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Dark Energy Survey Year 3 Results: Cosmological constraints from second and third-order shear statistics

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We present a cosmological analysis of the third-order aperture mass statistic using Dark Energy Survey Year 3 (DES Y3) data. We perform a complete tomographic measurement of the three-point correlation function of the Y3 weak lensing shape catalog with the four fiducial source redshift bins. Building upon our companion methodology paper, we apply a pipeline that combines the two-point function ξ_{\pm} with the mass aperture skewness statistic $\langle \mathcal{M}_{\text{ap}}^3 \rangle$, which is an efficient compression of the full shear three-point function. We use a suite of simulated shear maps to obtain a joint covariance matrix. By jointly analyzing ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ measured from DES Y3 data with a Λ CDM model, we find $S_8 = 0.780 \pm 0.015$ and $\Omega_m = 0.266^{+0.039}_{-0.040}$, yielding 111% of figure-of-merit improvement in Ω_m - S_8 plane relative to ξ_{\pm} alone, consistent with expectations from simulated likelihood analyses. With a w CDM model, we find $S_8 = 0.749^{+0.027}_{-0.026}$ and $w_0 = -1.39 \pm 0.31$, which gives an improvement of 22% on the joint S_8 - w_0 constraint. Our results are consistent with $w_0 = -1$. Our new constraints are compared to CMB data from the Planck satellite, and we find that with the inclusion of $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ the existing tension between the data sets is at the level of 2.3σ . We show that the third-order statistic enables us to self-calibrate the mean photometric redshift uncertainty parameter of the highest redshift bin with little degradation in the figure of merit. Our results demonstrate the constraining power of higher-order lensing statistics and establish $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ as a practical observable for joint analyses in current and future surveys.

I. INTRODUCTION

The large-scale structure of the Universe has been extensively studied over the past decades by a variety of cosmological surveys. In this context, the analysis of cosmic shear has proven to be a powerful tool in placing competitive constraints on cosmological theory models [1–5]. The traditional way to perform such analyses is by measuring second-order summary statistics, such as the two-point correlation function (2PCF), on survey catalogs, and subsequently testing the model predictions for the statistic being used. However, second-order statistics are not able to capture the rich non-Gaussian information imprinted by the non-linear growth of cosmic structure.

The cosmic shear three-point correlation function (3PCF) was studied by Schneider and Lombardi [6] and proposed as a complement to the two-point function that allows the probing of non-Gaussian features of the field. Other higher-order statistics (HOS) have since been proposed [7–12] and applied to mock and survey data, obtaining different levels of improvement over their corresponding second-order constraints [13–18]. Most of the available HOS are constructed from simulated maps, not being analytically built from theoretical models.

High signal-to-noise (S/N) measurements of the 3PCF on Dark Energy Survey Year 3 (DES Y3) data have been performed by Secco *et al.* [1], showing the potential of this data set to undergo a joint 2PCF and 3PCF analysis. Such an analysis has the advantage of being directly modeled from theory, and also allows us a more direct study of the contributions of different scales. This stands in contrast to the approach of Gatti *et al.* [19], which requires mass map reconstruction from the Y3 shear catalog in order to extract the third moment of the convergence field. Our previous work has established a fast algorithm to compute the model prediction of the three-point correlation function [20], obtaining an improvement on runtime of six orders of magnitude relative to brute-force integration [see also 21, 22, for a similar idea for the fast modeling of three-point correlation function]. We make use of these advances and build a pipeline for a full cosmological analysis of the shear two-point functions ξ_{\pm} and of the $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ statistic, which is shown to be an efficient compression of the full 3PCF [23], on DES Y3 data.

In a companion paper, which we will refer to as Paper I [24], we have developed the methodology, built the joint covariance matrix, and performed modeling of systematics along with thorough simulated likelihood analyses.

We have found our methodology to be competitive with other higher-order analyses, yielding a factor-of-two improvement on the joint Ω_m - S_8 constraint. This level of improvement is similar to that found by Gatti *et al.* [25] through simulation-based inference combining scattering transforms, wavelet phase harmonics, and third moments of the convergence map. It is also comparable to the $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ analysis results from the KiDS collaboration [26]. In this paper, we apply our methodology to DES Y3 data and obtain improved constraints on Ω_m and S_8 , maintaining our predicted factor-of-two improvement on the joint constraint.

The structure of this paper is as follows. In Section II, we present our data and covariance matrix. In Section III, we review the theoretical modeling of the weak lensing summary statistics. In Section IV, we present our measurements of the three-point correlation function and of the mass aperture statistic in the Y3 shape catalog. In Section V, we briefly review our parameter inference methodology and present our blinding procedure. In Section VI, we present and discuss our cosmological constraints, with concluding remarks in Section VII.

