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1 Title page

2

3 **Neuropsychiatric diagnoses after montelukast initiation in paediatric asthma patients**

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29 **Keywords:** Adverse events; Asthma; Cohort study; Mental disorders; Montelukast; Paediatric

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31 contributed to the important intellectual content of the manuscript and agreed on the decision to submit.

32

1 **Key messages**

2 **What is already known on this topic**

3 Montelukast has been associated with various adverse neuropsychiatric outcomes, but the evidence base is  
4 inconclusive in children and young persons (CYP) with asthma.

5 **What this study adds**

6 In this cohort study of over 107,000 CYP, montelukast was associated with a 32% higher incidence of  
7 neuropsychiatric outcomes compared to those not exposed to montelukast, a statistically significant difference.  
8 Number needed to harm was 58.

9 **How this study might affect research, practice or policy**

10 Decision to prescribe montelukast in CYP with asthma should be based on a careful consideration of potential risks  
11 and benefits.

12

1 **ABSTRACT**

2 **Background:** The evidence base on montelukast-associated adverse outcomes is inconclusive in children and young  
3 persons (CYP) with asthma. We aimed to investigate 1-year incidence of neuropsychiatric diagnoses after initiation of  
4 montelukast as an adjunct therapy to inhaled corticosteroids in CYP aged 3–17 years with asthma.

5 **Methods:** This propensity score matched cohort study was conducted using electronic health records between  
6 2015–2019 in the TriNetX Analytics Network patient repository in the USA. Neuropsychiatric diagnoses were  
7 identified using the International Statistical Classification of Diseases, Tenth Revision, Clinical Modification (ICD-10-  
8 CM) codes. We estimated risk ratios (RR), absolute risk increase (ARI), and number needed to harm (NNH) with 95%  
9 confidence intervals (95%CI).

10 **Findings:** The mean age [SD] at index prescription in the 107,384 CYP with asthma was 8.7 [4.0] years (93,461 [87%]  
11 mild to moderate asthma; 62,301 [58%] male; 53,485 [50%] white; 33,107 [31%] Black/African American).  
12 Montelukast was associated with excess incidence of any neuropsychiatric outcome (71 per 1000 persons with  
13 montelukast and 54 per 1000 persons with no montelukast; RR, 1.32 [95%CI, 1.25–1.39]; ARI per 100 persons, 1.71  
14 [95%CI, 1.44–1.98]; 1-year NNH, 58 patients [95%CI, 51–69]). The highest excess risk in the montelukast group was  
15 for sleep disorders (RR, 1.63 [1.50–1.77]; ARI per 100 persons 1.17 [1.00–1.33]; NNH, 85 patients [75–100]).  
16 Montelukast use was also associated with excess incidence of anxiety disorders (RR, 1.10 [1.03–1.17]) and mood  
17 disorders (RR, 1.13 [1.03–1.24]).

18 **Conclusions:** In CYP with asthma who were treated with inhaled corticosteroids, adjunct treatment with montelukast  
19 was associated with a higher incidence of neuropsychiatric outcomes compared to those who were not exposed to  
20 montelukast.

21

22

## 1 INTRODUCTION

2 Inhaled corticosteroids (ICS) are the first-line therapy in children and young people (CYP) with chronic asthma.[1]  
3 Oral leukotriene receptor antagonists (LTRAs) can be used as an adjunct therapy in individuals who do not achieve  
4 adequate asthma control with daily ICS.[1] Montelukast is the most common LTRA prescribed worldwide and it is  
5 indicated for the maintenance treatment of asthma and for preventing asthma attacks in children aged 12 months  
6 and older. Other indications of montelukast are prevention of exercise-induced bronchoconstriction and relief of  
7 symptoms of allergic rhinitis. Montelukast is considered to be safe and well-tolerated, with mainly mild and  
8 infrequent side effects reported, such as headache and gastrointestinal symptoms.[2] In 2018, there were 2.3 million  
9 CYP under the age of 17 years in the U.S. who were prescribed montelukast.[3]

