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Citation for final published version:

Borodich Suarez, Sofia, Heravi, Saeed and Pepelyshev, Andrey 2023. Forecasting industrial production indices with a new singular spectrum analysis forecasting algorithm. *Statistics and Its Interface* 16 (1) , pp. 31-42. 10.4310/21-SII693

Publishers page: <https://dx.doi.org/10.4310/21-SII693>

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Forecasting industrial production indices with a new singular spectrum analysis forecasting algorithm

SOFIA BORODICH SUAREZ, SAEED HERAVI, AND ANDREY PEPELYSHEV*

Existing time series analysis and forecasting approaches struggle to produce accurate results in application to time series with complex trend, such as those commonly displayed by indices of industrial production (IIPs). In this study, a new version of the Singular Spectrum Analysis (SSA) technique is developed, namely the Separate Trend and Seasonality (SSA-STS) forecasting algorithm. Its performance is compared to those of benchmark, classical times series forecasting methods, including Basic SSA (the core version of SSA), ARIMA, Exponential Smoothing (ETS) and Neural Network (NN). The methods in this study are applied to both simulated and real data. The latter includes twenty four monthly series of seasonally unadjusted IIPs of various sectors for the UK, Germany and France. Using the out-of-sample forecasts, the results of this newly developed SSA-STS algorithm were compared to the other aforementioned forecasting schemes by the means of pooled Root-Mean-Square-Error (RMSE). The pooling is done based on the number of steps ahead the forecasts extend, allowing for the performance of the methods to be evaluated on short and long horizons. The Kolmogorov-Smirnov Predictive Accuracy (KSPA) statistical test is applied to certify whether the errors produced by SSA-STS are statistically significantly smaller than those of all the benchmark methods. Since this new technique is based on separate trend and seasonality forecasting, it overcomes the difficulties in forecasting series with complex trends and seasonality, thus demonstrating a clear advantage over other methods in such particular cases.

AMS 2000 SUBJECT CLASSIFICATIONS: 62M20.

KEYWORDS AND PHRASES: Singular spectrum analysis, Forecasting, Root mean square error.

1. INTRODUCTION

Time series are thought to be composed of a combination of a trend, a periodic (referred to as seasonality if the frequency of oscillations follows a seasonal/monthly schedule), and random residuals. Components are attributed to these groups based on their frequencies. However, identifying these components correctly for real life time series (made

up of ordered observations over time), that have not been generated from a linear recurrence relation (LRR), is not as evident as it may seem. Hence, the purpose of the Singular Spectrum Analysis (SSA) technique is to decompose time series into a sum of components with simple interpretable shapes and create a grouping that may be utilized for further actions, such as forecasting. There are many different modifications to the original Basic SSA methodology [12, 11]. Implementing the appropriate method, based on the particularities displayed by the time series under consideration, allows for a tailored approach, leading to better results.

A popular application for SSA are indices of industrial production (IIPs). These series are especially interesting, not only from an economic point of view, as they can be used as a proxy to evaluate the current state of the economy by policymakers, banks and economists in general, but also due to their particular structure. They often display complex trends with strong seasonal sub-cycles. This means that standard forecasting methods, such as ARIMA, struggle to accurately predict them, whereas SSA has an advantage due to its non-parametric nature. As shown in [16, 17], and further in this study, Basic SSA is simple to implement and indeed outperforms previously applied methods. However, applying it to such series can result in a poor reconstruction, as it fails to extract the tail-end seasonality correctly and does not capture the complexity of such trends [14], and therefore SSA with derivatives, or DerivSSA, is preferred [13]. The DerivSSA method improves separability of components by changing their contributions based on their frequency. This conveniently groups components into those that constitute to the trend and the seasonality, hence facilitating the creation of an accurate reconstruction, which could reflect as an improvement in subsequent forecasts via SSA, since they rely heavily on the quality of the reconstruction.

However, this approach is not tailored for forecasting. One may still encounter a similar problem when creating forecasts directly from the reconstructed series, as using a method that evaluates the series as a whole will not capture the eccentricities of a complex trend and seasonality components, thus failing to predict their different behaviors. Hence, there is some room for improvement.

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1 Therefore, the present paper is devoted to the develop- 57
 2 ment of a new version of SSA-type forecasting algorithm to 58
 3 address the need for a highly flexible method of accurately 59
 4 forecasting time series with complex trends and strong sea- 60
 5 sonality, such as the ones commonly seen in industrial pro- 61
 6 duction data. The new SSA-STS forecasting algorithm is 62
 7 based on utilising the SSA technique to separately forecast 63
 8 the trend and the periodic component of a series with the 64
 9 required form, then the combination should provide an accu- 65
 10 rate forecasts for both long and short horizons. The greater 66
 11 flexibility stems from the fact that users can select four pa- 67
 12 rameters, as supposed to the two in the traditional SSA 68
 13 forecasting algorithms. 69

14 It is important to note that the conventional algorithms 70
 15 applied to the trend and seasonal component separately do 71
 16 not give benefits compared to the proposed algorithm. 72

17 This paper is organized as follows. In the following sec- 73
 18 tion, a preliminary review of the pertinent literature, nec- 74
 19 essary for the understanding of the research, is given. Sec- 75
 20 tion 3 details the new SSA-STS forecasting algorithm, as 76
 21 well as the ideas and methods used in its creation. In Sec- 77
 22 tion 4, a description of the data being utilised is provided. 78
 23 In Section 5, the algorithm is applied to simulated and real 79
 24 industrial production data, along with benchmark methods, 80
 25 and a comparison is made. Finally, Section 6 contains the 81
 26 conclusions drawn from the results of the research. 82
 27

28 2. DEVELOPMENTS OF SSA AND 83 29 RELEVANT LITERATURE 84

30 SSA is a novel and powerful non-parametric time series 85
 31 analysis and forecasting technique. This flexible methodol- 86
 32 ogy is based on the decomposition of the series into a sum 87
 33 of simple components, allowing for easy interpretation and 88
 34 facilitating the further construction of forecasts of periodic 89
 35 patterns in the presence of trends, see [12, 11]. It has been 90
 36 successfully employed in various areas, ranging from math- 91
 37 ematics and signal processing to meteorology, genetics, eco- 92
 38 nomics and tourism research [20, 27]. A full review of the 93
 39 application of SSA for economic and financial time series 94
 40 is presented in [19]. The main advantage of SSA is that it 95
 41 can be used without making statistical assumptions, such as 96
 42 normality and stationarity of the data. 97

