Incidence, prevalence and co-occurrence of autoimmune disorders, trends over time and by age, sex and socioeconomic status. A population-based study in 22 million individuals.

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Abstract

Background: A rise in the incidence of selected autoimmune disorders has been described, however, contemporary estimates of the overall incidence of autoimmune diseases and trends over time, are scarce and inconsistent.

Methods: We used linked primary and secondary electronic health records of 22 million individuals from the Clinical Practice Research Datalink (CPRD), a cohort that is representative of the UK population in terms of age and sex. We calculated age-sex-standardised incidence and prevalence of 19 autoimmune disorders (AID) from 2000 to 2019 and used negative binomial regression models to investigate temporal trends and variation by age, sex, socioeconomic status, season of onset and geographical region. To characterise co-occurrence of autoimmune diseases, we calculated incidence rate ratios, comparing incidence rates of comorbid autoimmune disease among patients with a first (index) autoimmune disease with incidence rates in the general population, using negative binomial regression models, adjusted for age and sex.

Findings: Among the 22,009,375 individuals included in the study, we identified a total of 978,872 patients with a new diagnosis of at least one autoimmune disease between 2000 and 2019 (mean (SD) age: 54.0 (21.4) years, 64% women). Over the study period, age-standardised incidence rates of autoimmune diseases increased by 4%, similarly for men and women. The largest increases were seen in Graves’ disease, coeliac disease and Sjogren’s syndrome, for which incidences have doubled over the past two decades. Two conditions exhibited a significant decrease in incidence (Hashimoto’s thyroiditis and pernicious anaemia). Taken together the 19 autoimmune disorders examined affected 10.2% of the population over the study period (13.1% of women, 7.4% of men). A socioeconomic gradient was evident across several diseases, including Graves’ disease, pernicious anaemia, rheumatoid arthritis, and systemic lupus erythematosus. Seasonal variations were observed for type 1 diabetes (more commonly diagnosed in winter) and vitiligo (more commonly diagnosed in summer), and regional variations were observed for a range of conditions. Autoimmune disorders were commonly associated with each other, particularly Sjogren’s, systemic lupus erythematosus and systemic sclerosis. Patients with type 1 diabetes also had significantly higher rates of Addison’s, coeliac, and thyroid diseases, and multiple sclerosis stood out as having low rates of co-occurrence with other autoimmune diseases.

Interpretation: Autoimmune diseases affect about one in ten individuals. Their burden continues to increase over time, albeit at varying rates across individual diseases. The socioeconomic, seasonal, and regional disparities observed among several autoimmune disorders, implicate environmental factors in disease pathogenesis. The interrelations between autoimmune diseases are commensurate with shared pathogenetic mechanisms or predisposing factors, particularly among connective tissue diseases and among endocrine diseases.

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Keywords: autoimmune disorders, immune mediated inflammatory diseases, incidence, prevalence, co-occurrence, cohort study, epidemiology, CPRD.
Research in Context

Evidence before this study

We searched Pubmed and Embase for reports published between 1 January 2000 to 30 July 2022 related to “autoimmune disorders” (any of the 19 individual conditions investigated) and “incidence”, reviewed references of clinical practice guidelines and consulted with experts for relevant studies. Most studies investigated one autoimmune disorder at a time, and generally the more common autoimmune disorders, such as type 1 diabetes or psoriasis.

Studies generally relied on a small number of cases and presented different designs, case identification and diagnostic methods, rendering adequate synthesis and the calculation of pooled estimates and temporal trends difficult. Evidence was particularly scarce for rarer autoimmune disorders. We found no study that reported large-scale disease incidence and temporal trends of autoimmune disorders as a group of conditions.

Added value of this study

We present standardised incidence rates derived from a large, representative, general population cohort, setting a baseline for international comparison, monitoring of prevention strategies and for the design of public health policies. Temporal trends across a broad range of autoimmune diseases do not support the idea of an epidemic of autoimmunity at present day, and provide valuable reference rates for future studies investigating population-level impact of newly-introduced risk factors, such as the covid-19 pandemic.

We provide robust evidence of socioeconomic, seasonal, and regional disparities for several autoimmune diseases (particularly Graves’ disease, pernicious anaemia, rheumatoid arthritis, systemic lupus erythematosus, and type 1 diabetes). Such variations are unlikely to be attributable to genetic differences alone and suggest involvement of potentially modifiable risk factors in the pathogenesis of autoimmune diseases.

