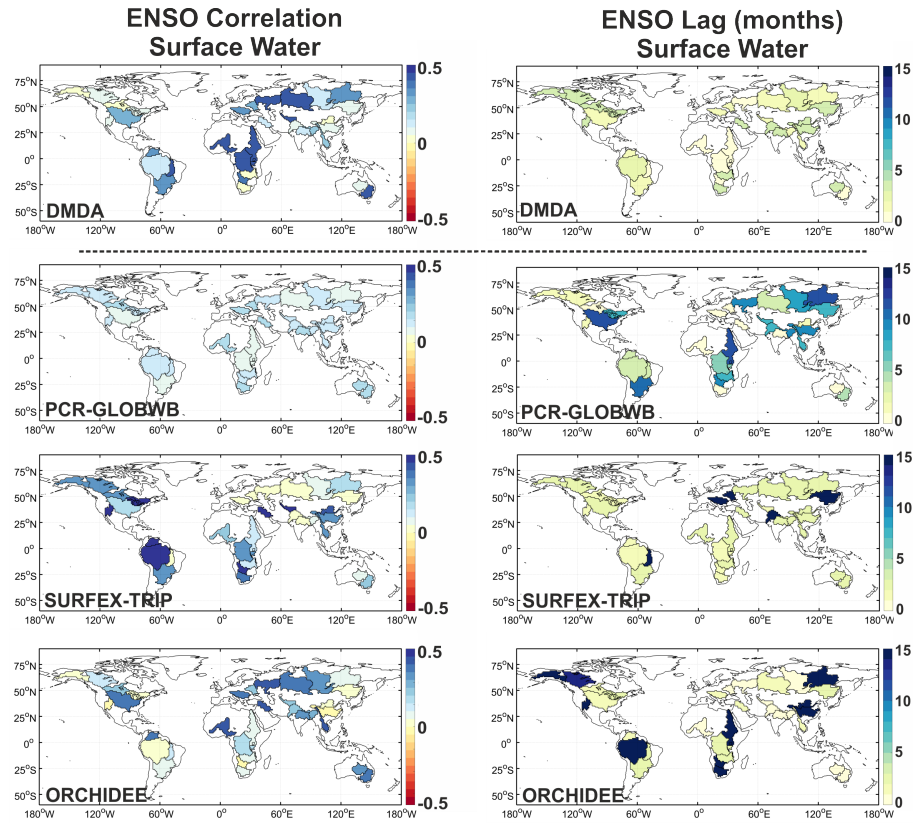


Figure 7: Correlation coefficients and their lags between the ENSO (-Niño 3.4 index) and surface water estimates derived from the DMDA method and hydrological models used in this study for the period of 2002–2012.

Table 1: Overview of models used in this study and their water storage components.

Model	Water Storage Compartments						
	GroundWater	Soil layer	Surface Water	Canopy	Snow	Snow layer	Water Use
PCR-GLOBWB	Yes	2	Yes	Yes	Yes	1	No
W3RA	Yes	3	No	No	Yes	1	No
HBV-SIMREG	Yes	1	No	No	Yes	1	No
SURFEX-TRIP	Yes	14	Yes	Yes	Yes	12	No
LISFLOOD	Yes	2	No	No	Yes	1	Yes
ORCHIDEE	No	11	Yes	No	Yes	6	irrigation

Table 2: An overview of satellite altimetry observation used to validate DMDA results.

Lake	River Basin	Lake mid point	Latitude range of pass	Satellite pass	Cycle
Nasser	Nile	23.31°N 32.83°E	[22.91°N 23.66°N]	94	48
Tana	Nile	12.11°N 37.40°E	[11.95°N 12.19°N]	94	38
chad	Niger	13.01°N 14.38°E	[12.94°N 13.05°N]	248	25
Kainiji	Niger	10.49°N 4.50°E	[10.40°N 10.50°N]	135	21
Malawi	Zambezi	10.84°S 34.40°E	[12.042°S 9.70°S]	44	4
Tanganyika	Zambezi	6.41°S 29.23°E	[8.44°S 4.461°S]	222	11
Guri	Orinoco	7.37°N 117.12°W	[7.06°N 7.67°N]	152	69
Winnipeg	Nelson	53.18°N 98.21°W	[52.82°N 53.55°N]	195	9
Winnipegosis	Nelson	51.91°N 100.01°W	[51.85°N 52.05°N]	195	17
Erie	St. Lawrence	42.11°N 81.48°W	[41.60°N 42.54°N]	193	45
Ontario	St. Lawrence	43.56°N 77.47°W	[43.35°N 43.83°N]	15	36
Tharthar	Euphrates	33.87°N 43.37°E	[33.75°N 34.00°N]	133	70
Urmia	Euphrates	37.25°N 45.45°E	[37.25°N 37.31°N]	133	4
Chany	Ob	54.96°N 77.33°E	[54.94°N 55.02°N]	5	28

Table 3: Magnitude of simulated predictors, observations, and DMDA results in a controlled synthetic simulation.