II. DATA

A. DES Y3 shape catalog

For this study, we use the DES Y3 weak lensing shape catalog described by Gatti *et al.* [27] [28]. This catalog was produced from the first three years of data collected through the DECam, which operated on the Blanco 4m telescope, located on the Cerro Tololo Inter-American Observatory. The shear was measured in the data through the METACALIBRATION pipeline, resulting in a catalog with 100,204,026 galaxies. The effective area of the sky covered through these observations is 4143 deg^2 , and the effective number density is $5.59 \text{ gal/arcmin}^2$.

The catalog is split into four tomographic redshift bins, with each bin being characterized by a redshift distribution function $n(z)$. The distribution functions for all redshift bins are presented by Secco *et al.* [29]. These functions are obtained through the self-organizing map $p(z)$ (SOMPZ) method, and combined with estimates from clustering redshifts and shear ratios, as described by Myles *et al.* [30]. Each of the four bins has a similar number density of galaxies.

B. CosmoGrid simulations for covariance measurement

We use Y3 weak lensing data to obtain cosmological constraints from the shear two-point and three-point correlation functions. As described in Section III, we perform an analysis with a joint 2PCF and mass aperture data vector.

In order to obtain adequate posteriors for our parameter inference process, we need to compute a joint covariance of our full data vector, which concatenates the two-point and three-point information. This covariance can be estimated either analytically, or from a large set of simulations, or through jackknife sampling of the real data. For this analysis, we use the same covariance matrix described in Paper I, which includes the second-order shear covariance (ξ_{\pm}), the third-order mass aperture covariance ($\langle \mathcal{M}_{\text{ap}}^3 \rangle$), and their cross-covariance. The full matrix is estimated using 796 mock realizations generated from the **CosmoGridV1** fiducial cosmology simulations [31]. Each realization consists of a DES-Y3-like cosmic shear map, cutting a DES-Y3 footprint from a full-sky simulated shear field and adding shape noise derived from the DES-Y3 shape catalog [27].

Measurements of ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ are performed using **TreeCorr**, with the angular binning and filters described in Paper I. The sample covariance is then computed as:

$$\mathbf{C} = \frac{1}{N_{\text{real}} - 1} \sum_{r=1}^{N_{\text{real}}} [\mathbf{d}^r - \bar{\mathbf{d}}] [\mathbf{d}^r - \bar{\mathbf{d}}]^T, \quad (1)$$

where \mathbf{d}^r is the joint data vector of the r -th realization, $\bar{\mathbf{d}}$ is the mean data vector, and $N_{\text{real}} = 796$.

To ensure the robustness of the covariance model, we apply scale cuts to the three-point correlation function prior to compressing it into $\langle \mathcal{M}_{\text{ap}}^3 \rangle$. We remove contributions from triangles with side lengths below $\theta \approx 8'$. This addresses inaccuracies caused by the resolution of the simulated maps. The impact of this cut and other validation checks, including comparisons with analytic models for ξ_{\pm} covariance [2], along with tests of baryonic feedback, are detailed in Paper I. We verify the convergence of our covariance matrix by performing a simulated analysis from a subsample of our total simulations (700 out of 796) and comparing the final constraints with our fiducial simulated analysis. The number of simulations is also sufficiently larger than our data vector length, which, after compression, equals 96.

We confirm that the cross-covariance between ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ is negligible, with no significant structure in the matrix. The correlation coefficient has a maximum value of 0.13, consistent with the low overlap in information content between the two statistics.

III. WEAK LENSING THEORY

A. Convergence and Shear

Weak gravitational lensing is a powerful probe of the universe's matter distribution, sensitive to both visible and dark matter. The convergence field, $\kappa(\boldsymbol{\theta})$ quantifies isotropic magnification of background galaxy shapes and is given by the line-of-sight integration of the matter

density contrast weighted by the lensing efficiency [32]:

$$\kappa(\boldsymbol{\theta}) = \frac{3\Omega_m H_0^2}{2c^2} \int_0^\infty d\chi q_i(\chi) \frac{\delta_m(\chi\boldsymbol{\theta}, \chi; z(\chi))}{a(\chi)}, \quad (2)$$

where Ω_m and H_0 is the matter density and Hubble parameter at redshift zero, c is the speed of light, χ is the comoving distance, $a(\chi)$ is the scale factor, and δ_m is the matter density contrast. The lensing efficiency, $q_i(\chi)$, depends on the source galaxy distribution, $p(\chi)$:

$$q_i(\chi) = \int_\chi^\infty d\chi' p_i(\chi') \frac{\chi' - \chi}{\chi'}, \quad \int d\chi p(\chi) = 1. \quad (3)$$

The weak lensing shear field in Cartesian coordinates is defined as $\gamma_c(\boldsymbol{\theta}) = \gamma_1(\boldsymbol{\theta}) + i\gamma_2(\boldsymbol{\theta})$, where γ_1 and γ_2 are shear components aligned with the x-axis and rotated by 45° , respectively. When projected onto a frame rotated by an angle ζ , the spin-2 transformation gives:

$$\gamma(\boldsymbol{\theta}; \zeta) = \gamma_t(\boldsymbol{\theta}; \zeta) + i\gamma_\times(\boldsymbol{\theta}; \zeta) = -\gamma_c(\boldsymbol{\theta})e^{-2i\zeta}, \quad (4)$$

where γ_t and γ_\times are the tangential and cross components of the shear.