We further demonstrate important interrelations between many autoimmune diseases, and confirm that co-occurrence of autoimmune disease is common, yet orders of magnitude differ widely between diseases. Associations were highest amongst connective tissue diseases, for patients with type 1 diabetes and Addison’s disease, coeliac and thyroid diseases. More generally, increased risk of developing Addison’s disease was observed following almost every autoimmune disease investigated. Multiple sclerosis stood out as having low rates of co-occurrence with other autoimmune diseases, suggesting a distinct pathophysiology.

Implications of all the available evidence

The burden of autoimmune disorders appears higher than previous estimates, and continues to increase over time, albeit at varying rates across individual diseases.

Socioeconomic, seasonal, and regional disparities in disease incidence point to potentially preventable factors involved in the pathogenesis of autoimmune diseases. Co-occurrences of diseases point to common genetic and environmental risk factors that interact and operate variably across these diseases.

Further research is needed to fully elucidate pathophysiological mechanism underlying the associations observed in this study.
Introduction

Autoimmune diseases arise when immune dysregulation causes host tissue damage. A wide range of autoimmune diseases are described that present with variable age of onset, tissue distribution and clinical and functional impact. Most of these diseases are incurable and require lifelong treatment.

Adequate public health and service delivery planning requires reliable information about contemporary population-level disease incidence. However, estimates of autoimmune disease incidence rates and their temporal trends, even in high-income countries, are scarce and inconsistent. Selected autoimmune disorders, such as type 1 diabetes, are reported to have increased over the past several decades, raising the question as to whether the overall incidence of autoimmune disorders is on the rise, driven perhaps by common environmental factors or behavioural changes. Even for type 1 diabetes, whose incidence is among the best studied within autoimmune diseases, reports rely on relatively small cohorts, and estimates vary by a factor 10 between studies in Europe alone. For many other autoimmune diseases, evidence concerning disease incidence and prevalence is more limited. The relatively modest absolute numbers of patients affected by individual autoimmune diseases presents as a major challenge to investigators, and hinders adequate synthesis across studies. As a result, reliable estimates of disease incidence and how they evolve over time, particularly as pertains to autoimmune diseases as a group, are not available.

Commonalities and differences between individual diseases also remain poorly understood and continue to be subject to much research. While emerging evidence has suggested that autoimmune diseases tend to co-occur within individuals, large-scale investigations across a broad spectrum of autoimmune diseases that could provide clues about shared pathogenesis and risk factors are not currently available.

To address these knowledge gaps, we analysed a large longitudinal database of primary and secondary care records that provides information on millions of individuals’ diagnoses with several years of follow-up. We investigated the incidence and prevalence for 19 of the most common autoimmune disorders (AID): Addison’s disease; ankylosing spondylitis; coeliac disease; childhood-onset type 1 diabetes; Graves’ disease; Hashimoto’s thyroiditis; inflammatory bowel disease (Crohn’s disease or ulcerative colitis); multiple sclerosis; myasthenia gravis; pernicious anaemia; polymyalgia rheumatica; primary biliary cholangitis; psoriasis; rheumatoid arthritis (including its specific subtypes such as Still’s disease, Caplan syndrome, rheumatoid spondylitis, and others); Sjogren’s syndrome; systemic lupus erythematosus; systemic sclerosis; vasculitis; and vitiligo. For selected conditions, we performed sensitivity analyses using more restrictive disease definitions and present these in the appendix. Diseases were considered individually and as a composite outcome of all AID combined. For the combined analyses, we calculated primary incidence (first recorded AID, reflecting the number of patients affected by AIDs) and cumulative incidence (all recorded AIDs, reflecting the cumulative number of AID diagnoses).
Although some of these diseases remain debated in terms of their autoimmune aetiology and may be more appropriately described as 'immune mediated inflammatory diseases', to assist readability we refer to this group of conditions as 'autoimmune diseases'.

For each condition, we identified diagnoses from primary care or secondary care records based on diagnostic codes from hospital/death (International Classification of Diseases, tenth revision (ICD-10)) and primary care (Read\textsuperscript{14} and SNOMED codes\textsuperscript{15}) coding schemes, and, for selected conditions, prescriptions of certain drugs (appendix). Specifically, individuals with childhood-onset type 1 diabetes were identified as those with at least one diagnostic code referring to type 1 or insulin-dependent diabetes, at least 1 insulin prescription, and aged 19 years or less at first diagnosis. Individuals with Hashimoto thyroiditis were identified as those with at least one levothyroxine prescription and no history of hyperthyroidism, pituitary disease, thyroid surgery, or thyroid-altering medication (amiodarone, lithium, sodium valproate, carbimazole, propylthiouracil, thalidomide, sunitinib). When conflicting arthritis diagnoses (ankylosing spondylitis and rheumatoid arthritis) were recorded in the same individual, the last recorded diagnosis was used. Incident diagnoses were defined as the first record of that condition in primary or secondary care records from any diagnostic position.