Hydrological Compartment	Model name	Min [mm]	Max [mm]	RMS [mm]
Groundwater (First model)	LISFLOOD	-10.5	16.1	7.9
Groundwater (Second model)	SURFEX-TRIP	-12.1	39.8	14.2
Groundwater (Expected value of DMDA)	PCR-GLOBWB	-39.5	70.4	24.2
Groundwater (DMDA result)	DMDA Output	-35.3	92.3	19.9
Groundwater (BMA result)	BMA Output	-46.0	130.2	43.8
Soil Moisture (First model)	LISFLOOD	-37.4	62.2	30.8
Soil Moisture (Second model)	SURFEX-TRIP	-45.7	79.9	41.5
Soil Moisture (Expected value of DMDA)	PCR-GLOBWB	-52.0	107.9	48.7
Soil Moisture (DMDA result)	DMDA Output	-58.5	113.8	51.2
Soil Moisture (BMA result)	BMA Output	-40.8	49.6	21.0
TWS (First model)	LISFLOOD	-46.8	75.5	37.2
TWS (Second model)	SURFEX-TRIP	-57.6	115.2	54.6
TWS (Expected value of DMDA results)	PCR-GLOBWB	-83.3	164.5	64.2
TWS (DMDA result)	DMDA Output	-77.8	153.8	63.2
TWS (BMA result)	BMA Output	-77.8	153.8	63.2
$ \Delta _{\text{Groundwater}}$	$ \text{LISFLOOD} - \text{Expected value} $	0	58.1	11.2
$ \Delta _{\text{Groundwater}}$	$ \text{SURFEX} - \text{Expected value} $	0	45.8	10.3
$ \Delta _{\text{Groundwater}}$	$ \text{DMDA} - \text{Expected value} $	0	31.2	5.3
$ \Delta _{\text{Groundwater}}$	$ \text{BMA} - \text{Expected value} $	0	87.6	20.4
$ \Delta _{\text{Soil Moisture}}$	$ \text{LISFLOOD} - \text{Expected value} $	0	46.8	9.6
$ \Delta _{\text{Soil Moisture}}$	$ \text{SURFEX} - \text{Expected value} $	0	29.3	5.7
$ \Delta _{\text{Soil Moisture}}$	$ \text{DMDA} - \text{Expected value} $	0	29.2	5.2
$ \Delta _{\text{Soil Moisture}}$	$ \text{BMA} - \text{Expected value} $	0	89.5	18.6
$ \Delta _{\text{TWS}}$	$ \text{LISFLOOD} - \text{Expected value} $	0	94.7	18.6
$ \Delta _{\text{TWS}}$	$ \text{SURFEX} - \text{Expected value} $	0	60.9	14.1
$ \Delta _{\text{TWS}}$	$ \text{DMDA} - \text{Expected value} $	0	24.2	6.2
$ \Delta _{\text{TWS}}$	$ \text{BMA} - \text{Expected value} $	0	31.4	8.4

Table 4: The amplitude of linear trend [mm/yr] and its uncertainty, fitted to the DMDA-derived groundwater, soil Moisture, and surface water, during 2002–2012.

Basin		DMDA	DMDA	DMDA
ID	Name	GroundWater	Soil Moisture	Surface Water
1	Amazon	0.17 ± 0.12	-1.92 ± 0.09	1.43 ± 0.06
2	Amur	0.46 ± 0.06	2.61 ± 0.09	0.25 ± 0.03
3	Aral	0.02 ± 0.08	-1.43 ± 0.22	0.21 ± 0.12
4	Brahmaputra	-0.44 ± 0.16	-7.00 ± 0.69	-0.13 ± 0.21
5	Caspian-Volga	-2.06 ± 0.15	-2.98 ± 0.16	-0.02 ± 0.07
6	Colorado	0.80 ± 0.11	-0.75 ± 0.09	0.82 ± 0.08
7	Congo	-0.72 ± 0.08	0.59 ± 0.03	0.06 ± 0.06
8	Danube	-0.47 ± 0.18	-0.75 ± 0.21	-0.08 ± 0.04
9	Dnieper	-0.5 ± 0.29	-2.27 ± 0.28	-0.03 ± 0.18
10	Euphrates	-5.36 ± 0.23	-5.75 ± 0.39	-2.09 ± 0.09
11	Lake Eyre	0.55 ± 0.16	2.42 ± 0.19	0.77 ± 0.04
12	Ganges	-14.77 ± 0.25	2.69 ± 0.40	0.29 ± 0.05
13	Indus	-8.26 ± 0.16	1.10 ± 0.13	-0.06 ± 0.07
14	Lena	1.74 ± 0.11	1.94 ± 0.05	0.20 ± 0.08
15	Mackenzie	0.51 ± 0.06	0.12 ± 0.05	-0.05 ± 0.10
16	Mekong	1.58 ± 0.43	-0.79 ± 0.33	0.83 ± 0.17
17	Mississippi	1.25 ± 0.09	1.36 ± 0.09	0.33 ± 0.02
18	Murray	0.06 ± 0.06	6.66 ± 0.15	-1.47 ± 0.04
19	Nelson	0.70 ± 0.18	2.45 ± 0.15	0.11 ± 0.03
20	Niger	-1.14 ± 0.15	0.75 ± 0.15	0.32 ± 0.05
21	Nile	0.45 ± 0.06	0.77 ± 0.06	-0.05 ± 0.02
22	Ob	-1.42 ± 0.08	-1.54 ± 0.06	0.05 ± 0.07
23	Okavango	1.74 ± 1.31	3.92 ± 0.55	-1.42 ± 0.37
24	Orange	1.32 ± 0.05	1.28 ± 0.06	-0.85 ± 0.05
25	Orinoco	0.87 ± 0.11	3.45 ± 0.26	-0.22 ± 0.19
26	Parana	0.68 ± 0.08	0.03 ± 0.13	1.04 ± 0.04
27	St. Lawrence	1.49 ± 0.18	1.07 ± 0.07	0.48 ± 0.05
28	Tocantins	2.41 ± 0.47	2.37 ± 0.35	0.08 ± 0.21
29	Yangtze	0.55 ± 0.23	-0.30 ± 0.09	0.20 ± 0.02
30	Yellow	-3.50 ± 0.14	-0.27 ± 0.05	0.08 ± 0.21
31	Yenisei	-0.26 ± 0.07	1.79 ± 0.06	0.75 ± 0.11
32	Yukon	-4.73 ± 1.08	-1.52 ± 0.20	-1.11 ± 0.23
33	Zambezi	1.19 ± 0.38	0.65 ± 0.31	0.35 ± 0.25