43 The evolution of ideas leading to SSA as we know it, 98
 44 started in the 18th century with a chain of key advance- 99
 45 ments in signal processing and time series analysis [8]. Some 100
 46 notable examples are, the spectral decomposition of time se- 101
 47 ries [23, 22] and the embedding theorem [24, 29]. In general, 102
 48 [4] and [9] are regarded as the initial, official publications of 103
 49 SSA, although many findings in previous papers contributed 104
 50 to this development, and their authors may also be credited 105
 51 [6, 1, 32]. The work of Ghil and Vautard throughout the late 106
 52 80s and early 90s further established the theory of SSA and 107
 53 its applications in climatology, [30, 31]. 108
 54 109
 55 110
 56 111
 112

A completely separate approach and chain of develop- 57
 ments took place in Russia during this time, culminating in 58
 a variant of SSA known as Caterpillar. This peculiar name 59
 originated from the analogy between a caterpillar's move- 60
 ment and the moving window utilised in this analysis. Al- 61
 though this was only publicised after the fall of the Soviet 62
 Union in [7], the precedent historical influences indicate that 63
 consistent work was carried out throughout the 70s, 80s and 64
 90s (see, e.g. [3, 2, 5]). 65

In the following years up until the present, major devel- 66
 opments on the methodology of SSA were achieved, detailed 67
 in [12, 14, 11]. Including, the development of the Rssa pack- 68
 age in programming language R for fast implementation of 69
 the techniques has also widened the application of SSA to a 70
 variety of problems. Further modifications of the SSA tech- 71
 nique have been steadily researched and introduced to cope 72
 with different problems encountered, for example the SSA- 73
 AMUSE method [15], which again facilitates the separation 74
 of components. 75

Some of the most widely studied and forecasted time se- 76
 ries are IIPs, to which SSA and the benchmark methods 77
 used in this study, among others, have been applied. In- 78
 dustrial production data is commonly used by policymakers 79
 and economists as it is an important contributing factor to 80
 the gross domestic product (GDP) as a whole. Hence, great 81
 importance is given to the accuracy of IIP forecasts, since 82
 the ability to correctly predict their movement can prove to 83
 be highly beneficial. Extensive research has been carried out 84
 into the optimal forecasting method for IIPs. 85

For instance, the seasonality patterns of seasonally un- 86
 adjusted series for eight components of real industrial pro- 87
 duction in Germany, France and the UK are investigated in 88
 [25]. Their findings show that seasonality accounts for over 89
 80% of the variations in all series for the UK and Germany, 90
 except for vehicles. They also found stronger seasonality in 91
 France, due to declines in production in the summer in tra- 92
 ditional industrial sectors. 93

Using the same data, the accuracy of NN forecasts was 94
 studied in [21] and it was concluded that, in general, NN 95
 models dominate linear models in the ability to predict the 96
 direction of change, but not in actual forecasting perfor- 97
 mance. Again, with the same data, SSA and ARIMA fore- 98
 casts are compared in [16], where it was found that SSA 99
 produces more accurate forecasts than the other benchmark 100
 models and performs very well in predicting the direction of 101
 change. 102

In [26] vector ETS was applied to forecast several series, 103
 including that of aggregate industrial production and it was 104
 shown that the forecasting power quickly deteriorates as the 105
 horizon extends. This is unsurprising, since ETS relies on the 106
 recent history of the series, hence the error of this method 107
 increases with the horizon. 108

The proposed SSA-STS forecasting algorithm grasps 109
 ideas from the literature mentioned above, in particular from 110
 111
 112

1 the works of [13, 14, 11] and extends upon it. Hence, con- 57
 2 tributing a method for forecasting specific series, with 58
 3 complex trends and strong seasonalities, by considering these 59
 4 components separately throughout the reconstruction and 60
 5 forecasting steps, allowing for different parameters to be 61
 6 utilised, thus improving the precision of results. This is par- 62
 7 ticularly applicable to the forecasting of industrial produc- 63
 8 tion series, which typically display such characteristics. 64

9 We apply the new forecasting algorithm to the 24 time 65
 10 series of seasonally unadjusted monthly indices of industrial 66
 11 production (IIP) in Germany, France and the UK, and pro- 67
 12 ceed to compare its performance against other benchmark 68
 13 methods mentioned in this review. The results are presented 69
 14 in a way that the forecasting accuracy can be evaluated for 70
 15 short and long horizons. 71

17 3. METHODOLOGY 72

18 The SSA methodology consists of a family of methods, 73
 19 such that the most suitable method can be chosen, depend- 74
 20 ing on the structure of a given time series and researchers' 75
 21 needs. The general scheme of SSA consists of four steps: em- 76
 22 bedding, decomposition, grouping and reconstruction. The 77
 23 core version of the SSA methodology is called Basic SSA, 78
 24 which is briefly described below, along with other SSA mod- 79
 25 ifications, including the new proposed SSA-STS forecasting 80
 26 algorithm. 81

28 3.1 Basic SSA 82

29 Let x_1, \dots, x_T be a time series of length T . Given a win- 83
 30 dow length L ($1 < L < T$), we construct the L -lagged vectors 84
 31 $X_i = (x_i, \dots, x_{i+L-1})^T$, $i = 1, 2, \dots, K = T - L + 1$, and com- 85
 32 pose these vectors into the matrix $\mathbf{X} = (x_{i+j-1})_{i,j=1}^{L,K} = [X_1 : 86$
 33 $\dots : X_K]$. \mathbf{X} is a Hankel matrix, meaning all the elements 87
 34 along its diagonal $i+j=\text{constant}$ are equal. 88

