Charitable giving, social capital, and positional concerns

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\begin{abstract}
Research on the effects of positional concerns on individuals’ attitudes and behavior in certain policy-relevant areas is lacking. In this paper, we investigate the relationship between positional concerns, charitable giving, and social capital. We use data from the “Caucasus Barometer” survey administered in three post-Soviet transition economies: Armenia, Azerbaijan, and Georgia. Our analysis proceeds in two phases. First, controlling for absolute income and other individual and household characteristics, we show an association between positional concerns and charitable giving as well as between positional concerns and social capital. Second, we use an instrumental variable model that uses heteroskedasticity-based instruments generated through Lewbel’s method to provide supporting evidence of the causal impact of positional concerns on the outcome variables of interest.

We find that the relative deprivation of a household can have negative impacts on its members’ charitable giving and social capital.
\end{abstract}

1. Introduction

Individuals are concerned with their relative position in society and tend to compare themselves with relevant others (Luttmer, 2005). In the literature, such positional concerns are dubbed the ‘comparison effect’ (e.g., (Senik, 2004)) or ‘keeping up with the Joneses effect’ (e.g., (Aronsson & Johansson-Stenman, 2008)) and can have a detrimental impact on human behavior and feelings. In particular, the ‘comparison effect’ can increase consumption of redundant goods and services (e.g., (Corneo & Jeanne, 1997; Duesenberry, 1949; Frank, 1985; Alpizar, Carlson, & Johansson-Stenman, 2005; Solnick & Hemenway, 2005)), intensify temporary emigration (Antinyan & Corazzini, 2018; Stark & Yitzhaki, 1988; Stark & Taylor, 1991; Bhandari, 2004, Quinn, 2006), shrink life satisfaction (Luttmer, 2005; Clark & Oswald, 1996; Ferrer-i-Carbonell, 2005; Antinyan, 2016; Cojocaru, 2016), promote risk-taking (Hill & Buss, 2010; Müller & Rau, 2019) and intensify social protests (Van Stekelenburg & Klandermans, 2013).

In this paper, we study how one’s perceived positional advantage and deprivation relate to one’s recalled and self-reported charity donation decisions (i.e., whether an individual donated to charity or gave money to beggars or not in the last 6 months) in the three republics of the South Caucasus: Armenia, Azerbaijan, and Georgia.\textsuperscript{1} In addition, we investigate the link between one’s perceived relative position and social capital in the South Caucasus, measured by trust toward others and trust toward secular institutions.\textsuperscript{2} For our purposes, we use a nationwide (cross-sectional) survey administered in Armenia, Azerbaijan, and Georgia.

The contribution of our work is multifold. First, we contribute to the multidisciplinary literature on charitable giving, which strives to understand (i) why people donate (e.g., Andreoni, 2006; Bekkers & Wiepking, 2012); (ii) who donates (Wiepking & Bekkers, 2012); (iii) how to increase the donations (e.g., Meier, 2007; Karlan & List, 2007; Huck & Rasul, 2011). Regarding the characteristics of donors, the literature puts forth individuals’ gender, family composition, and absolute income as important predictors of charitable giving (Wiepking & Bekkers, 2012). Nonetheless, our reading of the literature suggests that little is known about whether one’s perceived relative income compared to her peers can serve as a predictor of

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\textsuperscript{2} From here onward, whenever speaking about the relative standing of the respondents in the study we will refer to “one’s perceived relative standing”. For the brevity and clarity of the paper we will not mention the entire phrase repeatedly.

https://doi.org/10.1016/j.socec.2022.101929
Received 15 February 2021; Received in revised form 4 August 2022; Accepted 16 August 2022
Available online 17 August 2022
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charitable giving. This is somewhat important, as relative income and absolute income may not necessarily coincide: individuals may perceive themselves to be poorer relative to others than they actually are (e.g., Karadja, Mollerstrom, & Seim, 2017). In other words, the charitable donations of individuals \( i \) and \( j \) with identical demographic characteristics and absolute wealth can vary, simply because individual \( i \) perceives herself to be poorer than her peers, while individual \( j \) does not do so.

Second, given that charity donations can be motivated by prosocial motives such as altruism (Andreoni, 2006) we add to the inconstant literature on positional concerns and social preferences. While some experiments illustrate a positive (negative) impact of relative deprivation (advantage) on prosocial preferences (e.g., Guinote, Cotzia, Sandhu, & Siwa, 2015), others find no relationship (e.g., Greitemeyer & Saggion, 2019). In addition, the existing research suffers from a number of methodological weaknesses such as external validity, small sample sizes and sampling, publication or design biases (e.g., Kordüfer, Egloff, & Schmukle, 2015 and the references therein). We depart from this literature in two fundamental ways. While the mentioned papers proxy prosocial preferences through various acts such as allocation of needles on voodoo dolls, number of pens an individual helps to pick up from the floor, choice of prosocial professions and the like, we focus on one’s recalled and self-reported charity donation decisions, which we believe is a more relevant variable for the economy (e.g., Andreoni, 2006; List, 2011). Furthermore, we use multiyear nationwide surveys from three countries which allows us to circumvent the problems related to external validity, small sample sizes, and sampling. In other words, we study a prevalent real world (self-reported) behavior and we can generalize our findings beyond the sample under investigation. However, our approach is not without problems either, which we will discuss later in the text.

Third, we extend the paper of Fischer & Torgler (2013) to lower middle-income and upper middle-income post-Soviet countries in transition. The authors illustrate the social capital-lowering effect of positional concerns in 26 (mostly) institutionally and economically well-developed countries and justifiably posit that their findings can have important implications for developing economies, though a direct test of this claim seems missing in the literature.