Table 5: Correlation between satellite altimetry observation and: I) TWS , II) Surface Water (SW) derived from GRACE, DMDA, and individual models, during 2002–2012.

Basin	Water storage	Correlation between Altimetry Obs. and:							
		GRACE	DMDA	PCR-GLOBWB	SURFEX-TRIP	LISFLOOD	HBV-SIMREG	W3RA	ORCHIDEE
Nile (Nasser Lake)	TWS	0.358	0.381	0.326	0.239	0.095	-0.082	0.001	0.180
	SW	-	0.462	0.363	0.441	-	-	-	-0.046
Nile (Tana Lake)	TWS	0.682	0.718	0.602	0.569	0.517	0.302	0.231	0.635
	SW	-	0.492	0.340	0.603	-	-	-	0.455
St. Lawrence (Erie Lake)	TWS	0.353	0.261	0.271	0.010	-0.121	-0.114	-0.087	-0.010
	SW	-	0.432	0.483	0.126	-	-	-	0.227
St. Lawrence (Ontario Lake)	TWS	0.410	0.364	0.353	0.110	-0.063	-0.064	-0.023	0.037
	SW	-	0.582	0.572	0.273	-	-	-	0.239
Euphrates (Tharthar Lake)	TWS	0.698	0.569	0.225	0.021	0.103	-0.057	0.043	0.182
	SW	-	0.236	0.127	0.093	-	-	-	-0.282
Euphrates (Urmia Lake)	TWS	0.737	0.628	0.223	0.080	0.148	0.021	0.095	0.185
	SW	-	0.172	0.170	0.131	-	-	-	-0.325
Ob (Chany Lake)	TWS	0.393	0.482	0.371	0.303	0.336	0.338	0.348	0.328
	SW	-	0.296	0.278	0.177	-	-	-	-0.333
Zambezi (Malawi Lake)	TWS	0.552	0.632	0.362	0.277	0.346	0.225	0.246	0.391
	SW	-	0.382	0.247	0.410	-	-	-	0.394
Zambezi (Tanganyika Lake)	TWS	0.414	0.365	0.231	0.192	0.121	0.117	0.128	0.160
	SW	-	0.243	0.096	0.241	-	-	-	-0.093
Niger (Chad Lake)	TWS	0.576	0.558	0.436	0.318	0.308	0.065	0.188	0.519
	SW	-	0.657	0.511	0.616	-	-	-	0.689
Niger (Kainiji Lake)	TWS	0.132	0.102	-0.002	-0.149	-0.174	-0.383	-0.278	0.079
	SW	-	0.282	0.126	0.200	-	-	-	0.214
Orinoco (Guri Lake)	TWS	0.585	0.539	0.332	0.427	0.431	0.321	0.301	0.434
	SW	-	0.421	0.314	0.390	-	-	-	0.318
Nelson (Winnipeg Lake)	TWS	0.285	0.270	0.139	-0.185	-0.444	-0.440	-0.389	-0.279
	SW	-	0.104	-0.290	0.072	-	-	-	0.012
Nelson (Winnipegosis Lake)	TWS	0.216	0.249	0.238	0.135	-0.09	-0.164	-0.088	-0.065
	SW	-	0.098	-0.321	-0.015	-	-	-	-0.480