10 Asthma has been consistently associated with an increased risk of severe neuropsychiatric outcomes in CYP,  
11 including anxiety,[4,5] depression,[6,7] and even suicide.[8,9] The association between ICS and the risk of severe  
12 neuropsychiatric outcomes in CYP is not well established. However, ICS have been associated with hypothalamic-  
13 pituitary-adrenal (HPA) axis dysregulation.[10,11] It has been suggested that LTRAs could also be associated with  
14 HPA axis dysregulation, and that this could be the mechanism for montelukast-induced adverse drug reactions  
15 (ADRs) particularly in CYP.[12] Various severe neuropsychiatric outcomes have been associated with montelukast,  
16 including hallucinations, depression, and suicides.[13-18] Due to these concerns, the U.S. Food and Drug  
17 Administration (FDA) recommends limiting the use of montelukast in the treatment of allergic rhinitis to only those  
18 patients who cannot be treated effectively with alternative therapies.[3]

19 A total of 17 potential neuropsychiatric ADRs are listed in the labelling of LTRAs available in the U.S. as required by  
20 the FDA, along with a boxed warning of serious neuropsychiatric events reported in patients taking LTRAs.[3] The  
21 evidence based on the safety of LTRAs is almost exclusively based on montelukast because other LTRAs are rarely  
22 prescribed. In CYP with asthma, the most common ADRs associated with montelukast are anxiety, sleep problems,  
23 and various behavioural symptoms, including nervousness and aggression.[17] However, the evidence base in CYP is  
24 still considered inconclusive.[18] Therefore, further research is needed to clarify the nature and extent of the  
25 associations especially in CYP. The purpose of this study was to address this evidence gap by conducting a large  
26 propensity score matched cohort study of electronic medical records of over 800,000 paediatric asthma patients. We  
27 have previously shown that treatment with montelukast was associated with incident neuropsychiatric outcomes in

1 patients with asthma and allergic rhinitis who were aged 15–64 years.[19] This new study adds to the evidence base  
2 by focusing on CYP aged 3–17 years with asthma, by including symptoms not fulfilling diagnostic criteria for mental  
3 health disorders, and by presenting the findings separately by sex and age group. We will use two complementary  
4 approaches to analyse the data to improve assessment of potential sources of bias occurring after baseline matching.  
5 In our primary analysis, which is in principle similar to per-protocol analysis used in RCTs, participant eligibility is  
6 retained throughout the study period. This approach attempts to control various factors potentially confounding the  
7 observed associations during the one-year follow-up, and is thus our main analysis of ADRs associated with  
8 montelukast exposure. In our secondary analysis, which is in principle similar to intention-to-treat analysis used in  
9 RCTs, any changes in eligibility after baseline are ignored. Consistent associations across these two approaches will  
10 support the validity of the findings. To our knowledge, this is one of the largest study to date to explore the  
11 association between montelukast and neuropsychiatric ADRs in CYP with asthma.

12

## 1 METHODS

2 We used the TriNetX (TriNetX LLC., Cambridge, MA 02140, USA, [www.trinetx.com](http://www.trinetx.com)) patient repository and analytics  
3 platform to identify patients with asthma. The TriNetX Analytics Network provided access to pseudonymized  
4 information on diagnoses, dispensed prescription medicines, and clinical procedures on over 25 million patients aged  
5 3–17 years from 105 health care organizations (HCOs) mainly in the USA. This study is a secondary analysis of existing  
6 data, it does not involve intervention or interaction with human subjects, and is de-identified in line with the de-  
7 identification standard defined in Section §164.514(a) of the Health Insurance Portability and Accountability Act  
8 (HIPAA, 1996) Privacy Rule. The HIPAA Privacy Rule establishes the conditions under which protected health  
9 information may be used for research purposes. Based on the Privacy Rule, this retrospective study was exempt from  
10 Institutional Review Board approval and from informed consent.

11 Data were accessed and analysed via the TriNetX platform by using the TriNetX built-in query builder and algorithms.  
12 All diagnoses were identified using the International Classification of Diseases, tenth revision, clinical modification  
13 (ICD10-CM) codes. Dispensed prescription medicines were identified using the Veterans Affairs (VA) National  
14 Formulary codes and the normalized nomenclature for clinical drugs (RxNorm) codes. Clinical procedures were  
15 identified using the Current Procedural Terminology (CPT) codes. During the years 2015–2019, there were 844,429  
16 patients aged 3–17 years with asthma. From this population, we defined two separate asthma groups: one of  
17 patients treated with ICS and montelukast (i.e., montelukast group), and another of patients treated with ICS but not  
18 with montelukast (i.e., control group). Each study year (2015–2019), we used the first dispensed montelukast  
19 prescription in the data as the index prescription for asthma patients exposed to montelukast; whereas, for the  
20 control group, we used the first dispensed prescription for ICS or inhaled bronchodilators, whichever came first, as  
21 the index prescription.