Why do we focus on charitable donations and social capital? Similar to many lower- and middle-income countries (LMIC), the three republics of the South Caucasus face diverse socio-economic challenges that cannot be easily resolved since the governments do not have enough capacity and financial resources to provide essential public goods and services. Thus, charitable donations (both by individuals and organizations) become an important instrument through which various societal problems are settled. For example, 28% of deaths from non-communicable diseases (NCDs) is attributed to cancer in Armenia (World Health Organization, 2018). Not all families are able to keep up with the high costs of cancer treatment, since out-of-pocket payments (World Health Organization, 2018). Not all families are able to keep up with the high costs of cancer treatment, since out-of-pocket payments (World Health Organization, 2018). Not all families are able to keep up

Table 1

<table>
<thead>
<tr>
<th>Note</th>
<th>Armenia</th>
<th>Azerbaijan</th>
<th>Georgia</th>
</tr>
</thead>
<tbody>
<tr>
<td>GDP growth, %</td>
<td>2010-2015 average</td>
<td>4.4</td>
<td>2.8</td>
</tr>
<tr>
<td>GDP per capita</td>
<td>PPP, current USD, 2015</td>
<td>8,419</td>
<td>17,776</td>
</tr>
<tr>
<td>Gini index</td>
<td>Latest available year, 2015*</td>
<td>32.4</td>
<td>31.8</td>
</tr>
<tr>
<td>Unemployment rate, %</td>
<td>ILO estimates, 2014</td>
<td>17.1</td>
<td>5.2</td>
</tr>
<tr>
<td>Inflation, %</td>
<td>2010-2015 average</td>
<td>5.8</td>
<td>4.1</td>
</tr>
<tr>
<td>Poverty headcount ratio</td>
<td>At national line, 2012</td>
<td>32.4</td>
<td>6.0</td>
</tr>
<tr>
<td>Export (% of GDP)</td>
<td>Goods and services, 2015</td>
<td>29.8</td>
<td>37.8</td>
</tr>
<tr>
<td>Corruption perception index</td>
<td>2016 Rank (Score in parenthesis)</td>
<td>113 (33)</td>
<td>123 (30)</td>
</tr>
</tbody>
</table>

Note. Source for all data except Corruption perception index (CPI), - World Development Indicators database. Source for CPI - Transparency International. (*) The most recent available data for the Gini index for Azerbaijan dates back to 2008.
3. Behavioral predictions

The mainstream economic theory neglects the significance of social influences on individual decisions and postulates that absolute consumption or income is what matters for individuals: the higher the consumption or income, the higher the utility of an individual. Nonetheless, such prominent social scientists as Adam Smith, Albert Hirschman, Karl Marx, Thorstein Veblen, and James Duesenberry, among many others, acknowledge the importance of social influences on human behavior (e.g., McCormick, 1983; Hirschman & Rothschild, 1973). Following this line of reasoning, individuals are concerned with their relative position in the society and tend to compare themselves with relevant others (also called a reference group). If one is in a disadvantaged (or advantageous) situation vis-à-vis the reference group, then she considers herself relatively deprived (or advantaged). These comparisons can result in two diametrically opposite effects – a “keeping up with the Joneses” effect or a “tunnel” effect.

According to the “keeping up with the Joneses” effect, individuals who stand high in the social hierarchy tend to mimic the behavior of those who stand high in the social hierarchy to maintain a high social status relative to others in the society (Veblen, 1899; Trigg, 2001). Similarly, individuals make upward comparisons and change their consumption patterns because of the need to eliminate the feelings of inferiority created by other people consuming superior goods (e.g., McCormick, 1983; Mason, 2000). Relative disadvantage may trigger negative feelings such as dissatisfaction with life (Lutten, 2005; Clark & Oswald, 1996; Ferrer-i-Carbonell, 2005; Antinyan, 2016; Cojocaru, 2016) and intensify deleterious behavior, such as the consumption of redundant goods and services (e.g., Corneo & Jeanne, 1997; Duesenberry, 1949; Frank, 1985; Alpizar et al., 2005; Solnick & Hemenway, 2005), temporary emigration (Antinyan & Corazzini, 2018; Stark & Yitzhaki, 1988; Stark & Taylor, 1991; Bhandari, 2004; Quinn, 2006), and risk-taking (Hill & Buss, 2010; Müller & Rau, 2019).

According to the “tunnel” effect, the situation of an individual’s reference group serves as an informational device to form expectations regarding prospective opportunities (Hirschman & Rothschild, 1973). For instance, if an individual’s neighbor or acquaintance improves her material wellbeing, this can serve as a signal that this individual will also improve her material wellbeing in the future. Thus, according to the “tunnel” effect, relative disadvantage should generate positive feelings since it can serve as a signal of upward mobility. For instance, in the context of post-Soviet economies, (Senik, 2004) documents that an individual’s reference group income exerts a positive influence on her satisfaction with life. In other words, relatively disadvantaged individuals (i.e., those who compare themselves with a reference group richer than they are) increase their life satisfaction. Thus, the income of others is used as an informational device à la Hirschman and Rothschild. Despite the “tunnel” effect documented in Russia, (Antinyan, 2016) and (Antinyan & Corazzini, 2018) illustrate that the “keeping up with the Joneses” effect prevails in the LMICs of the South Caucasus. More specifically, relatively disadvantaged individuals decrease (i.e., “keeping up with the Joneses” effect) rather than increase (i.e., “tunnel” effect) their life satisfaction (Antinyan, 2016) and are more likely to express intentions of temporary migration (Antinyan & Corazzini, 2018). Stemming from these findings, the remaining discussion in this section is developed from the perspective of the “keeping up with the Joneses” effect.