**Study population**

Included in the study were men and women with records labelled as ‘acceptable’, approved for HES and ONS linkage, and registered with their general practice for at least 12 months during the study period (01/01/1998 to 30/06/2019). For incidence calculations, we excluded all individuals who had a diagnosis of the disease of interest prior to study start date (1 January 2000), or within the first 12 months of registration with their general practice.

**Covariates**

Smoking status and body mass index (BMI) were abstracted from electronic health records as the most recent measurement within 2 years prior to diagnosis. Socioeconomic status was defined as the Index of Multiple Deprivation (IMD) 2015 quintile,\textsuperscript{16} a composite measure of relative deprivation at a small area level, covering an average population of 1500 people, ranked in ascending order of deprivation score and grouped in equal fifths, with quintiles 1 and 5 representing the least and most deprived areas, respectively. Data on sex are as reported by the patient when they registered with their general practitioner. Ethnicity was extracted from both primary and secondary care records. When ethnicity differed between primary and secondary care records, secondary care data was used.

**Statistical analyses**

Baseline characteristics are presented as frequencies (%) for categorical data, medians and interquartile range (IQR) for non-normally distributed continuous data, or means and standard deviation (SD) for normally distributed continuous data, over the whole autoimmune disease cohort and stratified by sex, socioeconomic quintile, and period of diagnosis. Number and percentage of records with missing data are displayed for variables with missing entries.

Observed incidence rates were computed by dividing the number of incident cases by the number of patient-years in the cohort. Category-specific rates were computed separately for subgroups of age, sex, socioeconomic status, region, calendar year of diagnosis and season of diagnosis. Winter was defined as the period from January to March, and Summer as June to August. Time at risk was calculated to start at the latest of patient’s registration date plus 12 months, birthyear, or study start date; and stopped at the earliest of death, transfer out of practice, practice’s last collection date, incidence of the disease of interest, or study end date. Observed prevalence rates were computed considering all patients ever diagnosed with autoimmune disease (numerator) among patients alive and registered with a general practitioner on June 30\textsuperscript{th} in each year (denominator). To allow comparison with other studies, we further calculated incidence and prevalence rates of childhood onset type 1 diabetes restricting the denominator to those aged 19 or less (appendix).

Standardised rates were computed by applying direct age and sex standardisation\textsuperscript{17} to the 2013 European Standard Population\textsuperscript{18} using 5-year age bands up to 90 years of age.
Negative binomial regression models were used to examine overall and category-specific incidence rate ratios (IRR) and corresponding 95% confidence intervals (CI). Models were adjusted for calendar year, age (categorised into five years age-bands), sex, socioeconomic status and region. Negative binomial models were chosen over Poisson models to account for overdispersion in the data. Sensitivity analyses comparing Poisson and negative binomial models showed very similar results.

To characterise co-occurrence of autoimmune diseases, we have run a series of cohort studies to calculate incidence rate ratios for the development of a second (comorbid) autoimmune disease according to disease status of a first (index) autoimmune disease. Incidence rate ratios were calculated using negative binomial regression models adjusted for age, sex and calendar year. We performed separate analyses for each pair and sequence of autoimmune diseases, following methods by Somer et al.\textsuperscript{7} For these analyses, time at risk started at the latest of patient registration plus 12 months or diagnosis of the index disease, and stopped at the earliest of the incidence of the comorbid autoimmune disease, death, the patient ceasing registration with the general practice, the practice’s last collection date, or the end of the study.

Study findings are reported in accordance with the REporting of studies Conducted using Observational Routinely-collected health Data (RECORD) recommendations.\textsuperscript{19} Statistical analyses were performed in R (version 3.6.2) and validated with SAS (version 9.4).

Role of the funding source

The funders of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report.

Results

A total of 22 009 375 individuals contributed data between 01/01/2000 and 31/12/2019 with 135 691 152 patient-years of follow-up. Among those, we identified 1 123 789 new diagnoses of autoimmune diseases, affecting a total of 978 872 patients. The mean (SD) age at AID diagnosis was 54.0 (21.4) years, and 64% of patients were women (Table 1).

Temporal trends

The number of patients newly diagnosed with one (or more) autoimmune diseases increased only modestly over time (IRR comparing 2017-2019 vs 2000-2002: 1.04 [1.00, 1.09]). The number of new autoimmune disease diagnoses increased by 22%, largely due to an increasing number of secondary autoimmune disease diagnoses among patients already affected by a first autoimmune disease (IRR comparing 2017-2019 vs 2000-2002: 1.22 [1.18, 1.28]). Coeliac, Graves’, and Sjogren’s syndrome showed the greatest increases, whereas Hashimoto’s thyroiditis and pernicious anaemia declined modestly over the study period (Figure 1). The observed increase in coeliac disease, Graves’ disease, and Sjogren’s syndrome was largely driven by a higher number of diagnoses in women. The increase in Graves’ disease was largely driven by the very old (80 years and older), whereas the increase in coeliac disease was largely driven by diagnoses in the very young (up to 40 years of age) (appendix).