35 The columns X_j of \mathbf{X} can be considered as vectors be- 89
 36 longing to the L -dimensional space \mathbb{R}^L . The singular-value 90
 37 decomposition (SVD) of the matrix $\mathbf{X}\mathbf{X}^T$ yields a collec- 91
 38 tion of L eigenvalues and eigenvectors. For a given inte- 92
 39 ger r , $r < L$, we choose the r largest eigenvalues and corre- 93
 40 sponding eigenvectors of $\mathbf{X}\mathbf{X}^T$. These r components of the 94
 41 SVD decomposition can be separated into several groups, 95
 42 e.g. a trend group and a periodic group. The chosen eigen- 96
 43 vectors determine an r -dimensional subspace in \mathbb{R}^L ; call this 97
 44 subspace S_r . The L -dimensional data $\{X_1, \dots, X_K\}$ is then 98
 45 projected onto this r -dimensional subspace S_r and the sub- 99
 46 sequent averaging over the diagonals gives us some Han- 100
 47 kel matrix $\tilde{\mathbf{X}}$, which we consider as an SSA reconstruction 101
 48 of \mathbf{X} . The time series corresponding to $\tilde{\mathbf{X}}$ is called the re- 102
 49 constructed series and usually serves as an estimator of the 103
 50 signal when the observed time series is noisy. 104

52 3.2 SSA forecasting 105

53 There are two ways of constructing forecasts based on 106
 54 the SSA decomposition of the series described above, see 107
 55 108
 56

57 [12, Ch 2]. The most obvious way is to use the linear recur- 58
 59 rent formula which the last terms of the series reconstructed 60
 61 from $\tilde{\mathbf{X}}$ satisfy. We however prefer to use the so-called 'SSA 61
 62 vector forecast' [12, Sect. 2.3.1]. The main idea of this fore- 62
 63 casting algorithm is as follows. A selection of r eigenvec- 63
 64 tors of $\mathbf{X}\mathbf{X}^T$ leads to the creation of the subspace S_r . SVD 64
 65 properties allow us to assert that the L -dimensional vec- 65
 66 tors $\{X_1, \dots, X_K\}$ lie close to this subspace. Consider the 66
 67 vectors Z_1, \dots, Z_K , where Z_i is defined as the projection of 67
 68 X_i onto the subspace S_r . The vector forecasting algorithm 68
 69 then sequentially constructs the vectors $\{Z_{K+1}, Z_{K+2}, \dots\}$ 69
 70 so that they stay in the chosen subspace S_r and the han- 70
 71 kelization of the matrix $(Z_1, \dots, Z_K, Z_{K+1}, Z_{K+2}, \dots)$ gives 71
 72 the vector forecast. 72

73 3.3 DerivSSA 74

75 The application of Basic SSA to a given time series may 76
 77 lead to a poor separation [especially, at the tail-ends] of in- 77
 78 dividual components in the SVD decomposition. To resolve 78
 79 this problem, DerivSSA was developed, which improves the 79
 80 separability of trend components from periodic ones, [11, 80
 81 Sect. 2.5]. Specifically, DerivSSA arranges the components 81
 82 based on their derivatives, such that periodic components 82
 83 with high frequencies are assigned a larger contribution and 83
 84 the trend components are placed to the end of the decompo- 84
 85 sition. The algorithm of DerivSSA is similar to that of Basic 85
 86 SSA, with the replacement of the matrix \mathbf{X} by the matrix 86
 87 $\mathbf{X}_D = [X_1 : \dots : X_K, X_2 - X_1 : \dots : X_K - X_{K-1}]$. 87
 88

89 3.4 SSA-STS forecasting algorithm 90

91 The new SSA-STS forecasting algorithm proposed in this 91
 92 study is based on separate trend and seasonality forecast- 92
 93 ing. Initially, perform DerivSSA on a time series with some 93
 94 choice of parameters L and r . The application of DerivSSA 94
 95 produces a list of components, where the first r_s components 95
 96 are grouped into the seasonal component and the following 96
 97 r_t are grouped into the trend component. Then, conduct- 97
 98 ing separate forecasts for the seasonal component, by Ba- 98
 99 sic SSA with some L_s and the same r_s , and for the trend 99
 100 component, by SSA with double centering with some L_t 100
 101 and r_t , [11, Sect. 2.3]. Finally, the sum of the two forecasts 101
 102 is taken as the final forecast of the given time series. The 102
 103 above description of the SSA-STS algorithm contains six pa- 103
 104 rameters L, r, L_t, r_s, L_s, r_s , which provides a lot of freedom 104
 105 and flexibility. For reducing the computational complexity 105
 106 of the search for optimal values of parameters, we introduce 106
 107 the R code of the SSA-STS algorithm with four parameters 107
 108 L_t, r_t, L_s, r_s as follows. 108
 109
 110
 111
 112

```

1
2   s=ssa(TimeSeries,Ls)
3       # Perform the SSA decomposition
4   deriv=fossa(s, nested.groups=c(1:(rs+rt)))
5       # Perform DerivSSA
6   rec=reconstruct(deriv,
7                   groups=list(seasonality=1:rs,
8                               trend=(rs+1):(rs+rt)))
9       #Obtain the trend and seasonal component
10  s_trend = ssa(rec$trend,Lt,
11               column.projector=1,row.projector=1)
12       # Perform trend decomposition by SSA
13       with projections
14  f_trend=forecast(s_trend,groups=list(1:Lt),
15                  method='vector',len=Nahead)
16       # Forecasting the trend
17  s_seas = ssa(rec$seasonality,Ls)
18       # Perform the SSA decomposition of the
19       seasonal component
20  f_seas=forecast(s_seas,groups=list(1:rs),
21                 method='vector',len=Nahead)
22       # Forecasting the seasonal component
23  ssa_sts_forecast=f_trend$mean+f_seas$mean
24       # the sum of forecasts of the trend and
25       the seasonal component

```

The individual steps of the SSA-STS algorithm are well-known, see e.g. [14, 11], however, to the best of our knowledge, such an algorithm has not been suggested earlier in the literature.

Our algorithm relies on the improved separability of components provided by DerivSSA to correctly extract the trend and periodic/seasonality. However, in the forecasting step, since there is no need to separate any further and sufficiently many components are used, Basic SSA can be applied without negative consequences. Note that, if a given time series has a relatively simple trend, Basic SSA yields a good decomposition, thus the forecasts produced by SSA-STS and Basic SSA are similar. We would expect the SSA-STS forecasts to be superior only in the presence of a complex trend in a time series.

3.5 Parameter choice

For successful application of SSA methodology, the choice of parameters L (window length) and r (number of components considered) is of utmost importance, as an appropriate selection can improve the separability of components of the SVD decomposition and further generate more accurate forecasts [10]. Hence, to compare the results of Basic SSA and SSA-STS forecasting algorithms it is critical that optimal parameters are selected for both methods. For the purpose of forecasting IIPs, it is a good idea to find parameters for both SSA-STS and Basic SSA algorithms on the basis of accuracy of retrospective forecasts, that is, cutting the time series at different points and minimizing the aggregated RMSE of residuals between the forecasts and the given time series.