The Fourier transform of the shear field is defined as:

$$\gamma_c(\boldsymbol{\theta}) = \int \frac{d^2\boldsymbol{\ell}}{(2\pi)^2} \gamma_c(\boldsymbol{\ell}) e^{-i\boldsymbol{\ell}\cdot\boldsymbol{\theta}}, \quad (5)$$

with a similar relation for the convergence field. In Fourier space, the shear and convergence fields are related by:

$$\gamma_c(\boldsymbol{\ell}) = \kappa(\boldsymbol{\ell}) e^{2i\beta}, \quad (6)$$

where β is the polar angle of the Fourier mode $\boldsymbol{\ell}$.

Under the Limber approximation, we write the convergence power and bi-spectra as

$$P_\kappa(\ell) = \frac{9\Omega_m^2 H_0^4}{4c^4} \int_0^\infty d\chi \frac{q_i(\chi) q_j(\chi)}{a^2(\chi)} P_\delta\left(\frac{\ell}{\chi}, z(\chi)\right), \quad (7)$$

$$B_\kappa(\ell_1, \ell_2, \ell_3) = \frac{27\Omega_m^3 H_0^6}{8c^6} \int_0^\infty d\chi \frac{q_i(\chi) q_j(\chi) q_k(\chi)}{a(\chi)^3 \chi} \times B_\delta\left(\frac{\ell_1}{\chi}, \frac{\ell_2}{\chi}, \frac{\ell_3}{\chi}, z(\chi)\right). \quad (8)$$

B. Correlation functions

The cosmic shear two-point and three-point correlation functions probe the matter distribution in real space. In order to theoretically model these observables, we start with the matter power spectrum and bispectrum. We model the first with the revised Halofit formula [33], and the second with the BiHalofit formula [34]. The convergence power and bi-spectrum are computed, next, from Eqs. 7 and 8.

By decomposing the convergence power spectrum into E and B modes, we can compute the shear two-point functions ξ_+ and ξ_- by

$$\xi_+(\boldsymbol{\theta}) = \int_0^\infty \frac{\ell d\ell}{2\pi} J_0(\ell\theta) [P_\kappa^E(\ell) + P_\kappa^B(\ell)], \quad (9)$$

$$\xi_-(\boldsymbol{\theta}) = \int_0^\infty \frac{\ell d\ell}{2\pi} J_4(\ell\theta) [P_\kappa^E(\ell) - P_\kappa^B(\ell)], \quad (10)$$

For the modeling of the shear three-point functions, we make use of the multipole formalism, as implemented in our `fastnc` code [20]. We use the natural components of the three-point function, as defined by Schneider and Lombardi [6], and write them in the \times -projection introduced by Porth *et al.* [35]. We obtain

$$\Gamma_0^\times(\theta_1, \theta_2, \phi) = - \int \frac{d^2\boldsymbol{\ell}_1}{(2\pi)^2} \frac{d^2\boldsymbol{\ell}_2}{(2\pi)^2} e^{-i\boldsymbol{\ell}_1\cdot\boldsymbol{\theta}_1 - i\boldsymbol{\ell}_2\cdot\boldsymbol{\theta}_2} \times B_\kappa(\ell_1, \ell_2, \alpha) e^{2i\sum_i \beta_i} e^{-3i(\varphi_1 + \varphi_2)}, \quad (11)$$

with similar expressions for Γ_1 , Γ_2 and Γ_3 .

The bispectrum, now written in terms of two ℓ modes and the angle α between them, can be expressed as a sum over Legendre polynomials. Following Sugiyama *et al.* [20], we compute a sequence of multipoles of the natural components of the 3PCF for each combination of two triangle sides θ_1 and θ_2 . The full three-point function for θ_1 , θ_2 , and the opening angle ϕ is calculated next as a sum over all the computed multipoles.

C. Mass aperture statistic

In this work, we extract higher-order lensing information using the mass aperture statistic, a localized probe of the projected matter distribution. The skewness of the mass aperture, $\langle \mathcal{M}_{\text{ap}}^3 \rangle$, acts as a compressed summary of the non-Gaussian signal contained in the shear three-point function.

The utility of $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ as a compressed statistic has been demonstrated in recent work. As shown by Heydenreich *et al.* [23], this quantity captures nearly the full constraining power of the shear 3PCF, matching the performance of its leading principal components. This makes $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ a practical and powerful summary for cosmological inference.