22 Patients were eligible at baseline if they were aged 3–17 years at index prescription, had an asthma-related visit and  
23 were prescribed ICS during the same year, and had no recorded history of LTRA prescriptions (i.e., had incident/new  
24 prescription for montelukast or did not have known previous exposure to LTRAs in the control group). All participants  
25 were required to have at least six months of recorded patient history in the participating HCOs, defined as any  
26 patient encounter, before the index prescription. Nearly all (95% in the montelukast group and 98% in the control

1 group) of the CYPs had a recorded patient history of up to two years before the index prescription, including the  
2 minimum period of six months (eFigure 1).

3 *Primary analysis.* The main results are based on analyses where participants were required to remain in their original  
4 treatment group (i.e., those in the control group were not allowed to have montelukast treatment during the one-  
5 year follow-up to avoid exposure misclassification/contamination bias), were required to have persistent asthma  
6 during follow-up to avoid confounding by asthma status (i.e., had an asthma-related visit or were prescribed asthma  
7 medication also in the following year after the index year), had recorded follow-up information for the full study  
8 period (i.e., were not lost to follow-up for the outcomes), did not have diagnosed obstructive sleep apnoea (OSA)  
9 and did not have tonsillectomy or adenoidectomy during follow-up. Those with tonsillectomy or adenoidectomy  
10 during follow-up and those with diagnosed OSA were excluded to reduce confounding particularly in sleep-related  
11 outcomes but also in relation to anxiety and depression. Furthermore, excluding those with OSA excluded patients  
12 for whom montelukast was possibly prescribed off-label to treat symptoms of OSA rather than to treat symptoms of  
13 asthma. There were 53,793 patients in the montelukast group who fulfilled the eligibility criteria and 313,490  
14 patients in the control group who fulfilled the eligibility criteria and were thus included in propensity score matching  
15 (eFigure 2). Over 99% (364903/367283) of these patients were in the USA. After matching, there were 53,692 CYPs  
16 included per group.

17 *Secondary analysis.* In the secondary analysis, we used only baseline information to define eligibility and ignored  
18 potential changes in eligibility after baseline. Results from these analyses were used to validate the findings from the  
19 main analyses by comparing consistency of observed associations. The baseline eligible groups were matched for the  
20 same set of covariates as those in the primary analysis, which resulted in 73,443 matched CYPs per group. Of the  
21 patients in the control group, 1886 CYPs (2.6%) were prescribed montelukast during the one-year follow-up and thus  
22 changed their exposure status during the outcome measurement period. For 15,606 (21,2%) CYPs in the control  
23 group and 10,973 (14.9%) CYPs in the montelukast group, persistent asthma status could not be determined during  
24 follow-up because they did not have an asthma-related visit or had prescriptions for asthma medication during the  
25 following year after the index prescription. These post-baseline changes in eligibility could potentially affect the  
26 observed associations. The mean follow-up time, defined by the last known fact in the EHR, was 11.4 months



1 (standard deviation, 2.4) in the montelukast group and 11.4 (standard deviation, 2.5) in the control group, meaning  
2 that lost to follow-up did not likely bias the associations.

3 Primary outcome measures were 12-month incident psychotic disorders (ICD10-CM: F20–F29), mood disorders (F30–  
4 F39), anxiety disorders (F40–F48), sleep disorders (F51, G47), and suicidality defined by codes for intentional self-  
5 harm, suicidal ideation, and events of undetermined intent (X71–X83, T14.91, R45.851, Y21–Y33). For potential  
6 mental health symptoms not fulfilling diagnostic criteria for any of the primary diagnoses, we created two symptom  
7 groups using the ICD10-CM codes for emotional state (R45) and categorised these into externalising symptoms  
8 (nervousness, restlessness and agitation, irritability and anger, hostility, violent behaviour, impulsiveness, and  
9 homicidal ideations) and internalising symptoms (unhappiness, demoralisation and apathy, low self-esteem, worries,  
10 excessive crying, anhedonia, and emotional lability). We used dispensed prescriptions for sedatives and hypnotics,  
11 antidepressants, antipsychotics, stimulants, and medicines commonly prescribed for sleep problems as proxies for  
12 clinically important neuropsychiatric symptoms receiving treatment (eTable 1).