3.6 Other benchmark methods

The benchmark forecasting algorithms to which SSA-STS is compared to in this paper are well-known methods which have been applied to a wide range of time series, including those of industrial production.

The traditional forecasting methods are autoregressive integrated moving average (ARIMA), exponential smoothing (ETS) and neural network (NN) models. ARIMA is a parametric method and makes several assumptions about the data (e.g. stationarity, normality etc.) which are often too stringent for the reality.

ETS is motivated by the idea that older observations have less forecasting power than newer ones, hence the weights used in the weighted sum for the prediction shrink exponentially with time. This works well for series without complex, changing trends or seasonality, however there is also a great risk of errors accumulating when using this technique.

In contrast, NN is a highly flexible machine learning algorithm which allows for non-linear relationships, however also requires a large amount of previously observed data to perform well. This, along with Basic SSA, may prove to be the best rivals to SSA-STS for forecasting IIPs, due to their adaptability.

4. THE DATA

Data availability statement: the data used in this study is taken from Eurostat, the official statistical agency of the European Community.

Eight major components of real industrial production in three European countries, France, Germany and the UK, are considered in this study. The time series data are seasonally unadjusted monthly indices in Food Products, Chemicals, Basic Metals, Fabricated Metals, Machinery, Electrical Machinery, Vehicles and Electricity/Gas (Utilities) industries. The selected industries are important and diverse, seen as the eight time series included account for at least half of total industrial production in each country. In all cases, the sample period ends in July 2019. However, the starting dates for data in the UK, Germany and France data are different and reflect the availability of consistent data from Eurostat. The same time series data, with smaller sample periods, which ended earlier, have been previously employed and studied by [25, 21, 16, 28]. Note that data after August 2019 is not included in the sample, since forecasting during the COVID-19 pandemic is an untractable task and would not provide fair comparisons.

Figure 1 presents the time series data used in this study. As can be seen from these graphs, the movements of all these series are dominated by seasonality. However, except Utilities (Electricity and Gas supply), they all have complex trend pattern with different periods of expansions and contractions, as well as a sharp decline in production in 2008/2009 due to the financial and sub-prime mortgage banking crisis.

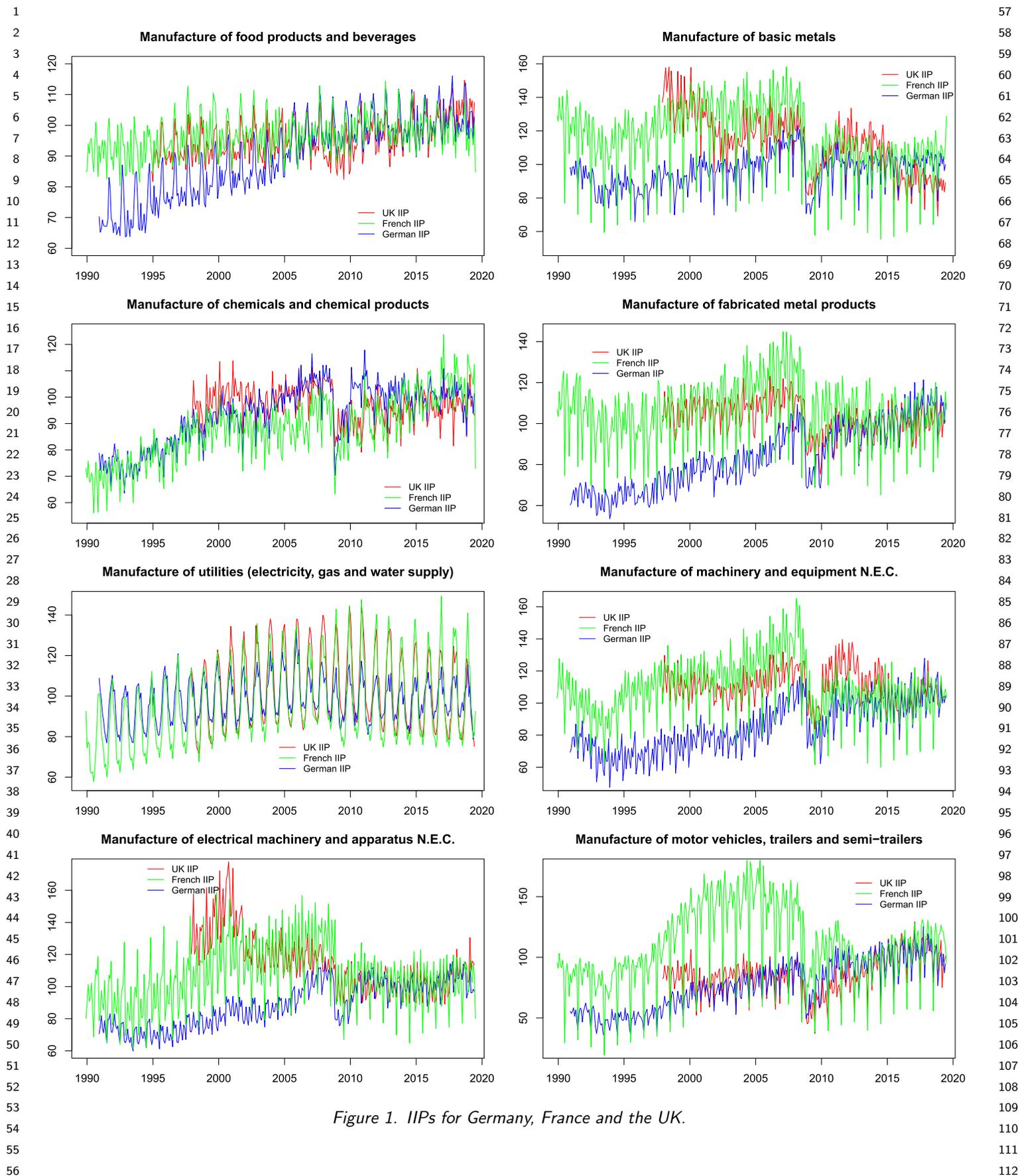


Figure 1. IIPs for Germany, France and the UK.