What is the impact of relative disadvantage on charitable giving and trust toward others? Relative disadvantage may generate feelings of frustration and social injustice (Smith & Pettigrew, 2015; Hill & Buss, 2006; Klandermans, 2015; Moscatelli, Albarello, Prati, & Rubini, 2014, Kawachi, Kennedy, & Wilkinson, 1999). For instance, relatively disadvantaged individuals may be rather frustrated since they may perceive that they are unjustifiably exploited by those who are economically better-off and that such unequal distribution of the resources in the society is the outcome of unequal distribution of power between economic agents rather than the result of market forces (Fischer & Torgler, 2013). Indeed, according to nationally representative surveys administered in the USA, relative deprivation triggers strong feelings of an unjust society (Newman, Johnston, & Lown, 2015). The perception of undeserved disadvantage may also trigger anger, evoking hostile emotional reactions (Greitemeyer & Sagioglou, 2017). For instance, violent crimes may often be explained by relative deprivation that leads to an individual’s personal frustration and hostility (Greitemeyer & Sagioglou, 2017) and the references therein. In sum, the negative emotional reactions (e.g., envy, frustration, anger, hostility, perception of social injustice) activated by relative deprivation can induce a disadvantaged individual to develop rather negative feelings and behavior toward other members of the society (including individuals in her reference group) (Fischer & Torgler, 2013). Under these circumstances, disadvantaged individuals can behave less altruistically toward others and exhibit lower levels of trust in others. Nevertheless, it’s worth noting that negative emotional reactions triggered by relative disadvantage may not be the only mechanism explaining the relationship between relative disadvantage and charitable giving. More specifically, individuals may be guided by egalitarian concerns (Dawes, Fowler, Johnson, McElreath, & Smirnov, 2007; Epper, Fehr, & Senn, 2020; Müller & Renes, 2021). For instance, an individual who perceives herself as poorer than her reference group may ideally want to be the target for charity rather than give money to charity. Based on this discussion we can formulate the following hypotheses:

Hypothesis 1. Relative deprivation makes individuals less altruistic toward other members of the society needing help.

Hypothesis 2. Relative deprivation decreases individuals’ trust in other members of the society.

What is the impact of relative disadvantage on trust toward secular institutions? Relatively disadvantaged individuals may hold the state institutions responsible for unfair distribution of societal wealth and therefore may blame these institutions for the relative income disadvantage they suffer compared to their reference group (e.g., Fischer & Torgler, 2013).

This problem can be especially stark in post-Soviet countries, because since the fall of the Soviet Union the institutions in these countries have generally been in the hands of a wealthy minority that does not usually act in the best interest of the society. In other words, the sense of exploitation because of unfair distribution of the societal wealth may be higher in post-Soviet republics compared to Western democracies. To conclude, the frustration and the feeling of exploitation that individuals develop because of relative deprivation can decrease the trust not only on the horizontal but also on the vertical level (Fischer & Torgler, 2013). Stemming from this discussion, we develop the following hypothesis:

Hypothesis 3. Relative deprivation decreases individuals’ trust in secular institutions.

The existing literature mainly discusses the impact of relative deprivation on individuals’ attitudes and feelings, which is in line with Duesenberry’s original claim that comparisons are mainly asymmetric and upwards (i.e., poorer individuals can be negatively influenced by the income of richer reference group members, while the opposite may not necessarily be true). Nevertheless, the empirical evidence obtained from the South Caucasus suggests that the comparisons in this region can be asymmetric, as relatively disadvantaged individuals are positively influenced by the presence of poor reference group members (e.g., Antinyan & Corazzini, 2018; Antinyan, 2016). Thus, one may be tempted to conjecture that relative advantage may be positively related to social capital and social preferences. Given the lack of empirical and theoretical literature about the impact of relative advantage, we do not

3 We thank an anonymous referee for this important comment.
form testable hypotheses. Instead, we will try to rationalize the findings ex post facto.

4. Empirical strategy

4.1. The regression equation and the estimation approach

The main purpose of the current study is to empirically test the effect of positional concerns on charitable donations and social capital. First, we estimate a regression equation of the following form:

\[
Y_{ijt} = \beta_0 + \beta_1 \times \text{BelowAverage}_{ijt} + \beta_2 \times \text{AboveAverage}_{ijt} + \beta_3 \times X_{ijt} + \beta_4 \times T + \beta_5 \times C_i + \beta_6 \times \text{REG} + \epsilon_{ijt}
\]

(1)

\(\text{BelowAverage}_{ijt}\) and \(\text{AboveAverage}_{ijt}\) indicate the relative position of individual \(i\) in country \(j\) at time \(t\); \(X_{ijt}\) includes the absolute income of individual \(i\) in country \(j\) at time \(t\) as well as additional individual and household socio-demographic control variables, such as age, gender, number of household members, education, marital status and employment status. The descriptive statistics for these variables are available in Table A1 of the SI Appendix. These independent variables are specified based on data availability and are in line with previous studies on happiness, migration, and social capital (e.g., Antinyan & Corazzini, 2018; Ferrer-i-Carbonell, 2005; Antinyan, 2016; Fischer & Torgler, 2013). Furthermore, most of the specified variables are important drivers of social preferences and trust. For example, social preferences and trust can vary by gender (e.g., Croson & Gneezy, 2009; Buchan, Croson, & Sollnick, 2008), absolute income (Smeets, Bauer, & Gneezy, 2015; Gangadharan, & Nikiforakis, 2011; Johansson-Stenman, Mahmud, & Martinsson, 2005), regional differences within a country (e.g., Bigoni, Bortolotti, Casari, Gambetta, & Pancotto, 2016, 2019) and education (Huang, Van den Brink, & Groot, 2009). \(T\), \(C_i\), and \(\text{REG}\) dummy variables control for (other unobserved) temporal, national, and regional differences. The dependent variable \(Y_{ijt}\) is either the charity donation decision of individual \(i\) in country \(j\) at time \(t\), or the self-reported social capital. We still cannot account for individual personal traits that may drive respondents’ perceptions, feelings and behavior. In this regard, the literature on life satisfaction confirms the impact of positional concerns on life satisfaction both with panel data estimators that allow one to control for individual personal traits, as well as with cross-sectional data estimators that omit individual personal traits (e.g., Senik, 2004; Clark & Oswald, 1996; Ferrer-i-Carbonell, 2005; McBride, 2001). These results provide us with reasons to think that utilizing cross-sectional data and omitting individual personal traits should not substantially bias our results.