13 Sensitivity analyses were conducted within the main categories of neuropsychiatric outcomes: within mood disorders  
14 (bipolar disorder, major depression, dysthymic disorder, other and unspecified mood disorders), within anxiety  
15 disorders (phobic anxiety, generalised anxiety, obsessive-compulsive disorder (OCD), other and unspecified anxiety),  
16 within sleep disorders (insomnia, hypersomnia, parasomnias, movement disorders, other and unspecified sleep  
17 disorders), and suicidality (intentional self-harm, suicidal ideation, and events of undetermined intent). Furthermore,  
18 models for the primary outcomes were stratified by sex and age at index prescription (aged 3–9 years and 10–17  
19 years at index prescription, for which the cut-off age was based on the mean age at index prescription) to observe  
20 potential differences between males and females and by age group. A post-hoc sensitivity analysis was conducted  
21 separately for those whose race was recorded as Black or African American because this group represented 31% of  
22 the study population.

### 23 Statistical methods

24 Propensity score matching was used to control for various confounders that could bias the comparisons between the  
25 montelukast and control groups. These covariates included demographics, severity of asthma, covariates associated  
26 with treatment allocation, and various neuropsychiatric and somatic comorbidities (Table 1). For the propensity

1 score matching, we used 1:1 nearest neighbour matching with a caliper of 0.1 standard deviations. Based on  
2 standard differences ( $\geq 0.1$ ) before matching, distributions of patient characteristics differed across the montelukast  
3 group and the control group, e.g., in terms of asthma severity, covariates associated with montelukast indication,  
4 history of neuropsychiatric diagnoses, and somatic comorbidities (Table 1). Of the patients in the montelukast group,  
5 over 99% (53692/53793) were successfully matched with a patient in the control group. After matching, the groups  
6 were balanced for all the covariates included, i.e., all standard differences were below 0.1. We calculated risk ratios  
7 (RR) with 95% confidence intervals (95%CI) for the associations between montelukast and neuropsychiatric  
8 outcomes. For RRs above one, we calculated absolute risk increase (ARI) with 95% CI and number needed to harm  
9 (NNH, calculated as  $1/\text{ARI}$ ) with 95%CI. For the RRs, statistical significance was defined as 95%CIs not crossing 1; and  
10 for the ARIs, 95%CIs not crossing 0. There were no missing information on age, 0.1% had missing information on sex,  
11 and 12% had missing information on race. Those with missing information on sex or race were included in the  
12 propensity score matching as 'unknown sex' or 'unknown race' to balance the cohorts for missing information (Table  
13 1). All sensitivity analyses were matched for the same set of covariates as the main models.

14

## 1 FINDINGS

2 Of the 107,384 patients who were included in the primary analysis, 62,301 (58%) were males; 53,485 (50%) were  
3 white; 33,107 (31%) were Black or African American; their mean age was 8.7 years (standard deviation (SD), 4.0) at  
4 the index prescription; 93,461 (87%) had mild to moderate asthma; 61,673 (57%) had vasomotor or allergic rhinitis;  
5 35,320 (33%) had dermatitis or eczema; and 68,458 (64%) had a history of oral glucocorticoid prescriptions (Table 1).  
6 The most common neuropsychiatric comorbidity was disruptive behaviour disorder (attention deficit hyperactivity  
7 disorder (ADHD), (8198 patients, 8%) or conduct disorder (CD) (2710, 3%)), followed by sleep disorders (6799, 6%)  
8 and generalised anxiety disorder (4718, 4%). Patients in the montelukast group had a median of two dispensed  
9 prescriptions (mean, 2.2; SD, 2.6) during the 1-year follow-up period (data not shown).

10 During the 1-year follow-up, 3240 (7%) patients in the montelukast group and 2462 (5%) patients in the control  
11 group experienced at least one of the neuropsychiatric outcomes (incidence rate, 71 per 1000 persons vs 54 per  
12 1000 persons, respectively), corresponding to about 30% excess risk (RR, 1.32 [95%CI, 1.25–1.39]; ARI per 100  
13 persons, 1.71 [95%CI, 1.44–1.98]; NNH, 58 patients [95%CI, 51–69]) (Table 2). The highest excess risk in the  
14 montelukast group was for sleep disorders (RR, 1.63 [1.50–1.77]; ARI per 100 patients, 1.17 [1.00–1.33]; NNH, 85  
15 patients [75–100]). Excess risks were also observed for other neuropsychiatric outcomes, including anxiety and  
16 mood disorders, and externalising and internalising symptoms, but with lower ARIs and higher NNHs. Similar findings  
17 of excess risks in the montelukast group were observed for incident psychotropic medicines across all drug groups  
18 (Table 3).