Age at diagnosis

Autoimmune disorders developed over the whole life spectrum, from the first to the 95\textsuperscript{th} year of life. Median age at first autoimmune disease presentation varied greatly among individual autoimmune disease. For many conditions, incidence increased with age; this was the case for Graves’ disease, pernicious anaemia, and rheumatoid arthritis. Only six conditions were commonly diagnosed before the age of five, these were: Addison’s, coeliac disease, type 1 diabetes, psoriasis, vitiligo and vasculitis (largely due to Henoch-Schonlein purpura, Kawasaki disease, and glomerulonephritis). For others, such as multiple sclerosis, psoriasis, and lupus, the incidence peaked in the middle age. Finally, some conditions presented a bi-modal age-distribution with a peak in childhood or early adulthood and another one later in life; this was the case for coeliac disease, inflammatory bowel disease and vasculitis (Figure 2).

Incidence by sex
Most autoimmune disorders were more common in women than men (IRR for women compared to men: 1.74 [1.72, 1.77] for all diseases combined). Thyroid disorders, Sjogren’s syndrome, lupus, and systemic sclerosis had the highest incidence rate ratio in women compared to men. Only three diseases were more common in men than women, namely ankylosing spondylitis; type 1 diabetes; and myasthenia gravis (Figure 3).

Incidence by socioeconomic status

Overall, the most deprived socioeconomic groups had a higher incidence of autoimmune diseases (IRR for the most deprived compared to least deprived quintile: 1.14 [1.11, 1.16] for all autoimmune diseases combined). A marked socioeconomic gradient was visible across several individual diseases, including Graves’ disease, pernicious anaemia, rheumatoid arthritis, and lupus. For other diseases, such as Hashimoto’s thyroiditis and inflammatory bowel disease, no difference was observed among socioeconomic groups despite relatively large number of cases; and for coeliac and polymyalgia rheumatica, disease incidence was highest in the least deprived group (Figure 3).

Regional differences

Overall, variation by geographic region was limited. Notable exceptions were polymyalgia rheumatica, type 1 diabetes, and coeliac disease, which were considerably more common outside of the capital city, as well as pernicious anaemia, which appeared to be more common in northern regions. Sjogren’s syndrome, lupus and vitiligo, on the other hand, presented lower incidence rates outside of London (Figure S4).

Seasonal differences

Most autoimmune diseases were diagnosed throughout the year with no significant differences between winter and summer months. Seasonal variation was observed only for type 1 diabetes, which was more commonly diagnosed in the winter months (January to March) compared to summer months (June to August), and for vitiligo, which was more likely to be diagnosed during the summer (Figure S5).

Co-occurrence of autoimmune diseases

Autoimmune disorders were commonly associated with each other. Increased risk of developing a second autoimmune disorder was seen across many autoimmune diseases, but orders of magnitude differed widely between diseases. Associations were generally highest among connective tissue diseases, particularly between Sjogren’s, systemic lupus erythematosus and systemic sclerosis. Patients with type 1 diabetes also had significantly higher rates of Addison’s, coeliac and thyroid diseases. More generally, Addison’s disease occurred at a considerably higher incidence among patients with pre-existing autoimmune disease than in the general population. Multiple sclerosis stood out as having low rates of co-occurrence with other autoimmune diseases and even showed an inverse association with some autoimmune disorders, but this was only significant for vitiligo (Figure 5, Table S1).

Prevalence trends over time

Together the 19 autoimmune disorders investigated in this study affected 10.2% of the population over the study period (13.1% of women, 7.4% of men). Age- and sex- standardised prevalence increased over time from 7.7% in 2000-2002 to 11.0% in 2017-2019 (RR comparing 2017-2019 vs 2000-2002: 1.41 [1.37, 1.44]).