Our dependent variables are either binary or measured on a scale of ‘1’ to ‘10’ or of ‘1’ to ‘5.’ In case of binary variables, we estimate linear probability (LPM) and probit models. If variables are measured on a scale of ‘1’ to ‘10’ (or of ‘1’ to ‘5’), we treat them as interval variables and estimate OLS models. Furthermore, to test the robustness of the results, we preserve the ordinal nature of the scale and estimate ordered probit models. Stemming from research on life satisfaction, there should be (qualitatively) very little difference between the results of OLS and ordered probit (or logit) models (Ferrer-i-Carbonell & Frijters, 2004). Throughout the text, we only interpret those coefficients that are robust across OLS and ordered probit models as well as across probit and linear probability models. Further, we account for the potential heteroskedasticity of residuals by introducing robust standard errors. For each outcome variable we run pooled sample regressions and country-specific regressions. In the tables, we report only the coefficients of relative and absolute income variables for the sake of brevity and clarity of the text. The complete tables with all the variables are available in the S3 Appendix.

4.2. Data

We use cross-sectional data from the nationwide “Caucasus Barometer” survey, developed by the Caucasus Research Resource Centers (CRRC from here onward). The “Caucasus Barometer” has been administered in the South Caucasus since 2004. Starting from 2013 the survey has been administered once every two years and omits Azerbaijan because of bureaucratic complexities. To the best of our knowledge, the “Caucasus Barometer” is the only survey that collects comparable data from Armenia, Azerbaijan, and Georgia, while containing a rich set of variables assessing the socio-demographic and economic conditions of the respondents at both individual and household levels. The data are collected through face-to-face interviews that take place in the households of the respondents. The sampling frame of the survey is the list of registered voters of voting precincts, while the respondents’ households are selected using the “random walk” procedure. The respondent in a household is selected with a Kish grid. In the following analysis, we use information from the 2010–2013 waves. An interested reader can refer to the CRRC webpage for more details (http://www.crcccenters.org/).

We would also like to note that the survey questionnaire is rather long and the questions that measure the variables of interest are quite far from each other. Thus, it is very unlikely that the responses are biased because of potential “priming” effects. Such priming bias could occur if, for instance, the participants had to answer the questions about their households’ relative and absolute income just before providing an answer to the charity donation question.

4.3. Dependent variables

To proxy the charitable behavior of the individuals, we use a survey question that directly elicits respondents’ recalled charitable giving: “Could you please tell me which of these activities you were involved in during the past 6 months?” “Contributed to a charity, including donations by SMS and giving money to beggars” is one of the activities mentioned in the questionnaire, with a binary “Yes” or “No” answer. Similar questions are asked in other surveys, such as the German Socio-Economic Panel (SOEP), the American Consumer Expenditure Survey (CEX), and the General Social Survey (GSS). Please note that large private donations to charities by ordinary individuals living in Armenia, Azerbaijan, and Georgia are not widespread phenomena; meanwhile donating small amounts of money through SMS or giving money to beggars on the streets is rather common. Thus, we most likely deal with small charity donations on average, though we cannot exclude that a few individuals in the sample made large contributions.

We would like to warn the readers that a glimpse of caution is required when working with such survey questions, since helping others can be perceived as a sensitive issue. A respondent may not report her true behavior so as not to harm her identity because of concerns about social approval. The magnitude of this concern can be rather heterogeneous among the respondents. Despite this problem, such questions are considered as important sources of information and are actively used to uncover the determinants of individuals’ social preferences (e.g., Korn dorfer et al., 2015). Furthermore, recent empirical studies demonstrate that survey questions can be as good predictors of individual preferences (including social preferences, such as altruism) as incentivized experiments (e.g., Falk, Becker, Dohmen, Huffman, & Sunde, 2016).

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4 To validate the use of a linear probability model, we compare the deviation between fitted values obtained from the LPM and predicted probabilities obtained from the probit model. Our results indicate very little deviation between the LPM fitted values and the predicted probabilities obtained from the probit model. Lastly, only a small fraction of fitted values obtained from LPM models is negative. Both observations confirm the quality of LPM coefficients. Please refer to Figure E1 in SS Appendix for the complete results.
In line with (Fischer & Torgler, 2013), we acknowledge that social capital has multiple facets: horizontal trust, vertical trust, and norm compliance. However, the survey does not contain questions to measure the norm compliance of respondents, so we do not discuss this facet in the paper. We proxy horizontal (generalized) trust with the following question measured on a scale of “1” (you can’t be too careful) to “10” (most people can be trusted): “Generally speaking, would you say that most people in / country can be trusted?” The second facet of social capital is captured by questions that assess respondents’ trust in state institutions: “I will read out a list of social institutions and political unions. Please indicate your level of trust toward each of them on a 5-point scale, where ‘1’ means ‘Fully distrust,’ and ‘5’ means ‘Fully trust’.” The institutions were shuffled for each respondent to exclude potential order effects. We concentrated our attention on the three principal decision-making institutions in the region: the parliament, the president, and the executive government (prime minister and the ministers). We constructed a combined index by averaging the trust toward the seminal institutions and treating it as our variable of interest.

Figs. 1–3 illustrate the average values of the variables of interest broken down by countries and years. Interestingly, the outcome variables are different in levels and exhibit different temporal dynamics across countries.

### 4.4. Relative and absolute income

To identify the respondents’ relative position vis-à-vis the reference group, the literature suggests two approaches.

First, one can use the objective measure of welfare, which compares an individual’s absolute income (or consumption) with the average income (or consumption) of her reference group (e.g., Ferrer-i-Carbonell, 2005). If a researcher follows this approach, an inherent difficulty is to define the reference group of individuals (e.g., Clark & Senik, 2010). To this date, no unequivocal definition of a reference group does not exist. The reference group of an individual can include the entire population of the country (e.g., Easterlin, 1995), professional peers (e.g., Senik, 2004) or neighbors living in the same community (e.g., Luttmer, 2005; Bhandari, 2004; Quinn, 2006).

Second, one can opt for the subjective measure of welfare, which serves as a viable source of additional information, however, it is largely neglected by scholars (see the excellent discussion in (Ravallion & Lokshin, 2010)). In a subjective measure, survey respondents are asked to compare their well-being with that of their comparison group. Thus, it is left to the participant to define who she compares herself with and by what standard their economic status is to be judged (e.g., Ravallion & Lokshin, 2010). In other words, the subjective measure of relative wealth evokes individuals’ first-person view of the social world and themselves within it (Schneider, 2019). Despite the obvious methodological differences, both approaches yield qualitatively similar results when assessing the impact of relative concerns on subjective well-being (e.g., Ravallion & Lokshin, 2010).