Table 1. Descriptive statistics of the production data

Series	Mean	Geom Mean	S.D.	R^2
UK				
Food	0.65	0.57	3.89	0.60
Basic	-2.45	-3.22	11.71	0.50
Metals				
Chemicals	0.13	-0.06	5.99	0.64
Fabricated				
Metal	-0.24	-0.49	6.88	0.64
Utilities	0.13	-0.02	5.46	0.84
Machinery	-0.46	-0.96	9.85	0.72
Electrical				
Equipment	-0.77	-1.08	7.78	0.76
Vehicles	0.84	-0.65	14.77	0.78
Germany				
Food	1.39	1.32	3.63	0.60
Basic	0.13	-0.49	10.38	0.71
Metals				
Chemicals	0.99	0.70	7.40	0.67
Fabricated				
Metal	1.82	1.42	8.58	0.58
Utilities	0.14	0.02	4.93	0.78
Machinery	1.12	0.59	9.82	0.82
Electrical				
Equipment	1.11	0.78	7.83	0.86
Vehicles	2.16	1.21	12.79	0.71
France				
Food	0.25	0.18	3.67	0.58
Basic	-0.69	-1.05	8.01	0.96
Metals				
Chemicals	1.55	1.34	6.31	0.75
Fabricated				
Metal	-0.18	-0.47	7.45	0.92
Utilities	1.03	0.86	5.93	0.81
Machinery	-0.28	-0.86	9.98	0.92
Electrical				
Equipment	0.34	0.04	7.40	0.93
Vehicles	1.01	-0.32	14.26	0.94

Mean, Geom Mean and S.D. are the arithmetic mean, geometric mean and standard deviation of the annual percentage change in the series over the sample period.

Table 1 contains information for each annual difference series and therefore refer to percentage changes in the original series. The first three columns are the sample arithmetic mean, geometric mean and standard deviation.

The sample means indicate the different growth/decline in production for these industrial sectors over the sample period, and it can be seen that some experienced a substantial rise during this period. In particular, production of Vehicles and Fabricated Metal in Germany show increase of 2.16% and 1.82% per year (based on arithmetic mean). However, others suffered declines, for instance, Basic Metal in the UK shows a sharp fall, with arithmetic average of 2.5% and geometric average of 3.2%.

The sample standard deviations indicate higher volatility

for all the three countries in production for the Vehicles industrial sector, with the lowest volatility in Food products. This is expected and aligns with the economic theory of consumer elasticity of demand.

The fourth column in Table 1 shows the seasonal coefficient of determination, which is obtained by regressing the monthly changes in production data against twelve monthly dummies, computed via a regression. The results generally show stronger seasonality for the French industrial production series than those for the corresponding series of Germany and the UK. In particular, monthly dummy variables account for over 90% of the variations in Basic Metals, Fabricated Metals, Machinery and Vehicles in France, which are associated with declines in production during the summer for the traditional industrial sectors. The results are in line with those reported in [25] for the data period ending December 1995.

5. APPLICATION

This section tests the new SSA-STS forecasting algorithm against well known forecasting methods, including Basic SSA, ARIMA, ETS and NN, on simulated and real data.

5.1 Numerical examples on simulated data

We now turn to the main purpose of this paper, evaluating the performance of forecasts. We use the out-of-sample forecast Root Mean Square Error (RMSE) to measure performance and support the findings via the statistical Kolmogorov-Smirnov Predictive Accuracy (KSPA) test [18].

First, we apply the SSA-STS forecasting algorithm to artificially generated data that emulates the movement of an IIP, to ensure that the algorithms performance is satisfactory and indeed produces exceedingly accurate results in regard to the benchmark methods it is being compared to. Our interest centers on the treatment of the trend in forecasting. The simulated series consists of the sum of an appropriate trend with differing complexity, periodic pattern and noise.

Consider an artificial time series mimicking UK industrial production data, with observations spanning from January 1998 to July 2019. More specifically, the time series $I\alpha$, which was generated as the sum of a trend, the seasonal component of the form $5.7 \cos(2\pi t/12) + 6.4 \cos(2\pi t/4 + 1.2)$ and the Gaussian white noise with variance of 1. The trend is defined by the vector of length 259, mirroring the number of observations in the IIP series for the UK, with the following form

$$100 + \alpha \text{ cumsum}(S_{-0.17,61}, S_{0.35,68}, S_{-3,11}, S_{1.6,25}, S_{-0.6,52}, S_{0.2,42}),$$

where $S_{z,k}$ is the vector of length k with all elements equal to z and ‘cumsum’ is the cumulative sum operator. Note that, the considered model of the trend has spline form, which is a very flexible model since the spline can nicely approximate

any nonlinear trend. The parameter α , ranging from 0.0 to 1.0, indicates the degree to which the trend differs from the constant value of 100. The higher the value of α , the more complex the trend will be.

As previously mentioned, in our study we compare the forecasts obtained by five methods: SSA-STS, Basic SSA, ARIMA, ETS and NN. To assess these methods, out-of-sample forecasts are computed for the last 60 observations. That is, we compare the forecasts by computing the aggregated RMSE of retrospective 12-month ahead forecasts with cutting points from July 2014 to July 2018. We re-estimate parameters of ARIMA, ETS and NN forecasts at each cutting point using the functions `auto.arima`, `ets` and `nnetar` from the R package `forecast`. However, the same parameters were selected for SSA-STS and Basic SSA forecasting algorithms in the post-sample period, these parameters are chosen optimally on the basis of retrospective forecast accuracy. They are shown in Table 2, as well as that, we also choose parameters $L_t = 12$ and $r_t = 3$ for the SSA-STS forecasting algorithm and maintain throughout.

The sensitivity analysis [which is not presented here due to lack of space, but is available upon request] shows the accuracy of SSA-STS and Basic SSA forecasts weakly depends on the parameters chosen for these methods.