19 The sensitivity analyses by diagnostic subgroups showed that the excess risk in sleep disorders applied to  
20 subcategories of sleep disorder diagnoses, including hypersomnia and insomnia (eTable 2). Similarly, excess risks  
21 were seen across anxiety-related diagnoses, including generalised anxiety and OCD. A relatively large proportion of  
22 the diagnoses in sleep disorders, anxiety disorders, and mood disorders were in the diagnostic categories of ‘other’  
23 or ‘unspecified’ disorders. Incident suicidal ideation, rather than self-harm, was more common in the montelukast  
24 group than in the control group, but the association was not statistically significant (RR, 1.25 [0.98–1.61]). When the  
25 models for main neuropsychiatric outcomes were stratified by sex and age, increased risks were observed in the  
26 montelukast group in males and females (eTable 3), and in those aged 3–9 years and 10–17 years (eTable 4). In CYP

1 whose race was Black or African American, montelukast was associated with a comparable excess risk of any  
2 neuropsychiatric outcome (RR, 1.29 [1.17–1.42]) (eTable 5) than what was observed in the total population.

3 The associations observed for the main neuropsychiatric outcomes in the secondary analysis showed that the RRs for  
4 anxiety disorders, mood disorders, sleep disorders, and for any neuropsychiatric outcomes showed comparable  
5 associations as in the primary analysis (results presented in Table 2 vs. eTable 6). However, differences in ARIs and  
6 thus also in the NNH were observed between the two study designs. In the primary analysis, the NNH was 58 (95%CI,  
7 51–69) for any neuropsychiatric outcome, whereas in the secondary analysis it was 45 (95%CI, 41–52), respectively.

8 The largest difference was seen in the NNH for sleep problems (85 (95%CI, 75–100) in the primary analysis and 52  
9 (95%CI, 47–57) in the secondary analysis). In other words, the changes in eligibility after baseline did not change the  
10 proportionate risk difference across the groups for the main outcomes, but the primary and secondary analysis  
11 populations likely differed in their baseline risks, as suggested by the differences in ARIs. However, the consistency of  
12 associations in the RRs between the two designs support the validity of the observed increased risks associated with  
13 montelukast treatment.

## 1 DISCUSSION

2 We compared the incidence of neuropsychiatric diagnoses between those with and without montelukast as an  
3 adjunct therapy to ICS in CYP with asthma. Patients treated with montelukast had a 32% higher 1-year incidence of  
4 any neuropsychiatric outcomes compared to those treated with ICS alone. The highest excess incidence was for sleep  
5 disorders. NNH to cause one excess neuropsychiatric outcome was 58 in the montelukast group. The NNH was  
6 lowest for sleep disorders (NNH, 85), followed by anxiety disorders (NNH, 200), and mood disorders (NNH, 417). The  
7 observed associations were consistent across males and females, and across age groups.

8 These findings support the conclusion that montelukast is associated with neuropsychiatric outcomes in CYP with  
9 asthma, and that sleep problems and anxiety are common ADRs during treatment with montelukast.[17,20-23] In  
10 contrast with some previous studies,[17,18] we found an increased incidence of mood disorders in CYP with asthma  
11 who were treated with montelukast. Although the increased risk of mood disorders was smaller than for sleep  
12 disorders or anxiety, these disorders have a marked impact on the overall health and wellbeing of affected  
13 individuals; particularly if the symptoms are intensive and persistent, and occur during critical transitioning phases of  
14 childhood and adolescence. The excess incidence of antidepressant prescriptions in the montelukast group support  
15 the finding of increased symptoms of depression and anxiety. In our previous study, in a population of mainly adults,  
16 we did not find an increased risk of depression.[19] It is, therefore, possible that montelukast exacerbates a pre-  
17 existing HPA axis dysregulation particularly in CYP who have an initial suboptimal asthma symptom control, for which  
18 montelukast is commonly prescribed for.[12]

19 In sleep, anxiety, and mood disorders, diagnostic codes assigned for other and unspecific disorders were frequently  
20 used. This likely reflects diagnostic challenges on one hand and practices in assigning working diagnoses for  
21 neuropsychiatric symptoms in paediatric populations on the other hand. Increased risk of unspecific anxiety  
22 associated with montelukast has been reported previously.[20] These diagnostic challenges may partly explain the  
23 conflicting evidence from previous studies in CYP, together with other methodological limitations, such as small and  
24 selected study populations and limited control for confounders.[18] Potential montelukast associated  
25 neuropsychiatric ADRs are often identified and reported as symptoms rather than diagnoses.[17] Subjective  
26 symptom experience and presentation of symptoms in CYP may produce challenges in identifying neuropsychiatric  
27 symptoms and thus correctly diagnosing these symptoms. Certain ADRs, such as sleep disorders, are more likely to

