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- Not all bipolar disorder outcomes are created equal:
- 2 occupational dysfunction and hospital admissions associate
- with different polygenic profiles

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#### **ABSTRACT**

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Objective: Hospital care due to severe mood episodes are consequences for many but not all persons with bipolar disorder. Likewise, some but not all patients suffer long-term occupational dysfunction that extend beyond acute mood episodes. It is not known whether these dissimilar outcomes of bipolar disorder are driven by different polygenic profiles. We assessed how polygenic scores (PGS) for major psychiatric disorders and educational attainment associate with occupational functioning and psychiatric hospital admissions in bipolar disorder. Method: A total of 4,782 bipolar disorder patients and 2,907 control subjects were genotyped and linked to Swedish national registers. Longitudinal measures from at least 10 years of registry data were used to derive percent of years without employment, with long-term sick leave, and the mean number of psychiatric hospital admissions per year. Ordinal regression was used to test the associations between outcomes and PGS for bipolar disorder, schizophrenia, major depressive disorder (MDD), attention-deficit/hyperactivity disorder (ADHD), and educational attainment. Replication analyses of hospital admissions were conducted in the Bipolar Disorder Research Network (N=4,219). **Results**: Long-term sick leave and unemployment in bipolar disorder were significantly associated with PGS for schizophrenia, ADHD, MDD, and educational attainment, but not bipolar disorder. By contrast, the number of hospital admissions per year associated with higher PGS for bipolar disorder and schizophrenia, but not with the other PGS. Conclusions: Bipolar disorder severity (indexed by hospital admissions) associates with a different polygenic profile than long-term occupational dysfunction. This has clinical implications, suggesting that mitigating occupational dysfunction requires other interventions than those deployed to prevent mood episodes.

#### INTRODUCTION

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Bipolar disorder is characterized by recurring periods of elevated or depressed mood. Acute mood episodes are managed in psychiatric outpatient or inpatient care depending on the severity. Manic episodes typically warrant hospital care. Hospitalisation is also considered when a depressive episode is too severe to be managed in psychiatric outpatient care. The need for inpatient care varies among persons with bipolar disorder and many bipolar disorder patients can be successfully managed in outpatient care only. The need for hospitalisation thus reflects the severity of acute mood episodes. Aside from the impairments associated with acute mood episodes, persons with bipolar disorder might suffer from persistent psychosocial and occupational deficits. Indeed, as many as 30-60% of bipolar disorder patients fail to regain full functioning in occupational and social domains after illness onset<sup>1</sup>. Only half of persons with bipolar disorder are within the workforce<sup>2,3</sup>. This is not only much lower than the general population, but also lower than persons with unipolar depression<sup>3</sup>. Bipolar disorder is in fact one of the main causes of disability among young adults<sup>4</sup>. Occupational disability is also an important driver of the societal costs associated with bipolar disorder, where one estimate found that 75% of the total societal costs could be attributed to sick leave and early retirement<sup>5</sup>. To mitigate occupational dysfunction in bipolar disorder, it is key to identify the driving mechanisms. One would assume that persons with more severe mood episodes have worse long-term functional outcome. But classical measures of illness severity—such as the number of psychiatric hospital admissions—in fact correlates poorly with psychosocial functioning<sup>6</sup>. Instead, we and others showed an association between executive functioning and occupational functioning<sup>7,8</sup>. This suggests that abiding functional sequels of bipolar disorder is determined by other factors than those that cause mood symptoms manifested during acute mood episodes.

Polygenic scores (PGS) can be calculated for a range of different traits and is a tool to unearth genetic components that drive specific subphenotypes within a diagnostic category. In bipolar disorder, specific subphenotypes such as number of hospital admissions<sup>9</sup> and psychosis<sup>10</sup> have previously been associated with polygenic risk for schizophrenia. It is, however, not known whether long-term occupational functioning associates with particular polygenic liabilities.