Table 2. Parameters of the SSA-STS and Basic SSA forecasting algorithms for the artificial time series $I\alpha$ with $\alpha = 0.1, 0.2, 0.4, 0.6, 0.8, 1$

	SSA-STS				Basic SSA	
	L	r	L_s	r_s	L	r
I0.1	132	6	36	4	96	6
I0.2	36	7	84	4	120	6
I0.4	48	9	84	4	36	8
I0.6	48	9	60	4	36	8
I0.8	60	10	36	4	36	8
I1.0	60	10	48	4	36	8

The aggregated RMSE for these forecasts are presented in Table 3. We can see that for $\alpha = 0.1$ or 0.2 [i.e. series with relatively constant trends] the Basic SSA and SSA-STS forecasts perform similarly, both better than the rest of the other methods, for the whole range of step ahead forecasts considered. However, as α increases and the trend becomes more complex, the SSA-STS forecast develops an advantage. The RMSE values for the other methods clearly growing at a faster rate than those of the SSA-STS forecasts, hence for the case of $\alpha \geq 0.4$ the SSA-STS forecast has a noticeably smaller RMSE, indicating that it is indeed more accurate than the other forecasting methods. However, a contender for best forecasting results is the ETS approach, for shorter (1 through 6) steps ahead. This method acquires a lower aggregate RMSE than SSA-STS for $\alpha = 0.8$ and 1 . Although, it fails to outperform SSA-STS consistently, as it can be seen that the new SSA-STS algorithm dominates ETS for other α values, as well as for all longer horizon forecasts

Table 3. Aggregated RMSE for different step ahead forecasts for the artificial time series $I\alpha$ with $\alpha = 0.1, 0.2, 0.4, 0.6, 0.8, 1$ by 5 forecasting algorithms

	SSA-STS	Basic SSA	ARIMA	ETS	NN
Aggregated RMSE for 1, 2, ..., 6-step ahead forecasts					
I0.1	1.10	1.09	1.28	1.16	1.20
I0.2	1.23	1.34	1.36	1.26	1.57
I0.4	1.34	1.62	1.55	1.48	1.78
I0.6	1.51	1.94	1.91	1.57	2.57
I0.8	1.75	2.29	2.10	1.69	3.71
I1.0	1.93	2.67	2.22	1.84	4.67
Aggregated RMSE for 7, 8, ..., 12-step ahead forecasts					
I0.1	1.19	1.11	1.40	1.23	1.24
I0.2	1.34	1.26	1.56	1.41	1.54
I0.4	1.55	2.17	2.29	1.84	2.33
I0.6	1.78*	2.79	3.50	2.18	3.37
I0.8	2.02*	3.47	4.34	2.50	5.57
I1.0	2.31*	4.20	4.42	2.91	6.89

Note:* indicates that forecasting errors for a method is statistically significantly smaller than forecasting errors for all other methods by the one-sided KSPA test based on a p -value of 0.1.

(7 through 12 steps ahead). On the other hand, forecasts produced using ARIMA are considerably worse than the other methods, as ARIMA does not deal effectively with periodic patterns and nonlinear trends.

The KSPA test is carried out to verify whether the SSA-STS out-performance in terms of accuracy is statistically significant. We use this test instead of the traditional Diebold-Mariano (DM) test, due to the lack of necessary underlying assumptions. The KSPA test has been developed and applied in [18] to successfully test the predictive accuracy of SSA and ARIMA.

Indeed, as expected the KSPA one-sided test showed that for series with complex trends ($\alpha = 0.6, 0.8, 1.0$) in long horizons, the SSA-STS algorithm produces significantly smaller errors than the rest of the benchmark methods applied. These results from the statistical test corroborate the claim that SSA-STS dominates forecasting in such cases.

5.2 Forecasting of IIPs

We now consider eight time series of real industrial production for Germany, France and the UK. All series end in July 2019. However, the time series data for the UK starts in January 1998, January 1991 for Germany and in January 1990 for France. Data up to July 2014 is used in the estimation, and the post-sample forecast RMSE is used to measure performance, which is then tested by the one-sided KSPA test. We estimate parameters of ARIMA, ETS and NN forecasts at each cutting point, as mentioned in the previous section. However, only one choice of parameters for the SSA-STS and Basic SSA forecasting algorithms is maintained for all cutting points. These parameters are shown in Table 4, additionally parameters $L_t = 12$ and $r_t = 3$ were

chosen for the analysis of the trend in the SSA-STS forecasting algorithm.

Table 4. Parameters of the SSA-STS and Basic SSA forecasting algorithms for 8 IPIs for 3 countries

	SSA-STS				Basic SSA	
	L	r	L_s	r_s	L	r
UK						
Food	144	20	84	18	84	20
Basic						
Metals	108	15	60	13	60	12
Chemicals	60	22	60	20	72	23
Fabricated						
Metals	48	22	48	20	48	19
Utilities	48	6	36	5	36	6
Machinery	120	20	48	16	36	13
Electrical						
Equipment	120	26	24	17	60	18
Vehicles	144	20	36	16	60	17
Germany						
Food	84	28	84	26	84	20
Basic						
Metals	48	18	48	16	60	17
Chemicals	48	21	60	19	48	21
Fabricated						
Metals	36	27	72	22	48	16
Utilities	60	24	60	23	36	16
Machinery	36	16	84	14	36	16
Electrical						
Equipment	48	18	36	16	36	13
Vehicles	96	24	72	17	108	8
France						
Food	36	21	72	18	84	24
Basic						
Metals	48	19	72	16	72	22
Chemicals	60	20	72	18	72	20
Fabricated						
Metals	72	20	24	16	72	15
Utilities	36	9	36	8	36	9
Machinery	60	21	48	19	60	16
Electrical						
Equipment	60	20	36	17	72	14
Vehicles	48	17	48	14	120	14

We compute the aggregated RMSE of retrospective 12-month ahead forecasts with cutting points at each month between July 2014 and July 2018 to evaluate the forecasting performance of the five methods at different horizons. Table 5 and Table 6 present the RMSE results of the five methods for all the three countries and eight sectors, for short horizon (1 to 6 months ahead) and long horizon (7 to 12 months ahead) forecasts, respectively.