1 be diagnosed, because a single symptom, such as difficulty falling asleep, can be used for a diagnosis of insomnia.  
2 However, as sleep problems are a known potential ADR of montelukast, clinicians may discontinue treatment with  
3 montelukast without a formal diagnosis of sleep disorders once these symptoms occur. There may also be  
4 considerable delay in arriving with a psychiatric diagnosis particularly in younger children, meaning that a longer  
5 follow-up than what is typically used in RCTs is needed to capture ADRs. Disruptive behaviours (ADHD and CD) were  
6 the most common baseline comorbidities in our data. Hyperactivity is commonly reported as a potential ADR  
7 associated with montelukast, but montelukast has not been found to be associated with an ADHD diagnosis.[24] This is  
8 not surprising as ADHD is a neurodevelopmental disorder. Hyperactivity or restlessness observed in CYP treated with  
9 montelukast is thus more likely associated with anxiety or sleep disorders. Our consistent findings on various  
10 neuropsychiatric outcomes, thus, strengthen the partly mixed evidence base from larger observational studies in  
11 paediatric populations.[21,25,26]

## 12 **Methodological considerations**

13 As with all studies based on register data, this study had limitations. We did not have information on the duration of  
14 montelukast treatment, adherence, or for the indication for which montelukast had been prescribed for (e.g.,  
15 whether primarily for asthma or allergic rhinitis). If patients received treatment from HCOs not participating in the  
16 TriNetX, before or after index prescription, this may have introduced bias due to misclassification in exposure status  
17 or loss to follow-up for the outcomes. However, because our eligibility criteria required a visit during the following  
18 year after the index prescription, lost to follow-up for the outcomes is unlikely. Because we used recorded diagnoses  
19 to identify neuropsychiatric outcomes, we only identified patients receiving treatment for these conditions. This has  
20 likely underestimated the number of patients experiencing these outcomes because patients with less severe  
21 symptoms may have not been identified. We focused on new treatment episodes with montelukast to reduce bias,  
22 because history of montelukast treatment is an indication of tolerability. Therefore, the associations may be  
23 different among patients with continuous long-term exposure to montelukast. It is possible that some residual  
24 confounding by indication remained in the models by unmeasured risk factors not captured by the data. Some case  
25 studies have reported that adverse montelukast-associated symptoms have resolved after cessation of treatment  
26 and returned after restarting treatment.[27] Our data did not enable us to establish a potential dechallenge-  
27 rechallenge association. We did not have information on outpatient deaths, meaning that we did not identify

1 completed suicides, if the individual was not treated in hospital before death. The HCOs participating in the TriNetX  
2 cannot be considered as representative of all HCOs in the USA., meaning that the true incidence of ADRs associated  
3 with montelukast remains unknown. Finally, it should be noted that the FDA issued safety warnings and other  
4 publicity around potential montelukast-associated adverse neuropsychiatric outcomes may have increased  
5 sensitivity of guardians and clinicians to monitor for and report neuropsychiatric symptoms in CYP treated with  
6 montelukast. This limitation applies to all observational studies on montelukast. A major strength of our data was the  
7 large number of CYP initially prescribed with montelukast, which allowed us to control for various potential baseline  
8 confounders and to establish the associations for various rare neuropsychiatric outcomes. Practically all patients in  
9 the montelukast group were successfully matched with a patient in the control group, which further reduced  
10 potential bias. Finally, consistent associations were observed across the two analysis approaches based on the  
11 principles of per-protocol and intention-to-treat analyses, which further supported the validity of our findings.