In this paper, we adopt the second approach and use self-reported information on a respondent’s perception of the relative standing of her household in comparison to that of the households around her. The survey participants answered the following question on a scale of “1” (Very Poor) to “5” (Very Good) “Relative to most of the households around you, would you describe the current economic condition of your household as ….?” Hence, the respondent must judge which households around her constitute the reference group, and how the economic status of her household compares to that of her reference group. Based on the responses, we construct the dummy variable Above Average if the individual perceives the conditions of her household as either ‘Very Good’ or ‘Good’ vis-à-vis the reference group. Similarly, we create the dummy variable Below Average if the respondent’s answer is either ‘Poor’ or ‘Very Poor.’ Individuals who answer ‘Fair’ to the abovementioned question, i.e., perceive their households to be neither disadvantaged nor advantaged compared with surrounding households, constitute the omitted category.

In addition to the relative income measures, we need to evaluate the absolute income of the respondent’s household. The inclusion of absolute income variables ensures separate identification of relative and absolute income effects. To categorize the households into different income groups, we use respondents’ answers to the following question: “Household income is the sum of the monetary income of all household members. Speaking about the monetary income of all your household members last month, to which of the following groups does your household belong?” The respondents do not report exact figures and are asked to indicate to which of the eight income groups their households belong. It can be of potential concern that the income of the last month may not capture the average income throughout the year. We check the income stability in Armenia and Georgia utilizing survey data collected by the national statistical services of these countries. The most important takeaway of this exercise is that the monthly income is generally stable during the year (at least for the time interval under consideration). For the sake of brevity, we relegate this discussion to S2 Appendix.

As for relative advantage, Fig. 5 suggests that relatively advantaged households are present in all income categories. One can also detect that the number of relatively advantaged households spikes in the two highest income groups. Nonetheless, relative advantage and absolute wealth never coincide. To state it differently, not all individuals who declare either zero or little household income consider their households’ economic condition worse than that of the reference group.

As for relative disadvantage, Fig. 4 suggests that relatively disadvantaged households are present in all income categories. One can also detect that the number of relatively deprived households gradually reduces in absolute income. Last, but not the least, relative and absolute poverty do not coincide. In other words, not all individuals who declare either zero or little household income consider their households’ economic condition worse than that of the reference group.

### 5. Results

#### 5.1. Charitable donations and positional concerns

Table 2 illustrates the relationship between the positional concerns and charitable donations. The highly negative and significant coefficient of Below Average Dummy in the pooled and country-specific regressions suggests a negative association between relative deprivation and donation decisions. This evidence is aligned with Hypothesis 1, which claims that relatively deprived individuals should be more selfish than non-deprived individuals. The most reasonable interpretation of the Below Average Dummy coefficient in the pooled regression is that the probability of charitable donations by relatively deprived individuals may be reduced by approximately 7 percentage points compared with individuals who feel neither deprived nor advantaged (i.e., the omitted category).

As indicated by the coefficient of the Above Average Dummy, we also document a positive relationship between relative advantage and donation decisions. The most reasonable interpretation of this coefficient, in the pooled regression, is that the probability of donations by relatively advantaged individuals may increase by approximately 7
percentage points compared with the omitted category. Nevertheless, the relationship between relative advantage and donations is heterogeneous across countries. While in Georgia and Azerbaijan, relative advantage is positively correlated with the social preferences of individuals compared to the reference category, in Armenia there is no effect. We do not have a meaningful explanation for the heterogeneous result across countries.

How can one rationalize the positive association between relative advantage and charitable donations? Presumably, living in a household that is better off than the reference group enhances the social status of the individuals belonging to the household. The experimental literature illustrates a positive link between social status and altruism: the higher the status of the decision-maker, the more she donates in a dictator game experiment (e.g., Liebe & Tutic, 2010).

Regarding absolute income, there is a positive relationship between income and charity donation decisions. The result that prosocial behavior may be correlated with absolute wealth is in line with the

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5 These effects are sizable considering that the average incidence of charitable donations in our data for the period considered is approximately 30 percent.
findings of (Korndörfer et al., 2015; Smeets et al., 2015; Trautman et al., 2013). Utilizing representative survey data that covered a broad range of countries, (Korndörfer et al., 2015) illustrate that higher-class individuals are more likely to make charitable donations and contribute a larger percentage of their family income to charity. (Trautman et al., 2013) use Dutch data and illustrate that individual wealth is positively correlated with prosocial behavior (e.g., volunteering). Smeets et al. (2015) play a dictator game with millionaires and conclude that the latter give away more than in any other example in the literature if matched with a low-income participant. Nonetheless, we would also like to note that both field and lab studies raise questions about the positive relationship between wealth and prosocial behavior, and the evidence is rather mixed (an interested reader can refer to Erkal et al., 2011; Andreoni, Nikiforakis, & Stoop, 2017 and the references therein).

5.2. Social capital and positional concerns

First, we fix our attention on the first facet of social capital and analyze the relation between positional concerns and horizontal (generalized) trust. Table 3 illustrates the estimation results.

The negative and significant coefficient of Below Average Dummy in pooled and country-specific regressions demonstrates a negative relationship between trust toward others and relative deprivation. This result backs Hypothesis 2, which claims that relatively deprived individuals should trust other members of the society less than others.

Regarding the relationship between positional advantage and...
generalized trust, unlike the existing literature (i.e., Fischer & Torgler, 2013) we do not demonstrate a trust-lowering effect of relative advantage in the countries under scrutiny. On the contrary, we even detect a positive link between positional advantage and trust in Azerbaijan. As discussed elsewhere in the text, positional advantage of individuals may enhance their social status in the reference group/neighborhood. Possessing high status may lead one to trust others more (e.g., Lount & Pettit, 2012), since a high-status individual may perceive others as having positive intentions toward her. Furthermore, in traditional societies, the expectations of a high-status individual may well be met, as status has more expositional features than in more contemporary settings, in the sense that a high-status individual is often treated better than the rest.