Discussion

Our large-scale population-based study provides several novel insights into the burden of autoimmune disorders, its variation over time, by individual diseases, and patient subgroups of age, sex and socioeconomic status. Our findings confirm and extend evidence from previous studies demonstrating an increasing incidence of several autoimmune disorders, and shows that the increase was particularly pronounced for Graves’ disease, coeliac disease, and rheumatic disorders. Our results reflect disease incidence based on diagnostic criteria, screening practices, availability and accuracy of diagnostic tests in place at that precise time and hence must be interpreted within this context. As such and in consideration of the increasing awareness for some of these conditions, improved coding practices, and earlier recognition of these conditions over the past two decades, the observed increase remains modest.
The epidemiology of type 1 diabetes is perhaps the best studied of all. Previous studies had reported varying estimates and trends over time, and several surveys from Scandinavian countries have reported steep increases since the 1950s with a plateau since mid 2000. In our study, we observed only a modest increase in disease incidence over the past two decades. Overall estimates were comparable with estimates from a recent simulation study on the incidence of type 1 diabetes in Western Europe and similar to those reported by the International Diabetes Federation Atlas for the UK, but lower than those recently reported in northern European countries. Another key finding was the reduction in Hashimoto’s thyroiditis over time, which could be due to more careful initiation of levothyroxine in older persons with subclinical hypothyroidism following trials demonstrating limited benefits in this population. Pernicious anaemia also showed an apparent decline in incidence over time, a decline that appears to coincide with increased use of dietary supplements over the same period, and possibly a more widespread recognition of other causes of vitamin B12 deficiency such as helicobacter pylori, although no causal inference can be made from our data.

The observed increase in incidence of rheumatic diseases will likely have important implications for health services and already substantial medication expenditures linked to biologics. Improved coding practices during the study period, the introduction, in 2013, of a quality audit in primary care (so-called Quality and Outcomes Framework, QOF) rewarding general practitioners for maintaining a register and evaluating cardiovascular/fracture risks in patients with rheumatoid arthritis, and novel classification criteria for ankylosing spondylitis, are all likely to have contributed to the observed trend. The increase in incidence of axial spondyloarthritis in women starting around the time of publication of ASAS (Assessment of SpondyloArthritis international Society) classification criteria - introducing the concept of non-radiographic axial spondyloarthritis - is consistent with other studies indicating that whilst ankylosing spondylitis is male-predominant, non-radiographic axial spondyloarthritis is similarly common in women and men.

Our stratified analyses by socioeconomic status, region, and seasonal variations in disease incidence provide further insights into the possible role of environmental factors in the development of autoimmune diseases. Four diseases - Graves’ disease, pernicious anaemia, rheumatoid arthritis, and systemic lupus erythematosus - presented a clear socioeconomic gradient with those in the most deprived group up to 50 % more likely to develop the disease than their affluent counterpart. Such socioeconomic disparities could indicate that diet, smoking, obesity, air pollution, or other currently unrecognised environmental exposures might play a role in the development of these diseases. Two conditions – coeliac disease and polymyalgia rheumatica – presented an inverse socioeconomic gradient, a phenomenon rarely observed in public health research, and which could be linked to increased awareness and testing for these diseases in more affluent populations or indeed lifestyle differences.

Seasonal variations on the other hand were limited. Vitiligo was more commonly diagnosed in the summer, perhaps due to increased visibility of depigmented skin areas in summer months, and type 1 diabetes was more commonly diagnosed in the winter and outside of the capital region, a finding compatible with hypotheses of viral triggers, diet, higher weight, or ethnicity playing a role in the disease’s pathogenesis. Other regional variations, such as coeliac disease more common outside of the capital, remain unexplained for now. While numerous reports have linked smoking to the incidence of certain autoimmune diseases such as rheumatoid arthritis, multiple sclerosis, or psoriasis, we did not observe a decline in these diseases’ incidences despite considerable reduction in smoking prevalence over the same period. This could be due to a parallel increase in other risk factors, such as obesity, over the same period, confounding via socioeconomic status, or currently unknown reasons.

Finally, co-occurrences of diseases provide valuable insights into a possibly common aetiology across some of these conditions. For example, we found high rates of co-occurrence among a range of connective tissue diseases, particularly Sjogren’s syndrome, systemic lupus erythematosus, and systemic sclerosis, regardless of diagnostic sequence and after accounting for age and sex. Interestingly polymyalgia rheumatica did not have much association with other connective tissue disorders, except for vasculitis. Myasthenia gravis and multiple sclerosis also tended to co-occur. However, the association was weak and given the nature of our data, which is based on routine clinical practice data, we cannot rule out the possibility of inaccurate diagnoses or coding.

Our study further provides robust evidence for an increased incidence of a range of autoimmune diseases, particularly coeliac disease and thyroid disorders, among patients with type 1 diabetes. Such observations have been reported before and may point to overlapping genetic or environmental risk factors as well as...
people with a first autoimmune disease being more likely to undergo screening for these conditions. While we did not observe associations between childhood-onset diabetes and diseases typically occurring at older age groups, our data was limited by follow-up duration, so that we cannot exclude that such associations might exist. Similarly, in our study, psoriasis presented limited co-occurrence with autoimmune diseases, possibly because of the large proportion of mild cases typically observed among cohorts based on primary care data.