The aim of this study is to assess the respective polygenic profiles of occupational functioning and severity of mood episodes in bipolar disorder. To this end, we calculated PGS for bipolar disorder, schizophrenia, Major Depressive Disorder (MDD), Attention Deficit Hyperactivity Disorder (ADHD), and educational attainment in a cohort of persons with bipolar disorder (N=4,782), control sample (N=2,907), and in the replication cohort Bipolar Disorder Research Network (BDRN, N=4,219). Long-term longitudinal data on occupational functioning (unemployment and long-term sick leave) and psychiatric hospital admissions in the main analyses were obtained by linkage to Swedish national registers.

# **METHODS**

#### **Population**

Persons diagnosed with bipolar disorder type 1, bipolar disorder type 2, or bipolar disorder not otherwise specified (NOS) in the Swedish Bipolar Collection (SWEBIC) were included in the study. Most study participants in SWEBIC were enrolled through the Swedish National Quality Register for bipolar disorder (BipoläR) 11. A smaller number were recruited from the St. Göran Bipolar Project<sup>12-14</sup>, and some were identified in the Swedish National Patient Register using a validated algorithm<sup>15</sup>. Controls were obtained from the general population in Sweden. They were selected at random from Swedish population registers, where the inclusion criteria included never being hospitalized for schizophrenia or bipolar disorder, both parents born in Scandinavia and age

18 years or older<sup>16</sup>. To avoid a "supernormal" control sample, we chose not to exclude subjects hospitalized for other psychiatric reasons. <sup>16</sup>

We linked patients in SWEBIC and the control sample to longitudinal Swedish population-based registers using the unique personal identification number assigned to all persons living in Sweden<sup>17</sup>. Both SWEBIC and the control sample collection were approved by the Regional Ethical Review Board in Stockholm, Sweden, and all participants provided written informed consent.

# Phenotypic measures

Unemployment and long-term sick leave

To capture occupational functioning in bipolar disorder, we included measures of employment and long-term sick leave gathered from the longitudinal integrated database for health insurance and labour market studies (LISA) <sup>18</sup>. We included registrations between the ages 25 and 65 (retirement age in Sweden) and only considered individuals with at least 10 years of registrations during 1993–2015.

For measures of employment, we included yearly reports in October and November capturing if a subject held an employment that year<sup>18</sup>. Long-term sick leave or early retirement for each year was defined as having more than 60 full sick leave days *or* receiving any reimbursement for early retirement. The number of years without employment or with long-term sick leave were divided by the total number of years with registrations to calculate the percentage of years without employment or on sick leave. Since the raw data had both ceiling and floor effects, with majority of patients towards the end of the extremes 0% and 100%, we grouped the percentage of years without employed or on long-term sick leave into four categories (<25%, 25–50%, 50–75% and >75% of the years) as outlined in Table 1. Given previous indications that occupational functioning decrease by age in bipolar disorder<sup>2</sup>, we conducted

sensitivity analyses where we stratified occupational functioning by two age spans (25–39, and 40–65). A detailed description of the phenotype definitions is given in Supplementary methods.

Number of psychiatric hospital admissions

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The patient register has full coverage of psychiatric hospital admissions since 1973. We retrieved information about psychiatric hospital admissions from the Swedish National Patient Register 1973-01-01 to 2016-12-31. To estimate how many years a subject had the possibility of being included in the patient register, we counted the number of years between 1973-01-01, or from the date subjects turned 15 years old, until subjects turned 70 years old, death, or 2016-12-31. Subjects with fewer than 10 years of possible inclusion in the register were excluded. We used the years of possible inclusion in the patient register to calculate the mean number of psychiatric hospital admissions per year. Since hospital admissions has a skewed distribution, we grouped subjects into four categories (0, 0–0.1, >0.1–0.3, and >0.3 psychiatric hospital admissions per year) where the categories correspond to no admissions (average of zero admissions), up to one admission per decade (average of >0 to 0.1 admissions per year), between >1 to 3 admissions per decade (average of >0.1 to 0.3 admissions per year), and patients with more than 3 admissions per decade (average >0.3 admissions per year). The intensity of hospital admissions might differ by age. We therefore conducted sensitivity analyses where we used age at hospital admissions to stratify the intensity during two age spans (15–39 and 40–70 years of age).