## 12 **CONCLUSIONS**

13 Although the observed overall risk increase for any adverse neuropsychiatric outcomes associated with montelukast  
14 was relatively small in absolute terms, these findings have important implications. First, because over 1/100  
15 individuals experienced adverse outcomes during treatment, adverse neuropsychiatric outcomes associated with  
16 montelukast can be considered to be common in CYP with asthma. Whether these excess risks warrant change in  
17 clinical practice at individual level requires consideration by clinicians, patients, and their families in light of the risk  
18 profile reported here. Thus, a careful risk-benefit assessment should be made before the decision on prescribing  
19 montelukast in CYP with persistent asthma.[28,29] Second, given the wide use of montelukast and the relatively low  
20 overall NNH of 58 in those treated with montelukast, suggests that there is a large number of montelukast-  
21 associated adverse neuropsychiatric outcomes at population level. The potential association with suicidal ideation  
22 warrants further research in larger populations exposed to montelukast. Because we observed only diagnosed  
23 conditions, i.e., those that were most severe or disruptive, the actual NNH is likely to be lower. Given the existing  
24 alternative add-on therapies, e.g., inhaled corticosteroids – long acting beta agonist combination therapy, these  
25 population level effects are important to consider when establishing treatment choices in consensus guidelines.

26

1 **STATEMENTS**

2 **Contributions**

3 TP, SF, and JF designed the study. SL provided methodological support. TP did the statistical analyses. TP wrote the  
4 first draft of the manuscript. All authors (TP, JF, CT, SL, KH, and SF) reviewed and revised the manuscript. SF is the  
5 guarantor. The corresponding author attests that all listed authors meet authorship criteria and that no others  
6 meeting the criteria have been omitted. TP and SL had full access to all the data in the study and take responsibility  
7 for the integrity of the data and the accuracy of the data analysis.

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9 TriNetX LLC (no grant number).

10 **Competing interests**

11 The authors have no conflict of interests to declare. SL and CT were employees of the TriNetX at the time when the  
12 study was conducted.

13 **Ethics approval**

14 Because this study used only de-identified patient records and did not involve the collection, use, or transmittal of  
15 individually identifiable data, this study was exempted from Institutional Review Board approval and from informed  
16 patient consent according to the U.S. federal law (Health Insurance Portability and Accountability Act 1996).

17 **Data sharing**

18 Data was provided by TriNetX ([www.trinetx.com](http://www.trinetx.com)), a federated data network. Access to TriNetX's de-identified  
19 patient data is available for the purpose of health care research with an approved user license.

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**Table 1.** Characteristics of montelukast and control groups before and after matching.

Characteristics. %	Before matching			After matching		
	Montelukast group (n=53793)	Control group (n=313490)	Std diff	Montelukast group (n=53692)	Control group (n=53692)	Std diff
Mean age at index prescription (SD)	8.6 (4.0)	9.2 (4.2)	0.126	8.6 (4.0)	8.7 (4.0)	0.009
Male	58.1	56.9	0.023	58.1	57.9	0.014
Unknown sex	0.1	0.04	0.018	0.1	0.1	0.010
White	49.2	46.3	0.059	49.2	50.5	<0.001
Black or African American	31.2	33.6	0.056	31.2	30.5	0.026
Other Race	7.6	9.0	0.016	7.6	7.3	0.015
Unknown race	12.1	11.1	0.030	12.1	11.7	0.005
Mild intermittent asthma	36.8	47.1	0.201	36.9	36.5	0.011
Mild persistent asthma	27.5	22.4	0.138	27.5	28.0	0.010
Moderate persistent asthma	23.1	11.6	0.323	23.0	22.1	0.012
Severe persistent asthma	3.7	1.0	0.184	3.6	3.0	0.023
Bronchodilators. inhalants	90.9	96.4	0.227	90.9	91.3	0.034
Glucocorticoids. oral	63.7	62.9	0.031	63.7	63.8	0.015
Wheezing	21.4	24.0	0.059	21.4	21.0	0.006
Acute upper respiratory infections	51.7	61.1	0.186	51.7	51.1	0.009
Vasomotor and allergic rhinitis	57.4	43.1	0.291	57.3	57.5	0.012
Chronic rhinitis. nasopharyngitis and pharyngitis	9.9	6.8	0.115	9.9	9.7	0.006
Snoring	10.5	8.9	0.053	10.5	10.1	0.009
Dermatitis and eczema	33.1	36.1	0.057	33.2	32.6	0.011
Suppurative and unspecified otitis media	26.1	31.5	0.116	26.1	25.5	0.020
Headache	7.4	9.5	0.075	7.4	7.2	0.014
Other headache syndromes	3.9	5.4	0.076	3.9	3.8	0.007
Overweight and obesity	8.2	10.6	0.080	8.2	7.9	0.007
Other functional intestinal disorders	14.1	16.4	0.060	14.1	13.6	0.014
Other and unspecified noninfective gastroenteritis and colitis	11.5	13.8	0.070	11.5	11.2	0.014
Gastro-oesophageal reflux disease	9.4	9.3	0.020	9.4	9.2	0.009
Nocturnal enuresis	1.5	2.1	0.041	1.5	1.5	0.007
Congenital malformations. deformations and chromosomal abnormalities	11.7	13.7	0.057	11.7	11.4	0.006
Depressive episode	1.9	3.1	0.069	1.9	1.8	0.008
Generalised anxiety disorder	4.5	6.1	0.065	4.5	4.3	0.009
Adjustment disorders	1.3	2.0	0.053	1.3	1.3	0.007
Sleep disorders not due to a substance or known physiological condition	0.8	0.9	0.016	0.8	0.8	0.007
Sleep disorders	5.7	5.6	0.010	5.7	5.4	0.003
Attention-deficit hyperactivity disorders	7.6	9.8	0.075	7.6	7.6	0.010
Conduct disorders	2.6	3.8	0.071	2.6	2.5	0.003
Injuries to the head	11.5	15.4	0.111	11.5	11.1	0.010
Suicidal ideations	0.4	0.8	0.041	0.4	0.4	0.011
Suicide attempt	0.1	0.1	0.011	0.1	0.1	0.010
Intentional self-harm	0.1	0.2	0.018	0.1	0.1	<0.001
Visit: Emergency	35.4	42.6	0.146	35.4	34.6	<0.001
Visit: Inpatient encounter	17.2	16.8	0.032	17.2	16.7	0.017