A summary RMSE Ratio (RRMSE) is also computed as the average ratio of the RMSE of the SSA-STS to that of other models for each country, and overall. Thus, a ratio of less than one is an indication that the SSA-STS model produces less errors on average for each country. RRMSE

Table 5. Aggregated RMSE for 1, 2, ..., 6-step ahead forecasts for 8 IPIs for 3 countries

	SSA-STS	Basic SSA	ARIMA	ETS	NN
UK					
Food	2.46	2.45	3.31	3.44	3.34
Basic	5.36*	6.41	6.59	6.42	9.37
Metals					
Chemicals	3.22	3.67	3.95	3.74	4.30
Fabricated					
Metal	4.24*	4.76	5.03	5.30	4.82
Utilities	4.44	5.04	4.86	4.76	5.24
Machinery	5.85	9.03	6.81	6.75	7.38
Electrical					
Equipment	4.67*	5.22	5.29	4.94	6.06
Vehicles	5.41	6.16	6.77	5.56	7.51
RRMSE		0.86	0.83	0.87	0.76
Germany					
Food	2.50	2.54	2.92	3.09	2.73
Basic	3.16	3.66	3.18	3.25	3.57
Metals					
Chemicals	3.01*	3.32	3.95	3.63	3.63
Fabricated					
Metal	3.48*	4.68	4.46	4.90	7.05
Utilities	3.15	3.10	2.99	4.47	3.88
Machinery	3.85*	4.49	4.70	5.47	6.54
Electrical					
Equipment	2.11	2.46	2.03	2.29	3.56
Vehicles	7.56	8.36	7.20	7.33	11.26
RRMSE		0.89	0.92	0.84	0.72
France					
Food	1.98	2.14	2.74	3.17	2.33
Basic	3.46	3.52	3.65	3.55	3.66
Metals					
Chemicals	4.44	4.60	4.45	4.65	5.53
Fabricated					
Metal	4.34	4.71	5.36	5.18	4.96
Utilities	5.22	5.38	4.94	6.14	5.76
Machinery	5.44*	5.83	6.13	6.18	7.06
Electrical					
Equipment	3.47*	3.99	4.27	4.33	4.56
Vehicles	6.75*	7.84	7.73	7.83	8.85
RRMSE		0.93	0.89	0.85	0.83
RRMSE (overall)		0.89	0.88	0.85	0.77
Score (overall)	19	1	4	0	0

Note: * indicates that forecasting errors for a method is significantly smaller than errors for all other methods by the one-sided KSPA test based on a p -value of 0.1.

(overall) is also computed for all 24 series. The Score (overall) is a score of the number of times out of 24 that each model yields the lower RRMSE.

In terms of producing smaller RMSE, the results in Tables 5 and 6 provide strong evidence that the SSA-STS model is generally superior to Basic SSA, ARIMA, ETS and NN models for all the three countries. Overall, the results

Table 6. Aggregated RMSE for 7, 8, . . . , 12-step ahead forecasts for 8 IIPs for 3 countries

	SSA-STS	Basic SSA	ARIMA	ETS	NN
UK					
Food	2.67	2.73	4.24	3.82	4.72
Basic	6.97*	7.78	10.10	9.69	13.03
Metals					
Chemicals	3.23*	3.88	5.21	4.41	5.27
Fabricated					
Metal	4.52	6.14	5.39	5.23	5.37
Utilities	5.13	5.11	5.16	5.47	6.11
Machinery	6.30	11.53	8.26	8.97	11.53
Electrical					
Equipment	5.03*	6.66	7.36	5.89	7.56
Vehicles	6.84*	7.88	9.44	7.22	9.64
RRMSE		0.83	0.74	0.81	0.66
Germany					
Food	2.53	2.60	3.32	3.34	3.02
Basic	3.32	3.84	3.13	3.39	4.43
Metals					
Chemicals	3.60	3.67	5.19	3.99	4.58
Fabricated					
Metal	3.98*	4.59	4.99	5.49	11.10
Utilities	3.353.57	3.27	4.47	4.37	
Machinery	4.12*	4.71	4.62	5.76	8.57
Electrical					
Equipment	3.38*	3.73	3.46	4.08	4.92
Vehicles	8.28	9.18	8.59	8.26	16.36
RRMSE		0.91	0.90	0.83	0.65
France					
Food	2.18*	2.14	3.00	3.28	2.40
Basic	3.53*	3.96	3.99	3.92	4.49
Metals					
Chemicals	4.27*	4.47	5.10	5.22	6.45
Fabricated					
Metal	4.21*	4.78	5.83	5.24	5.91
Utilities	5.20	5.56	4.92	5.28	5.04
Machinery	5.57	5.94	5.76	6.09	7.83
Electrical					
Equipment	3.75*	5.01	4.86	4.84	5.08
Vehicles	7.13*	8.00	8.76	8.60	11.90
RRMSE		0.91	0.85	0.84	0.77
RRMSE (overall)		0.88	0.83	0.83	0.69
Score (overall)	18	2	3	1	0

Note:* indicates that forecasting errors for a method is significantly smaller than errors for all other methods by the one-sided KSPA test based on a p -value of 0.1.

show that for 1 to 6 step ahead forecasts, SSA-STS outperformed Basic SSA, ARIMA, ETS and NN by 11%, 12%, 15% and 23%, whereas for 7 to 12 step ahead the difference was 12%, 17%, 17% and 31%, respectively. Hence, for longer horizons, SSA-STS performance exceeds all other methods by a larger margin. Although the Score (overall) statistics corroborates these results, by indicating that the SSA-STS

model produces lower RMSE for 19 and 18 cases out of the 24 (in aggregate for 1–6 and 7–12 step ahead forecasts respectively), one would expect the statistics to be switched if they were to support the findings that SSA-STS performs better for longer horizon forecasts.

Although, at face value the predictive accuracy advantage of SSA-STS is clear, we perform the one-sided KSPA test to ensure the results are statistically significant. Based on the 0.1 p -value, we see in Tables 5 and 6 that the SSA-STS method consistently had statistically significant lower errors for many IIPs over both or some horizons. As in the simulation study, for the real time series SSA-STS proved to be particularly superior when forecasting over longer horizons for time series with complex trends, in particular, for many of the French IIPs.

5.2.1 Depictions for UK IIPs

We show graphically the UK IIPs 12-month forecasts created at each cutting point in different colours for each forecasting scheme in Figure 2. Hence, the amount of deviation, or spread of the colourful lines from the reconstructed series (in black) confirms the adequacy of the forecasting methods deduced from the pooled RMSE results.

For instance, the Utilities IIP shows most of the forecasts coinciding for all the forecasting methods, as expected due to its simple trend and consistent periodic movement.

Whereas, for more complex IIPs, such as Chemicals and Machinery and Equipment, we observe the advantage of SSA-STS quite clearly. Since its retrospective forecasts concur, indicating that this method would be a reliable way of accurately forecasting these series, when compared to the benchmark methods. Interestingly, for the series of Machinery and Equipment the trend is especially complex and there seems to be some type of structural break around 2014, hence the benchmark methods fail to account for this. The steep red lines which are increasing when the true reconstruction declines (in black) depicts this problem precisely.