#### Genotyping and polygenic scoring

DNA extraction from whole blood samples and genotyping in SWEBIC patients<sup>19</sup> and the controls<sup>16</sup> has previously been described. In short, genotyping of patients and controls was done using three genotyping arrays: Affymetrix 6.0 chips (Affymetrix, Santa Clara, CA,

USA), Illumina OmniExpress chips (Illumina, San Diego, CA, USA), and Infinium 150 PsychArray-24 v1.2 BeadChip (Illumina, San Diego, CA, USA). Quality control was done 151 using the Ricopili pipeline<sup>20</sup> and genotypes were imputed to the HRC 1.1 reference panel 152 using the Sanger imputation server<sup>21</sup>. 153 We used publicly available summary statistics from genome-wide association studies to 154 calculate PGS for psychiatric disorders and educational attainment in our study sample. These 155 156 PGS were chosen to capture disorders on the psychotic/affective spectrum (bipolar disorder, schizophrenia, and MDD), a commonly comorbid disorder (ADHD), and cognitive abilities 157 158 (educational attainment). Summary statistics for bipolar disorder (bip2021, excluding our SWEBIC sample)<sup>22</sup> and schizophrenia (scz2022, excluding all Swedish samples)<sup>23</sup> were 159 received from Psychiatric Genomics Consortium (PGC). We also used publicly available 160 summary statistics for MDD (mdd2018, excluding 23andMe) <sup>24</sup>, ADHD (adhd2019) <sup>25</sup>, 161 educational attainment<sup>26</sup> and Alcohol Use Disorders Identification Test (AUDIT, excluding 162 23andMe)  $^{27}$ . To minimize multiple testing, we only tested the P-value thresholds ( $P_T$ ) that 163 164 previously shown to best predict their respective phenotypes: P<sub>T</sub><0.1 for bipolar disorder, 165 schizophrenia, ADHD and AUDIT; P<sub>T</sub><0.5 for MDD; and P<sub>T</sub><0.05 for educational attainment. 166 We only included variants with INFO score >0.9 for all PGSs (INFO not available for 167 educational attainment). 168 We calculated PGSs within each genotyping wave separately using hard call genotypes. We 169 included autosomal single nucleotide polymorphisms (SNPs) with INFO score>0.9, MAF>0.01, strand unambiguous SNPs, >98% genotype success rate, and not in the extended 170 MHC region (chr6:25-34 Mb, hg19). To exclude variants in linkage disequilibrium (LD), we 171 clumped based on  $r^2 < 0.2$  within a 500 kb window using the SWEBIC cohort as LD reference. 172 PRSice (v. 2.3.3) <sup>28</sup> was used for LD clumping and polygenic scoring. PGSs standardized ([x 173

- mean] / SD) within each genotyping wave, including both bipolar disorder cases and controls, were used for statistical analyses.

#### Statistical analyses

We conducted ordinal regression for association analyses between the included phenotypes and PGSs using R (v. 4.0.2). We adjusted regression analyses for sex, the first six ancestry principal components, genotyping wave, and year of birth grouped into 6 categories (<1955, 1955–62, 1963–69, 1970–77, 1978–88, >1988). We used Bonferroni correction for multiple testing: correcting for six PGS and three phenotypes (18 tests) yielded a P-value threshold for significance  $P_T$ <0.0028. To correct for possible influence of age of onset, we included information about age at first symptom before and after age 24 (see description in supplementary methods). Results after correcting for age of first symptom are presented in Table S1 and Table S2.

# **BDRN** replication cohort

The replication cohort included bipolar disorder patients from the UK Bipolar Disorder Research Network (BDRN) study<sup>29,30</sup>. Information about number of psychiatric hospital admissions and age at interview was available for 4,219 subjects with DSM-IV bipolar disorder type 1 (N=2,906), bipolar disorder type 2 (N=1,239), or bipolar disorder NOS (N=74). All the subjects used in the analysis have at least 10 years history of disease. Hospital admissions per year was calculated by dividing number of hospital admissions by the number of years between age 15 and age at interview. Number of admissions per year was grouped into four categories as in the main analysis. Participants were genotyped on Affymetrix GeneChip 500K Mapping Array Set, Illumina Omni Express Array, and Illumina PsychChip. For each array strict quality control (QC) was performed separately. QC used PLINK 1.9<sup>31</sup> software excluding SNPs with MAF<0.01, 