Another consideration is whether treatment of an index autoimmune disease affects risk of developing a comorbid disease. Although nonspecific immunosuppressives could in theory decrease risk of a comorbid autoimmune disease, targeted therapies could have differential effects. For example, an intriguing finding was the high incidence of Addison’s disease following almost every other autoimmune disorder, which could perhaps be related to glucocorticoid induced adrenal insufficiency (which is typically recorded as Addison’s disease in the UK). Such a mechanism might also explain the higher rates of cardiovascular diseases observed in patients with Addison’s disease in this same cohort. Nevertheless, given the relatively small number of patients with Addison’s disease, these results must be interpreted with caution. Overall, this research confirms that some autoimmune diseases co-occur with one another at a rate greater than expected by chance or surveillance bias alone, but it reveals that this phenomenon is not generalized across all autoimmune diseases.

A major strength of this study is the selection of a statistically powerful data source with over 130 million person-years of data to investigate the incidence, prevalence, and co-occurrence of autoimmune disorders. The very large size of our cohort allowed us to perform stratified analyses of unprecedented granularity, over a broad spectrum of conditions, as well as allowing examination of the influence of age, sex, and socioeconomic status, as well as trends over 20 years. The use of routinely reported diagnoses also captures the burden of disease as experienced by physicians and health services, and likely increases the generalizability of findings. One of the key limitations of our study was the limited diversity in ethnic backgrounds in our cohort and the unavailability of additional variables potentially relevant to autoimmune disease pathogenesis, such as smoking, body mass index or blood biomarkers such as vitamin D deficiency. The non-randomised design of our study further meant that we were unable to account for the effect of drug therapy, such as antirheumatic drugs, steroids, or biologics, on the incidence of a second autoimmune disease. Research using electronic health records is also reliant on the accuracy of clinical coding carried out during consultations and hospital admissions. The validity of diagnoses underlying our study has been carefully assessed and was considered appropriate in light of the over two hundred independent studies that have investigated the validity of diagnoses recorded in CPRD which reported an average positive predictive value of about 90% for a broad range of conditions. These include recent studies on several of the diseases studied here, which have demonstrated that algorithms based on diagnostic codes perform well at identifying patients with these conditions in primary care records. Generally, observed age distributions were consistent with prior studies and add to the validity of our approach. Nevertheless, our results, particularly trends over time and co-occurrence of diseases must be interpreted within the context and limitations of routinely collected health records data and the possibility that some level of miscoding is present. Finally, a large number of tests and subgroup analyses within one study must also be interpreted with adequate caution.

Our findings present an important new piece in the puzzle of autoimmune disease aetiology, a group of conditions that are apparent in near 10% of the population and which consume considerable health resources. The socioeconomic, seasonal, and regional disparities observed among several autoimmune disorders, point towards involvement of environmental factors in the pathogenesis of selected autoimmune diseases. The interrelations between many autoimmune diseases further suggest a shared pathogenesis, particularly among connective tissue diseases as well as between diabetes, coeliac and thyroid disorders. To this day, the exact causes of many of the autoimmune diseases studied here remain unknown and require further research.

Contributions

NC, GC and JCM conceived and designed the study. NC, GM, and GV designed the statistical analysis plan and NC performed the statistical analysis. All authors contributed to interpreting the results, drafting the manuscript, and the revisions. NC and GC had full access to the data in the study and had final responsibility
for the decision to submit for publication. NC, GM, GV, and JYV had permission to access the data and NC and GM verified the data (CRPD requests that access to data is given only as absolutely necessary). All authors gave final approval of the version to be published and accept responsibility to submit the manuscript for publication.

Declarations of interest

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Data sharing

Access to CPRD data is subject to a license agreement and protocol approval process that is overseen by CPRD’s Independent Scientific Advisory Committee (ISAC). A guide to access is provided on the CPRD website.

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Table 1: Baseline characteristics of patients with incident autoimmune disease during 2000-2019.