6. CONCLUSIONS

Building on the work in [13, 14, 11], we develop a greatly flexible SSA-STS forecasting algorithm for the application to series with complex trends and the presence of seasonality, such as commonly seen in IIP data. Its performance is investigated against that of popular methods, including Basic SSA, ARIMA, ETS and NN, on real and simulated data. The real data used consists of IIPs of eight major sectors in Germany, France and the UK. In the simulation study a series of similar form is considered, but allowing for variation in trend complexity. The superiority of SSA-STS is evident on both sets of data, with some minor exceptions. Those being in cases where there is no trend complexity, which eradicates the SSA-STS algorithm’s advantage. A good example of such a series is the Utility IIPs, that have more systematic behavior than the majority of the other series, due to very

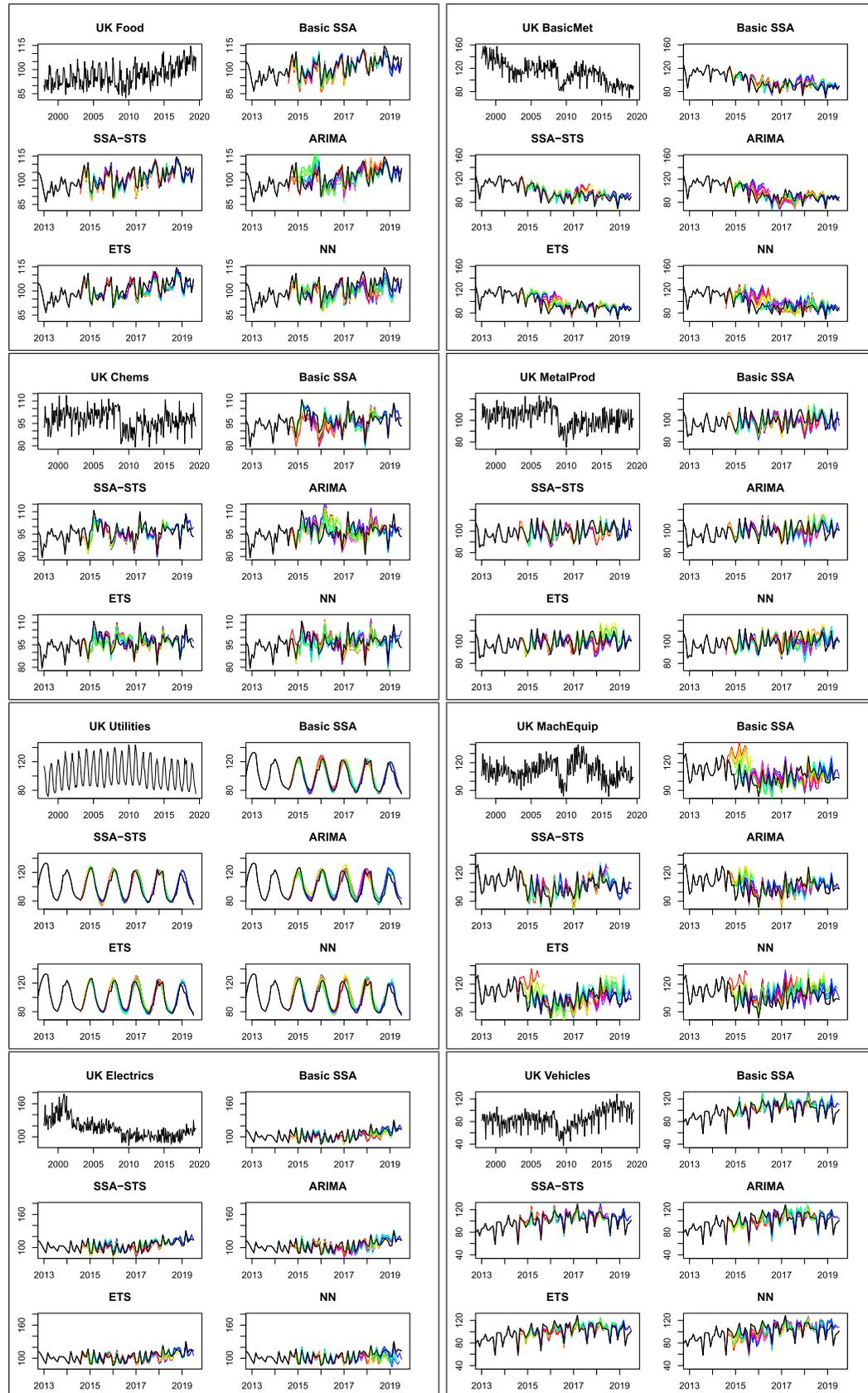


Figure 2. 8 UK IIPs and retrospective 12-month ahead forecasts by 5 methods.

1 slight variations in the periodic annually and the absence
2 of a trend and structural breaks. Hence, for such series, the
3 other methods may create similar or preferable results than
4 the SSA-STS algorithm.

5 Since the dominance of SSA-STS stems from the sepa-
6 rate evaluation of the trend and seasonal sub-cycles, this
7 entails a larger set of parameter options. Hence, an impor-
8 tant constituent for the success of the SSA-STS algorithm
9 is the correct selection of parameters L , r , L_s , r_s , L_t , r_t , as
10 for many SSA-type methods this choice can determine the
11 performance.

12 At a more general level, we can conclude that the SSA-
13 STS forecasting algorithm outperforms other methods in in-
14 stances where a trend of complex shape is present. This is
15 indicated during the development stages of the algorithm,
16 as forecasting the trend and periodic separately would sug-
17 gest that forecasting each in isolation enhances the final
18 result, especially in the case of a complex trend. This is
19 substantiated by the results from the retrospective forecasts
20 discussed. Since the majority of industrial production in-
21 dicators have complex trend, the SSA-STS algorithm is a
22 favorable forecasting method for such data. However, it can
23 also be successfully applied for forecasting many other types
24 of real data, and in particular recommended for forecasting
25 series with trend complexities and structural breaks.

26 ACKNOWLEDGEMENT

27 The work of A. Pepelyshev was partially supported by
28 the Russian Foundation for Basic Research (project no. 20-
29 01-00096).

30 *Received 6 May 2021*

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