deviation from Hardy-Weinberg Equilibrium (HWE) at  $P \le 10$ -6, call rate  $\le 98\%$ . Individuals were excluded from the sample if they had increased or decreased heterozygosity of |F| > 0.1, a discrepancy between their genotypic and reported sex, genotype call rate < 98%, high pairwise relatedness (pi-hat > 0.2) or did not cluster with European population samples in principal component analysis of 2000 participants from 19 populations of the 1000 Genomes Project. After OC, data for each platform were phased using SHAPEIT version 3.4.0.1023<sup>32</sup> and imputed using IMPUTE2<sup>33</sup> with the 1000 Genomes Project reference panel (phase 3). Imputed genotype dosages were converted to the most probable genotypes (probability  $\geq 0.9$ ) with additional SNPs excluded if the imputation INFO score was <0.8, MAF<0.01 or HWEP<1e-6). Imputed data were then merged on common SNPs between platforms. The polygenic scoring in BDRN were based on the same summary statistics as in the main analyses. However, summary statistics for schizophrenia and bipolar disorder were based on results excluding overlapping subjects from the BDRN cohort. PGS generation used PLINK version 1.9<sup>31</sup> in PRSice<sup>34</sup>. Imputed genotypes were clumped for linkage disequilibrium (window, 250 kb; r2 = 0.1) and single-nucleotide polymorphisms most significantly associated with the different traits were retained. After clumping, PRSs were generated at different P value thresholds (P<sub>T</sub>) p < 1.00, p  $\leq$  .50, p  $\leq$  .20, p  $\leq$  .10, p  $\leq$  .05, p  $\leq$  .01, and p  $\leq$  .001 and converted to z scores. For each of the PGS, we included the same p-value thresholds in the statistical analyses as described for the main analyses in the SWEBIC cohort. To test the association between hospital admissions and the PGS, we conducted ordinal logistic regression analyses adjusting for ten ancestry principal components, genotyping platform, and sex using R.

#### **RESULTS**

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Sample characteristics

A total of 4,782 bipolar disorder cases with phenotypic data from the SWEBIC population and 2,963 control subjects were included for association analyses between PGSs for selected traits and number of psychiatric hospital admissions per year and occupational functioning (unemployment and long-term sick leave) (Table 1). The number of years with information regarding employment and sick leave ranged between 10 and 23 years with a mean of 20 years for in total 4,139 patients and 2,963 controls. The mean percentage of years with registered long-term sick leave was 49% in cases and 13% in controls, while the mean percentage of years registered as unemployed were 40% in cases and 15% in controls. For sensitivity analyses, measures of employment and sick leave in bipolar disorder patients were also included during two age spans (age 25–39, N=1,649; age 40–65, N=2,999). Information about psychiatric hospital admissions intensity between 1973 and 2016 were available for 4,782 bipolar disorder patients and 3,339 controls between ages 15 and 70. The mean years of possible inclusion in the national patient register in the patient group was 36 years (range 10–44 years) with a mean of 0.2 psychiatric hospital admissions per year, which corresponds to an average of two hospital admissions per decade. For sensitivity analyses, number of hospital admissions per year was analysed in two age spans (age 15–39, N=4,180; age 40–70, N=3,213), where 2,611 patients had at least 10 years hospital admission intensity measurements during both age spans. The control group had an average of 39 years' (range 12–44 years) worth of data on hospital admissions available; 134 control subjects had one or more psychiatric hospitalization for other indications than bipolar disorder or schizophrenia. Polygenic score analyses We tested the association between occupational functioning as well as psychiatric hospital admissions and PGS for four psychiatric disorders (bipolar disorder, schizophrenia, ADHD,

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and MDD), AUDIT, and educational attainment in bipolar disorder. This was followed up by