<table>
<thead>
<tr>
<th>Age at diagnosis (years)</th>
<th>All patients (N = 978,872)</th>
<th>Sex</th>
<th>Socioeconomic status quintile</th>
<th>Time period</th>
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<tr>
<td>Mean (SD)</td>
<td>54.0 (21.4)</td>
<td>54.4 (21.1)</td>
<td>53.3 (22.0)</td>
<td>55.5 (21.0)</td>
</tr>
<tr>
<td>Women, N (%)</td>
<td>625,879 (63.9%)</td>
<td>139,694 (63.5%)</td>
<td>110,322 (64.6%)</td>
<td>95,166 (67.0%)</td>
</tr>
<tr>
<td>Ethnicity</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>African/Caribbean</td>
<td>17,257 (1.8%)</td>
<td>11,732 (1.9%)</td>
<td>5,525 (1.7%)</td>
<td>751 (0.4%)</td>
</tr>
<tr>
<td>Asian</td>
<td>46,621 (5.0%)</td>
<td>29,458 (4.9%)</td>
<td>17,163 (5.2%)</td>
<td>6,695 (3.2%)</td>
</tr>
<tr>
<td>Mixed/Other</td>
<td>45,363 (4.9%)</td>
<td>27,201 (4.5%)</td>
<td>18,162 (5.5%)</td>
<td>10,202 (4.9%)</td>
</tr>
<tr>
<td>White</td>
<td>824,573 (88.3%)</td>
<td>533,414 (88.6%)</td>
<td>291,159 (87.7%)</td>
<td>189,790 (91.5%)</td>
</tr>
<tr>
<td>Missing</td>
<td>45,058 (4.6%)</td>
<td>24,074 (3.8%)</td>
<td>20,984 (5.9%)</td>
<td>10,375 (5.1%)</td>
</tr>
<tr>
<td>Socioeconomic status quintile</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1 (least deprived)</td>
<td>220,047 (22.5%)</td>
<td>139,694 (22.3%)</td>
<td>80,353 (22.8%)</td>
<td>32,134 (22.6%)</td>
</tr>
<tr>
<td>2</td>
<td>204,731 (20.9%)</td>
<td>129,945 (20.8%)</td>
<td>74,786 (21.2%)</td>
<td>30,146 (21.2%)</td>
</tr>
<tr>
<td>3</td>
<td>197,109 (20.1%)</td>
<td>125,845 (20.1%)</td>
<td>71,264 (20.2%)</td>
<td>28,870 (20.3%)</td>
</tr>
<tr>
<td>4</td>
<td>186,129 (19.0%)</td>
<td>120,073 (19.2%)</td>
<td>66,056 (18.7%)</td>
<td>26,957 (19.0%)</td>
</tr>
<tr>
<td>5 (most deprived)</td>
<td>170,856 (17.5%)</td>
<td>110,322 (17.6%)</td>
<td>60,534 (17.1%)</td>
<td>23,844 (16.8%)</td>
</tr>
<tr>
<td>Number of autoimmune disorders</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>833,157 (85.1%)</td>
<td>522,095 (83.4%)</td>
<td>311,062 (88.1%)</td>
<td>187,435 (85.2%)</td>
</tr>
<tr>
<td>2</td>
<td>124,742 (12.7%)</td>
<td>87,678 (14.0%)</td>
<td>37,064 (10.5%)</td>
<td>28,024 (12.7%)</td>
</tr>
<tr>
<td>3 or more</td>
<td>20,973 (2.1%)</td>
<td>16,106 (2.6%)</td>
<td>4,867 (1.4%)</td>
<td>4,588 (2.1%)</td>
</tr>
<tr>
<td>Body mass index (kg/m²)</td>
<td>27.8 (6.33)</td>
<td>27.8 (6.69)</td>
<td>27.7 (5.55)</td>
<td>27.1 (5.75)</td>
</tr>
<tr>
<td>Mean (SD)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Missing (%)</td>
<td>509,139 (52.0%)</td>
<td>316,940 (50.6%)</td>
<td>192,199 (54.4%)</td>
<td>120,581 (54.8%)</td>
</tr>
<tr>
<td>Smoking status</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Current smoker</td>
<td>126,240 (20.7%)</td>
<td>77,046 (19.7%)</td>
<td>49,194 (22.6%)</td>
<td>18,463 (14.1%)</td>
</tr>
<tr>
<td>Former smoker</td>
<td>167,829 (27.6%)</td>
<td>92,501 (23.6%)</td>
<td>75,328 (34.7%)</td>
<td>37,139 (28.4%)</td>
</tr>
<tr>
<td>Never smoker</td>
<td>314,512 (51.7%)</td>
<td>221,793 (56.7%)</td>
<td>92,719 (42.7%)</td>
<td>75,190 (57.5%)</td>
</tr>
<tr>
<td>Missing (%)</td>
<td>370,291 (37.8%)</td>
<td>234,539 (37.5%)</td>
<td>135,752 (38.5%)</td>
<td>89,255 (40.6%)</td>
</tr>
</tbody>
</table>