Turning to absolute income, we find a positive association between income and trust in Armenia (marginal) and Georgia and no relationship in Azerbaijan (though the coefficient is positive). The finding that trust can increase with absolute income is in line with the literature (e.g., Alesina & La Ferrara, 2002).

Next, we analyze the link between positional concerns and trust toward secular institutions (i.e., president, parliament, and government). As mentioned before, we construct a combined index by averaging trust toward state institutions and treat it as our dependent variable. Table 4 reports the results of the estimations.

According to the table, we document a negative association between relative deprivation and trust toward state institutions. This finding supports Hypothesis 3, which postulates a negative impact of relative advantage on trust toward state institutions. Unlike (Fischer & Torgler, 2013), we observe a positive association between relative advantage and trust in secular institutions. The relatively advantaged individuals may give credits to the state for their contributions.

### Table 2

Positional concerns and charitable donations.

<table>
<thead>
<tr>
<th>Region</th>
<th>Armenia LPM</th>
<th>Probit</th>
<th>Georgia LPM</th>
<th>Probit</th>
<th>Azerbaijan LPM</th>
<th>Probit</th>
</tr>
</thead>
<tbody>
<tr>
<td>Above Average</td>
<td>0.066***</td>
<td>0.066***</td>
<td>0.026</td>
<td>0.026</td>
<td>0.095***</td>
<td>0.111***</td>
</tr>
<tr>
<td></td>
<td>(0.015)</td>
<td>(0.025)</td>
<td>(0.026)</td>
<td>(0.028)</td>
<td>(0.031)</td>
<td>(0.022)</td>
</tr>
<tr>
<td>Below Average</td>
<td>-0.065***</td>
<td>-0.077***</td>
<td>-0.091***</td>
<td>-0.114***</td>
<td>-0.081***</td>
<td>-0.093***</td>
</tr>
<tr>
<td></td>
<td>(0.010)</td>
<td>(0.012)</td>
<td>(0.022)</td>
<td>(0.017)</td>
<td>(0.020)</td>
<td>(0.015)</td>
</tr>
<tr>
<td>Absolute Income</td>
<td>0.015***</td>
<td>0.016***</td>
<td>0.017**</td>
<td>0.019**</td>
<td>0.031***</td>
<td>0.034***</td>
</tr>
<tr>
<td></td>
<td>(0.004)</td>
<td>(0.005)</td>
<td>(0.008)</td>
<td>(0.007)</td>
<td>(0.008)</td>
<td>(0.008)</td>
</tr>
<tr>
<td>Controls</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Adj. R²</td>
<td>0.153</td>
<td>0.154</td>
<td>0.152</td>
<td>0.091</td>
<td>0.091</td>
<td>0.091</td>
</tr>
<tr>
<td>N</td>
<td>15,899</td>
<td>15,899</td>
<td>5624</td>
<td>5624</td>
<td>5800</td>
<td>5800</td>
</tr>
</tbody>
</table>

Note. This table reports results from Linear Probability and Probit Models (with robust standard errors) estimated on the pooled dataset as well as for each country separately. For Probit Models the marginal effects are reported. All the regressions include year, regional, and country dummies. Dependent variable: Binary variable that equals 1 if the respondent contributed to charity in the last 6 months and 0 otherwise. Independent variables: Absolute Income: Integer variable that indicates the household income of the respondent: high values indicate high income levels; Below Reference Group = 1 if the respondent is poorer than her reference group, 0 otherwise; Above Reference Group = 1 if the respondent is richer than her reference group, 0 otherwise. Significance Levels: * p < 0.1; ** p < 0.05; *** p < 0.01. For the extended table with controls (gender, age, household size, marital status, education, employment status, country in which the individual lives, region in which the individual lives, and the year in which the survey was conducted), please refer to Table C1 in S3 Appendix.
advantageous position versus the reference group. Why would relative advantage, rather than absolute income, be connected to giving credit to the state? If an individual is somewhat affluent, but she perceives others to be of the same wealth, she may again not realize her absolute wealth.

Table 3
Positional concerns and trust toward others.

<table>
<thead>
<tr>
<th>Region</th>
<th>OLS</th>
<th>Ordered Probit</th>
<th>Armenia OLS</th>
<th>Ordered Probit</th>
<th>Georgia OLS</th>
<th>Ordered Probit</th>
<th>Azerbaijan OLS</th>
<th>Ordered Probit</th>
</tr>
</thead>
<tbody>
<tr>
<td>Above Average</td>
<td>0.292***</td>
<td>0.124***</td>
<td>0.030</td>
<td>0.005</td>
<td>0.066</td>
<td>0.028</td>
<td>0.525***</td>
<td>0.235***</td>
</tr>
<tr>
<td>Below Average</td>
<td>-0.435***</td>
<td>-0.191***</td>
<td>-0.199*</td>
<td>-0.079*</td>
<td>-0.291***</td>
<td>-0.132***</td>
<td>-0.598***</td>
<td>-0.275***</td>
</tr>
<tr>
<td>Absolute Income</td>
<td>0.077***</td>
<td>0.034***</td>
<td>0.087*</td>
<td>0.040*</td>
<td>0.115***</td>
<td>0.050***</td>
<td>0.921</td>
<td>0.099</td>
</tr>
</tbody>
</table>

Note. This table reports results from OLS and Ordered Probit Models (with robust standard errors) estimated on the pooled dataset as well as for each country separately. All the regressions include year, regional, and country dummies. Dependent variable: Trust toward others measured on a scale of 1 (Fully distrust) to 5 (Fully trust). Independent variables. Absolute Income: Integer variable that indicates the household income of the respondent: high values indicate high income levels; Below Reference Group=1 if the respondent is poorer than her reference group, 0 otherwise; Above Reference Group=1 if the respondent is richer than her reference group, 0 otherwise. Significance Levels: * p<0.1; ** p<0.05; *** p<0.01. For the extended table with controls (gender, age, household size, marital status, education, employment status, country in which the individual lives, region in which the individual lives, and the year in which the survey was conducted), please refer to Table C2 in S3 Appendix.