- 246 conducting association analyses between PGS and occupational functioning in a control
- sample, as well as psychiatric hospitalisations in the patient cohort BDRN.
- In bipolar disorder cases, Figure 1 shows that the number of hospital admissions per year was
- significantly associated with higher PGS for bipolar disorder (OR=1.14, P=2.4x10<sup>-6</sup>) and
- schizophrenia (OR=1.23,  $P=6.5\times10^{-12}$ ), but not significantly associated with PGS for ADHD,
- MDD, or educational attainment. By contrast, unemployment and long-term sick leave were
- significantly associated with higher PGS for MDD (unemployment: OR=1.17,  $P=1.9\times10^{-7}$ ;
- sick leave: OR=1.20,  $P=3.0x10^{-10}$ ), schizophrenia (unemployment: OR=1.16,  $P=4.1x10^{-6}$ ),
- and ADHD (unemployment: OR=1.13, P=3.9 x 10<sup>-5</sup>; sick leave: OR=1.12, P=6.4x10<sup>-5</sup>), as
- well as with lower PGS for educational attainment (unemployment: OR=0.83, *P*=7.6x10<sup>-11</sup>;
- sick leave: OR=0.85,  $P=3.0 \times 10^{-9}$ ). PGS for bipolar disorder and AUDIT-PGS were not
- associated with unemployment or long-term sick leave. The results remained after correcting
- 258 for age of first symptom (Table S1).
- 259 In controls, both measures of occupational functioning were associated with PGS for MDD,
- ADHD, and educational attainment in the same direction as in bipolar patients (P<0.05). PGS
- 261 for bipolar disorder and schizophrenia were not associated with occupational functioning in
- 262 control subjects (Figure 1).
- In the BDRN cohort, we could replicate the association between psychiatric hospitalizations
- and PGS for bipolar disorder (OR=1.17, P=3.3 x10<sup>-8</sup>) and schizophrenia (OR=1.19,
- $P=2.5\times10^{-8}$ ; Table 3). In BDRN, a lower PGS for ADHD was also associated with increased
- 266 psychiatric hospitalizations (OR=0.91, *P*=0.0013).
- In Figure 1, we present the main association analyses separately for bipolar disorder type 1,
- 268 type 2, and NOS. For bipolar disorder type 1, the occupational phenotypes were significantly
- associated with PGS in the same direction as the whole group, whereas hospitalizations only

associated with SCZ-PGS (OR=1.17, P=6.9x10<sup>-4</sup>). With respect to bipolar disorder type 2, we 270 271 observe consistent trends with the overall sample between PGS and phenotypes. However, only certain PGS showed significant associations (P<0.05) with employment (ADHD-PGS 272 273 and EA-PGS), sick leave (MDD-PGS, ADHD-PGS, and EA-PGS), and hospitalizations (MDD-PGS and SCZ-PGS). Within the NOS category, we found similar results to the whole 274 group analysis with regards to employment, where the same PGS associated at P<0.05 (SCZ-275 276 PGS, MDD-PGS, ADHD-PGS, and EA-PGS). However, only MDD-PGS associated with sick leave and there were no significant associations at P<0.05 with hospitalizations. In 277 bipolar subtype specific analyses in BDRN, bipolar disorder-PGS, SCZ-PGS, and ADHD-278 279 PGS were associated (P<0.05) with hospitalizations in bipolar disorder type 1 (Table 3). Occupational dysfunction and hospital admission intensity might differ depending on age<sup>2</sup>. 280 We conducted sensitivity analyses stratified by intensities during the time before and after age 281 40 (Table S2). In the older age group, the effect sizes for the PGS associations with long-term 282 sick leave and employment resembled the main analyses and were generally larger (Table 283 284 S2). In the younger age group, however, not all PGS were significantly associated with occupational outcomes: Unemployment associated with SCZ-PGS (OR=1.11, P=0.043), 285 ADHD-PGS (OR=1.13, P=0.012), and EA-PGS (OR=0.86, P=6.5x10<sup>-4</sup>). Long-term sick 286 287 leave associated with MDD-PGS (OR=1.04, P=0.015), ADHD-PGS (OR=1.11, P=0.034), and EA-PGS (OR=0.85, P=4.7x10<sup>-4</sup>). Further, in the older age group hospital admission 288 intensity was positively correlated with MDD-PGS (OR=1.12, P=6.3x10<sup>-4</sup>) and 289 290 schizophrenia-PGS (OR=1.16, P=4.1x10<sup>-5</sup>), but negatively correlated with PGS for educational attainment (OR=0.89, P=4.3x10<sup>-4</sup>). In the younger age group, hospital admission 291 intensity was associated with PGS for bipolar disorder (OR=1.20, P=3.2x10<sup>-9</sup>), schizophrenia 292  $(OR=1.23, P=2.9\times10^{-10})$ , and educational attainment  $(OR=1.12, P=4.9\times10^{-5})$ . To test if these 293 results were driven by age of onset, we corrected the regression analyses for age at first 294