Std diff, standardised mean difference; SD, standard deviation of the mean. Asthma patients aged 3–17 years at index prescription in years 2015–2019. Covariates recorded at or before index prescription.

**Table 2.** One-year incident neuropsychiatric outcomes.

	Montelukast group		Control group		RR (95%CI)	ARI (95%CI)	NNH (95%CI)
	Patients in group	Patients with outcome	Patients in group	Patients with outcome			
Psychotic disorder	53394	50	53369	50	1.00 (0.68, 1.48)	n/a	n/a
Mood disorder	51987	898	52031	773	1.16 (1.05, 1.29)	0.0024 (0.0011, 0.0037)	417 (270, 909)
Anxiety disorder	50061	1842	50036	1591	1.16 (1.08, 1.24)	0.0050 (0.0031, 0.0069)	200 (145, 323)
Sleep disorder	50315	1519	50473	935	1.63 (1.50, 1.77)	0.0117 (0.0100, 0.0133)	85 (75, 100)
Suicidality <sup>a</sup>	52971	324	53147	258	1.25 (0.98, 1.61)	0.0013 (0.0005, 0.0020)	769 (500, 2000)
Externalising symptoms	52900	240	52987	192	1.25 (1.02, 1.53)	0.0009 (0.0003, 0.0016)	1111 (625, 3333)
Internalising symptoms	52919	207	52841	151	1.36 (1.09, 1.71)	0.0011 (0.0005, 0.0017)	909 (588, 2000)
Any neuropsychiatric outcome	45735	3240	45828	2462	1.32 (1.25, 1.39)	0.0171 (0.0144, 0.0198)	58 (51, 69)

n/a, not applicable; RR, rate ratio; CI, confidence interval; ARI, absolute risk increase; NNH, number needed to harm calculated as 1/ARI.

<sup>a</sup>Suicidality including suicidal ideation, intentional self-harm, and events of undetermined intent. Asthma patients aged 3–17 years at index prescription in years 2015–2019. The same individual could have multiple incident diagnoses from different diagnostic groups during follow-up.

**Table 3.** One-year incident prescriptions for psychotropics.

Psychotropics	Montelukast group		Control group		RR (95%CI)
	Patients in group	Patients with outcome	Patients in group	Patients with outcome	
Sedatives	49468	1467	48990	1170	1.24 (1.13, 1.35)
Antidepressants	51525	1084	51711	792	1.37 (1.25, 1.51)
Antipsychotics	52937	306	53040	228	1.34 (1.13, 1.59)
Stimulants	50362	1194	50298	876	1.36 (1.25, 1.48)
Sleep medications	52274	725	52383	517	1.40 (1.24, 1.58)
Any selected psychotropic	45316	2787	45095	2198	1.26 (1.19, 1.33)

RR, rate ratio; CI, confidence interval. Asthma patients aged 3–17 years at index prescription in years 2015–2019. The same individual could have multiple incident prescriptions from different drug classes during follow-up.