Table 4
Positional concerns and combined vertical trust (Parliament, President, Executive Government).

<table>
<thead>
<tr>
<th>Region</th>
<th>OLS</th>
<th>Ordered Probit</th>
<th>Armenia OLS</th>
<th>Ordered Probit</th>
<th>Georgia OLS</th>
<th>Ordered Probit</th>
<th>Azerbaijan OLS</th>
<th>Ordered Probit</th>
</tr>
</thead>
<tbody>
<tr>
<td>Above Average</td>
<td>0.272***</td>
<td>0.324***</td>
<td>0.100*</td>
<td>0.093*</td>
<td>0.229***</td>
<td>0.277***</td>
<td>0.361***</td>
<td>0.448***</td>
</tr>
<tr>
<td>Below Average</td>
<td>-0.295***</td>
<td>-0.315***</td>
<td>-0.217***</td>
<td>-0.202***</td>
<td>-0.156**</td>
<td>-0.179***</td>
<td>-0.397***</td>
<td>-0.424***</td>
</tr>
<tr>
<td>Absolute Income</td>
<td>-0.022**</td>
<td>-0.029***</td>
<td>0.009</td>
<td>0.009</td>
<td>-0.003</td>
<td>-0.004</td>
<td>-0.064***</td>
<td>-0.076***</td>
</tr>
</tbody>
</table>

Note. This table reports results from OLS and Ordered Probit Models (with robust standard errors) estimated on the pooled dataset as well as for each country separately. All the regressions include year, regional, and country dummies. Dependent variable: Combined Trust toward state institutions measured on a scale of ‘1’ (fully distrust) to ‘5’ (fully trust). Independent variables. Absolute Income: Integer variable that indicates the household income of the respondent: high values indicate high income levels; Below Reference Group=1 if the respondent is poorer than her reference group, 0 otherwise; Above Reference Group=1 if the respondent is richer than her reference group, 0 otherwise. Significance Levels: * p<0.1; ** p<0.05; *** p<0.01. For the extended table with controls (gender, age, household size, marital status, education, employment status, country in which the individual lives, region in which the individual lives, and the year in which the survey was conducted), please refer to Table C3 in S3 Appendix.

The analysis of marginal effects of relative standing confirm our findings. For the entire region (as well as for separate countries), ceteris paribus, relative deprivation contributes to having a higher probability of reporting a low level of trust toward others and trust in secular institutions. The results for relative advantage mirror the results for relative deprivation. Nonetheless, please note, that country by country results are not always statistically significant (this is aligned with our findings reported in Tables 3 and 4). The reader can refer to Figures E2, E3 in S5 Appendix for more details. Please note that while for Armenia it is positive, albeit small and insignificant, for Azerbaijan the coefficient is negative and significant. Since Azerbaijan is an autocratic regime, most likely richer (hence more informed and educated) individuals trust the secular institutions less.

5.3. IV estimation and further robustness checks

It can be argued that in the estimations that we discussed so far perceives others to be of the same wealth, she may again not realize her absolute wealth.

Regarding the negative coefficient of absolute income, this result is mainly driven by Azerbaijan (for Armenia the coefficient is positive and insignificant, while for Georgia it is negative, albeit small and insignificant). Since Azerbaijan is an autocratic regime, most likely richer (hence more informed and educated) individuals trust the secular institutions less.

6 The analysis of marginal effects of relative standing confirm our findings.

7 Regarding the negative coefficient of absolute income, this result is mainly driven by Azerbaijan (for Armenia the coefficient is positive and insignificant, while for Georgia it is negative, albeit small and insignificant). Since Azerbaijan is an autocratic regime, most likely richer (hence more informed and educated) individuals trust the secular institutions less.
charity donations may induce an individual to feel materially disadvantaged relative to her reference group. Ideally, the extant literature should rule out such concerns. For instance, adopting an IV approach, (Fischer & Torgler, 2013) illustrate that causality runs from relative income to social capital and not the other way around. As for the relationship between social preferences and positional concerns, there is bulk of theoretical and empirical evidence on positional concerns, which has generally found that an individual strives to outperform the members of her reference group (e.g., Frank, 1985, 2008) and suffers utility losses whenever she lags behind (e.g., Antinyan, 2016). Following this line of reasoning, we believe that an individual would have not made a (large) charity donation if it had resulted in actual or perceived relative deprivation. Lastly, as mentioned in Section 4.3 private donations to charities by ordinary individuals living in Armenia, Azerbaijan, and Georgia are not common; meanwhile donating small amounts of money through SMS or giving money to beggars is rather common. Given that we most likely deal with small charity donations, it is quite unlikely that an individual feels materially disadvantaged relative to her reference group after making a donation.

To fully rule out the reverse causality problem, we cannot use a conventional IV approach as the survey does not contain appropriate variables to determine respondent-level instruments. Under these circumstances, we estimate an IV model that uses heteroskedasticity-based instruments generated through Lewbel’s method (Lewbel, 2012). Recently, this method has been successfully applied by several authors, including (Kelly & Markowitz, 2007; Emran & Hou, 2008; Fortin & Ragued, 2017; Denny & Oppedisano, 2013; Millimet & Roy, 2016). Below, we detail the approach to enhance the clarity of the text.

The generic model can be expressed in the following form:

\[ Y_1 = X' \beta_1 + Y_2 X_1' + \epsilon_1 \]  
(2)

\[ Y_2 = X' \beta_2 + \epsilon_2 \]  
(3)

where \( Y_2 \) is an endogenous variable (\( \text{Corr}(Y_2, \epsilon_1) \neq 0 \)), \( Y_1 \) is an outcome variable, \( X \) is a vector of exogenous variables, \( \epsilon_1 \) and \( \epsilon_2 \) are potentially correlated error terms. In the system of equations above, (2) is the structural equation while (3) is the first-stage equation. In a standard two-stage model, some of the elements in \( X \) are not part of the structural Eq. (2). However, these variables have the power to explain the endogenous variable \( Y_2 \) and are used as instruments in Eq. (3). In our case, such instruments are missing. To overcome the problem, (Lewbel, 2012) provides a solution to the identification of the parameters, which requires the following conditions to be held:

1. \( E(X \epsilon_1) = 0 \). This is the standard exogeneity condition for \( X \).
2. \( \text{Cov}(Z, \epsilon_2) \neq 0 \). In other words, heteroskedasticity should be present in the first-stage model (testable by a standard Breusch – Pagan test).
3. \( \text{Cov}(Z, \epsilon_1 \epsilon_2) = 0 \). This condition is somewhat analogous to the excludability condition of instruments from the structural equation.