symptom (before/after age 24) and the main results remained associated after correction as seen in table S2.

#### **DISCUSSION**

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Hospitalization due to severe acute mood episodes and long-term occupational dysfunction are both serious consequences of bipolar disorder. Here we demonstrate that these key outcomes are differentially related to genetic factors: Psychiatric hospitalizations are associated with polygenic risk for bipolar disorder and schizophrenia, but not MDD, ADHD, or educational attainment. We replicate the association between hospitalizations and PGS for schizophrenia and bipolar disorder in the BDRN cohort. By contrast, occupational dysfunction correlated with polygenic scores for educational attainment, ADHD, MDD, and schizophrenia, but showed no association with polygenic risk for bipolar disorder. Interestingly, similar polygenic liabilities associated with occupational dysfunction in controls, suggesting that the polygenic links to occupational dysfunction demonstrated here are not specific for persons with bipolar disorder. The observation that polygenic risk for bipolar disorder is associated with psychiatric hospitalizations accords previous findings<sup>9,10</sup>. This association could be expected as most patients in the PGC training set for the PGS have bipolar disorder type 1, who typically have a higher number of hospitalizations than type 2. The polygenic risks associated with unemployment and long-term sick leave are novel findings that have bearing on clinical strategies for improving outcomes in bipolar disorder. The management of bipolar disorder primarily revolves around prevention of mood episodes with the assumption that patients will functionally recover with subsiding symptoms. But many patients remain functionally impaired despite best available treatment, and deficits in psychosocial functioning in bipolar disorder may persist long after symptom recovery<sup>35</sup>. The high rates of occupational

dysfunction not only affect patients' quality of life and socioeconomic status<sup>2</sup>, but is also a 319 major driver of the societal costs caused by the disorder<sup>5</sup>. 320 Polygenic risk for ADHD has previously been associated with rapid cycling<sup>36</sup> and age of 321 onset<sup>37</sup> in bipolar disorder. Our finding that polygenic scores for ADHD and educational 322 323 attainment are also associated with occupational dysfunction echoes the clinical observations that comorbid ADHD<sup>14</sup>, years of education, and cognitive performance<sup>2,7,38</sup> are strong 324 325 predictors of occupational functioning. Our result for educational attainment PGS in controls accords the positive correlation between education and occupational functioning<sup>39</sup>. Clinical 326 327 studies have also found that comorbid ADHD is associated with worse clinical outcome in 328 bipolar disorder, including earlier onset and higher frequency of depressive—but not manic episodes<sup>14</sup>. Given that hospitalization is more common for manic than depressive episodes, 329 this might partly explain why lower ADHD-PGS was associated with hospitalizations in the 330 replication BDRN cohort. Taken together, our findings extend these clinical observations by 331 demonstrating that not only a clinical ADHD diagnosis, but also higher polygenic risk for 332 333 ADHD, increase the risk for long-term impairment. Together these results suggest that a developmental perspective is needed to improve long-term outcome also in bipolar disorder<sup>40</sup>. 334 Occupational outcomes in bipolar disorder patients might be improved by addressing other 335 336 domains than core mood symptoms including ADHD symptoms—potentially also subsyndromal manifestations—and cognitive dysfunction. 337 Polygenic risk for MDD has previously been associated with rapid cycling<sup>36</sup> and suicide 338 attempts<sup>41</sup> in bipolar disorder. Here, we add occupational dysfunction to the list of negative 339 consequences of high polygenic liability for MDD. 340 Polygenic risk for schizophrenia was the only trait in our study that significantly associated 341 with both occupational functioning and hospital admissions in persons with bipolar disorder. 342 Although a genetic overlap between schizophrenia and bipolar disorder has been 343