Socioeconomic status was defined as the Index of Multiple Deprivation (IMD) 2015 quintile, with 1 referring to the most affluent and 5 to the most deprived socioeconomic quintile. For variables with missing entries, category percentages refer to complete cases, and are presented alongside the percentage of missing values.
Incidence rates are presented as incidence rates per 100 000 person-years at risk and are age-sex-standardised to the 2013 European Standard Population. ‘Any autoimmune disorder’ refers to the primary incidence of the 19 autoimmune disorders investigated in this study (that is the number of patients first diagnosed with one autoimmune disease). ‘N’ refers to the number of patients newly diagnosed with autoimmune disease during the study period.
Incidence rates were calculated per 1-year age band and divided into a colour-gradient of 20-quantiles to reflect incidence density by age. The definition of type 1 diabetes used in this study refers to childhood-onset type 1 diabetes and hence to individuals aged 19 years or less at the time of diagnosis. IQR = interquartile range.
Figure 3: Incidence of autoimmune disorders by sex and over time.

**Any autoimmune disorder** refers to the primary incidence of the 19 autoimmune disorders investigated in this study. The definition of type 1 diabetes used in this study refers to childhood-onset type 1 diabetes. Denominator populations include all age groups. Yearly incidence estimates were smoothed using loess (locally estimated scatterplot smoothing) regression lines. IRR refers to the incidence rate ratio of women compared to men and CI to the confidence interval.

<table>
<thead>
<tr>
<th>Autoimmune Disorder</th>
<th>IRR (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Addison</td>
<td>1.28 [1.19, 1.38]</td>
</tr>
<tr>
<td>Ankylosing spondylitis</td>
<td>0.68 [0.65, 0.71]</td>
</tr>
<tr>
<td>Coeliac</td>
<td>1.95 [1.68, 2.01]</td>
</tr>
<tr>
<td>Graves’ disease</td>
<td>3.26 [3.18, 3.36]</td>
</tr>
<tr>
<td>Hashimoto’s thyroiditis</td>
<td>3.12 [3.10, 3.14]</td>
</tr>
<tr>
<td>Inflammatory bowel disease</td>
<td>1.06 [1.05, 1.10]</td>
</tr>
<tr>
<td>Multiple sclerosis</td>
<td>2.31 [2.21, 2.42]</td>
</tr>
<tr>
<td>Myasthenia gravis</td>
<td>0.67 [0.65, 0.77]</td>
</tr>
<tr>
<td>Pernicious anaemia</td>
<td>1.38 [1.36, 1.39]</td>
</tr>
<tr>
<td>Polymyalgia rheumatica</td>
<td>1.63 [1.56, 1.71]</td>
</tr>
<tr>
<td>Primary biliary cholangitis</td>
<td>2.87 [2.69, 3.08]</td>
</tr>
<tr>
<td>Psoriasis</td>
<td>1.11 [1.09, 1.13]</td>
</tr>
<tr>
<td>Rheumatoid arthritis</td>
<td>1.94 [1.88, 1.99]</td>
</tr>
<tr>
<td>Sjogren</td>
<td>5.93 [5.59, 6.29]</td>
</tr>
<tr>
<td>Systemic lupus erythematosus</td>
<td>3.63 [3.46, 3.80]</td>
</tr>
<tr>
<td>Systemic sclerosis</td>
<td>3.30 [3.06, 3.55]</td>
</tr>
<tr>
<td>Type 1 diabetes</td>
<td>0.88 [0.82, 0.95]</td>
</tr>
<tr>
<td>Vasculitis</td>
<td>1.29 [1.26, 1.32]</td>
</tr>
<tr>
<td>Vitiligo</td>
<td>1.15 [1.11, 1.19]</td>
</tr>
</tbody>
</table>
Figure 4: Incidence of autoimmune disorders by socioeconomic quintile for selected disorders.

Age-sex-standardised incidence rates by socioeconomic status quintile (Index of Multiple Deprivation 2015) for selected autoimmune disorders. ‘Any autoimmune disorder’ refers to the primary incidence of the 19 autoimmune disorders investigated in this study. Yearly incidence estimates were smoothed using loess (locally estimated scatterplot smoothing) regression lines. IRR refers to the incidence rate ratio of the most deprived socioeconomic quintile compared to the least deprived quintile and CI to the confidence interval. Individual plots for each of the diseases investigated are presented in the appendix.
Figure 5: Incidence rate ratios for development of comorbid autoimmune disease among index populations with autoimmune disease compared with the general population.

Incidence rate ratios were calculated for development of a second (comorbid) autoimmune disease among index populations with pre-existing autoimmune disease compared with the general population using negative binomial regression models adjusted for age and sex. Type 1 diabetes refers to childhood-onset type 1 diabetes, that is among people aged 20 years or less at the time of diagnosis.