\( Z \) is either a subset of exogenous variables \( X \) or equal to \( X \). In our case we use a subset of variables. The list of variables, from which the instruments are generated, can be found in the captions under Tables D1–D3 in S4 Appendix. Estimations were performed using iivreg2h STATA module (Baum & Schaffer, 2021) following the estimation steps by Ventura (2018).

We run the estimations with one endogenous variable, in line with the original (Lewbel, 2012) method. In this regard, (Lewbel, 2018) shows that assumptions required for (Lewbel, 2012) estimator can be satisfied when the endogenous regressor is binary. For this purpose, we use the variable capturing the perception of the relative standing of the respondent’s household in its original form measured on a scale of “1” (Very Poor) to “5” (Very Good).

Tables D1–D3 in S4 Appendix illustrate the results of the OLS, LPM and the corresponding Lewbel specifications. For the sake of comparability of the results across different models, the relative standing variable is used in its original form in OLS and LPM models as well. The regression coefficients obtained using Lewbel’s approach are consistent with those from OLS and LPM models, which reinforces the argument that positional concerns can have causal impact on the dependent variables under scrutiny. Furthermore, the sign and the significance of the coefficients are aligned with the main results of the paper. The discussion of the relevant model identification tests is reported in S4 Appendix. We also run estimations with two endogenous binary variables for relatively deprived and advantaged groups. The two Lewbel specifications yield qualitatively comparable results.

On top of the IV estimation, we further check the robustness of the results discussed in Section 5. First, based on the graphical evidence in Fig. 5, one can argue that there is a positive correlation between absolute income and relative advantage. In other words, those who declare high household income may perceive their households to have relative advantage vis-à-vis the reference group. This phenomenon is especially vivid in the last two income groups (around 40–60% of the individuals in these income groups report being relatively advantaged). The inclusion of both absolute and relative income variables in the main analysis is meant to ensure the separate identification of the relative and the absolute income effects. Nonetheless, if these two variables are highly correlated, the separate identification of the two effects will not be possible. To assure the validity of our estimations, we drop the households that are in the last two income groups and re-estimate the econometric models discussed in Section 4. In this way, we get rid of the potentially problematic income groups in which the relative and absolute incomes overlap. If this exercise yields a similar relationship between relative position and the outcome variables, we will have more fertile ground to argue that relative positions matter. The estimation results are in line with our main conclusions.

Second, instead of using integer numbers from one to eight to indicate various income levels, we construct an alternative absolute income variable in line with (McBride, 2001). More specifically, we make household’s income equal to the natural logarithm of the middle income of its income category. For example, if household’s income belongs to USD $1999-100$ income range, then we set its income equal to the natural logarithm of USD 75.5. To accomplish this exercise, we drop the households with either 0 income (logarithm of 0 does not exist) or an income higher than USD 1200 (an upper bound does not exist). Our main conclusions remain intact, and all above mentioned results are available upon request.

6. Conclusion

In this paper, we study how one’s perceived positional advantage and deprivation relate to one’s recalled and self-reported charity donation decisions and social capital in the South Caucasus. These variables under scrutiny are crucial for redistributive policies (Fong, Bowles, & Gintis, 2005) and macroeconomic development (Guiso et al., 2004), and they can be of great relevance for policymakers not only in the region, but also worldwide.

We illustrate a negative relationship between relative deprivation and charity donation decisions, concluding that an individual’s perceived positional disadvantage can have detrimental consequences for her charitable behavior. Interestingly, we also illustrate a positive link between relative advantage and charity donation decisions, concluding that an individual’s perceived positional advantage may enhance her charitable behavior. Regarding the relationship between positional concerns and social capital, we illustrate that on the one hand relative disadvantage has a social capital-destroying effect, on the other
hand relative advantage has a social capital enhancing effect.

Our results demonstrate that relative economic deprivation is associated with lower levels of charitable giving and social capital. Furthermore, our data suggest that relatively deprived individuals outnumber relatively advantaged individuals in the region. Combined, these results indicate that mitigation of relative deprivation can on aggregate contribute to higher frequency of charitable giving and trust. This serves as an additional convincing argument to undertake measures to equalize income distribution in the countries under consideration. For example, in such cases, Duesenberry (1949) suggests that progressive income taxation enhances allocational efficiency. In a similar vein, Fortin & Raguel (2017) proposes a progressive consumption tax, to ‘... mould the frame of reference in mutually beneficial ways’ (p.1844).

We would also like to note several limitations of our study. First, caution is required when working with sensitive survey questions such as helping others, trusting others, or trusting the state. A respondent may not report one’s true behavior/preferences so as not to harm one’s identity because of concerns of social approval or fear of being punished (in the case of trust in secular institutions). The level of this concern can be rather heterogeneous among the respondents. In sum, even though such questions are considered important sources of information and are actively used (e.g., Korndörfer et al., 2015; Fischer & Torgler, 2013), the reader should exhibit caution when generalizing the findings. Second, we utilized repeated cross-sectional data and an IV approach to establish a causal link between the relative position, charitable donations, and social capital. More causal evidence in this direction is required through laboratory or field experiments. Third, the findings can be highly region-specific and more research is required to establish a link between relative position and their donation levels. For the sake of robustness, the researchers may also want to use different survey questions that measure social preferences as well as data collection methodologies (e.g., laboratory or field experiments).

Supplementary materials
Supplementary material associated with this article can be found, in the online version, at doi:10.1016/j.socec.2022.101929.

References