demonstrated<sup>10</sup>, these findings show that polygenic liability for bipolar disorder and schizophrenia differ with respect to functional outcomes. Consistent with our findings on occupational functioning, schizophrenia risk alleles have previously been associated with poorer cognitive performance, whereas bipolar disorder risk alleles have been associated with better cognitive performance<sup>42</sup>. To account for the potentially increasing risk of occupational dysfunction with higher age<sup>2</sup>, we stratified analyses by intensities before and after age 40. The effect sizes for PGSs associated with long-term sick leave in the main analyses (SCZ, MDD, ADHD, and educational attainment) were slightly larger in the older age group. Conversely, the association between hospital admissions and bipolar disorder-PGS was limited to the younger age group (age 15-39). In addition, educational attainment-PGS exhibited associations with hospital admissions in opposite directions within the two age groups. This suggest that our outcomes capture different traits depending on age. The intensity of hospital admissions appears to relate to more frequent severe mood episodes in the younger age group, but may be confounded by treatment response or cognitive ability to cope with the disorder in older age. Strengths of this study include that we capture complete data on hospital care, employment, and sick leave during at least 10 years for each subject in a large sample of bipolar disorder patients. There are also some limitations to consider. First, register data do not include reasons for being without employment. In fact, one reason could be due to university studies, which is why we restricted the inclusion to subjects at least 25 years of age as most students have finished their education by the age of 25. Further, the frequency of long-term sick leave and early retirement are to some extent influenced by changes in social security policy during the study period. This might influence the age-specific analyses. Moreover, as we do not have information on the specific reason for sick-leave or early retirement, some of these recordings might reflect sick leave for reasons other than bipolar disorder. Second, although we observe

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and replicate associations between PGSs and important bipolar disorder phenotypes, the variance explained is still insufficient to be considered for clinical use in personalized psychiatry. Finally, some of our findings in the sensitivity analyses could be chance findings that would need to be followed up in future studies.

In conclusion, our results suggest that the underlying polygenetic liability for the severity of bipolar illness in terms of psychiatric hospitalisations clearly differs from the occupational outcome. This is important information given that much of the lifetime burden of bipolar disorder is driven by prolonged periods of sick leave and/or unemployment rather than acute mood episodes. Even if clinical symptoms are improved, psychosocial impairments may persist<sup>43</sup>, which suggests that interventions other than those used to prevent mood episodes might be required to mitigate functional impairment. This might include addressing comorbid symptoms (e.g., ADHD) or cognitive dysfunction through cognitive remediation therapy. Clarifying the underlying mechanisms of poor outcome is an important step towards targeted treatment interventions, and might also be used for prognostic purposes.

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# FIGURE LEGEND

Figure 1. Results from polygenic score association analyses of occupational functionin					
and psychiatric hospital admissions in bipolar disorder and controls. *Bonferroni					
corrected $P$ -value threshold for significance $P < 0.0028$ (correcting for 18 tests). OR=odds					
ratio.					

# **TABLES**

**Table 1.** Description of the SWEBIC bipolar disorder cohort and controls for the included phenotypes.

	SWEBIC		Controls		
	BD (all)	BD-I / BD-II / NOS			
	(N=4,782)	(N=2,175 / N=1,638 / N=969)	(N=3,339)		
Phenotype					
Sex (% females)	2,986 (62%)	1,253 (58%) / 910 (66%) / 630 (65%)	1,667 (50%)		
Percent of years w	Percent of years without employment (1993–2015, age 25-65) <sup>a</sup>				
	N=4,139	N=1,905 / N=1,389 / N=845	N = 2,963		
<25 %	1,686 (44%)	769 (40%) / 644 (46%) / 353 (42%)	2,319 (78%)		
25-50%	739 (19%)	373 (20%) / 320 (23%) / 189 (22%)	383 (13%)		
50-75%	585 (15%)	298 (16%) / 225 (16%) / 148 (18%)	130 (4%)		
>75%	813 (21%)	465 (24%) / 200 (14%) / 155 (18%)	131 (4%)		
Percent of years w					
	N = 4,139	N=1,905 / N=1,389 / N=845	N = 2,963		
<25%	1,159 (30%)	588 (31%) / 498 (36%) / 248 (29%)	2,483 (84%)		
25-50%	647 (17%)	295 (15%) / 296 (21%) / 157 (19%)	150 (5%)		
50-75%	847 (23%)	404 (21%) / 344 (25%) / 207 (24%)	184 (6%)		
>75%	1,170 (31%)	618 (32%) / 251 (18%) / 233 (28%)	146 (5%)		
Number of hospita	al admissions pe	er year (1973–2016, age 15-70) <sup>b</sup>			
	N = 4,782	N=2,175 / N=1,638 / N=969	N = 3,339		
0	949 (20%)	164 (7%) / 566 (35%) / 219 (24%)	3,205 (96%)		
0-0.1	1,497 (31%)	633 (29%) / 526 (32%) / 338 (35%)	114 (3%)		
0.1-0.3	1,434 (30%)	803 (37%) / 370 (23%) / 261 (27%)	16 (<1%)		
>0.3	902 (19%)	575 (26%) / 176 (11%) / 151 (16%)	4 (<1%)		
			1		

<sup>&</sup>lt;sup>a</sup>Mean years with phenotype information: 20 years (10-23 years). <sup>b</sup>Mean years with phenotype

<sup>517</sup> information: 36 years (10-44 years).

# **Table 2.** Description of the bipolar disorder replication cohort Bipolar Disorder Research Network (BDRN).

Phenotype	BD (all)	BD-I / BD-II					
	N=4,219	N=2,906 / N=1,239					
Sex (% females)	2,859 (68%)	1,971 (68%) / 841 (68%)					
Number of hospital admissions per year (age 15-70)							
0	1091 (26%)	340 (12%) / 713 (58%)					
0-0.1	1446 (34%)	1101 (38%) / 324 (26%)					
0.1-0.3	1229 (29%)	1077 (37%) / 142 (11%)					
>0.3	320 (8%)	286 (10%) / 33 (3%)					

- Table 3. Replication association analyses between PGS and psychiatric hospital admissions in
- 522 the BDRN cohort (N=4,219).

	ALL (N=4,219)		BD-I (N=2,906)		BD-II (N=1,239)	
PGS	OR (95% CI)	Р	OR (95% CI)	Р	OR (95% CI)	Р
BD-PGS	1.17 (1.11-1.24)	3.3 x 10 <sup>-8</sup>	1.15 (1.08-1.24)	6.6x10 <sup>-5</sup>	0.92 (0.82-1.04)	0.17
SCZ-PGS	1.19 (1.12-1.27)	2.5 x 10 <sup>-8</sup>	1.19 (1.10-1.28)	9.9x10 <sup>-6</sup>	0.90 (0.79-1.02)	0.092
MDD-PGS	0.96 (0.91-1.02)	0.16	0.99 (0.92-1.06)	0.71	0.97 (0.87-1.09)	0.59
ADHD-PGS	0.91 (0.86-0.96)	0.0013	0.91 (0.85-0.98)	0.011	1.02 (0.91-1.15)	0.68
EA-PGS	0.97 (0.92-1.03)	0.38	0.99 (0.93-1.07)	0.85	0.95 (0.84-1.07)	0.40
AUDIT-PGS	0.98 (0.92-1.03)	0.41	1.01 (0.95-1.09)	0.71	1.00 (0.89-1.13)	0.99

BD=Bipolar disorder, SCZ=Schizophrenia, MDD=Major Depressive Disorder, ADHD=Attention Deficit Hyperactivity Disorder, EA=educational attainment, AUDIT=Alcohol Use Disorders Identification Test.