We obtain the mass aperture statistic from the natural components of the 3PCF following the procedure detailed by Jarvis *et al.* [36]. One advantage of using this procedure instead of computing it from a convolution of the bispectrum is that we maintain a consistent 3PCF binning effect between our theoretical model and our measurements. Finally, we run our full theoretical pipeline on a set of cosmologies taken from a Sobol sequence in parameter space [37]. We sample over Ω_m , S_8 , h_0 , Ω_b , and n_s in Λ CDM.

We use our results to train a neural network emulator for use in our cosmological inference pipeline. Our emulator computes the z -dependent mass aperture statistic $\langle \mathcal{M}_{\text{ap}}^3 \rangle(\theta, z)$. In order to extract our theory vector from it, we integrate the emulator predictions over the line-of-sight using the kernel $q(\chi)$, which includes both the lensing kernel and the intrinsic alignment kernel (See Section III D). This integration is given by

$$\langle \mathcal{M}_{\text{ap}}^3 \rangle(\theta)_{ijk} = \int \frac{d\chi}{\chi} \frac{q_i(\chi)q_j(\chi)q_k(\chi)}{a(\chi)^3} \langle \mathcal{M}_{\text{ap}}^3 \rangle(\theta, z(\chi)) \quad (12)$$

We use six hidden layers and a sequence of decreasing learning rates. We use a testing set of 171 cosmologies, each with 144 outputs, and find that the network error is below 0.29% for 99% of the samples, and below 1.02% for 100% of the samples. We also build another emulator for w CDM, for which we construct a new training set sampling over w_0 as an additional parameter. We find the error across the 262 testing set cosmologies to be less than 0.7% for 99% of the samples. The emulator building process and the final network hyperparameters, along with the ranges of the sampled parameters, are described in detail in our Paper I.

D. Modeling of systematics

Here, we briefly review the systematic parameters used in our analysis, referring the reader to our methodology work for the complete description. For redshift calibration and multiplicative shear biases, we follow the procedure of the fiducial two-point DES-Y3 analysis [1, 2, 30, 38].

We account for the possibility of mean shifts in the redshift distributions by introducing four photo- z shift parameters Δz_i , one for each redshift bin i :

$$p_i(z) \rightarrow p_i(z - \Delta z_i). \quad (13)$$

Residual multiplicative biases on the shear are dealt with by introducing four parameters m_i , one for each redshift bin i . They are included in the correlation functions by

$$\xi_{\pm}^{ij} \rightarrow (1 + m_i)(1 + m_j)\xi_{\pm}^{ij}, \quad (14)$$

$$\Gamma_{\mu}^{ijk} \rightarrow (1 + m_i)(1 + m_j)(1 + m_k)\Gamma_{\mu}^{ijk}. \quad (15)$$

We model intrinsic alignment through the non-linear alignment (NLA) paradigm, which yields us the free parameters A_1 (amplitude parameter) and α_1 (for redshift evolution). While more complex models have been studied in the context of two-point functions, we aim for a consistent modeling between second- and third-order correlations, as suggested by Pyne *et al.* [39], and so we adopt NLA both for 2PCF and 3PCF. Our NLA modeling is implemented by modulating the lensing kernel through [40, 41]

$$q_i(\chi) \rightarrow q_i(\chi) + f_{\text{IA}}(z(\chi))p_i(\chi)\frac{dz}{d\chi}, \quad (16)$$

where $f_{\text{IA}}(z)$ is given in terms of the intrinsic alignment (IA) free parameters by

$$f_{\text{IA}}(z) = -A_1 \left(\frac{1+z}{1+z_0} \right)^{\alpha_1} \frac{c_1 \rho_{\text{crit}} \Omega_{\text{m},0}}{D(z)}. \quad (17)$$

Here, we use $z_0 = 0.62$ as the pivot redshift (this is a matter of convention, we follow the choice of Secco *et al.* [29]). Our growth function $D(z)$ is normalized to unity at redshift zero, and $c_1 \rho_{\text{crit}} = 0.0134$ is a constant determined through SuperCOSMOS observations [42].

The question of whether such NLA modeling is sufficient for a Y3 analysis has been addressed by Secco *et al.* [29]. While the fiducial Y3 cosmic shear analysis makes use of the tidal alignment and tidal torquing (TATT) model [43], the performance of both models is compared, and the shift in constraints is found not to be significant. The IA signal is found to be low, and an analysis of statistical model selection finds a preference for models with fewer parameters. TATT is a more realistic model, but it also adds complexity in parameter space, which can lead to the degradation of constraints. While improving our IA model will be necessary for larger datasets, we leave the complete development of a TATT model of the 3PCF and the mass aperture statistic for future work. Advancements in this direction are being made, with a theoretical modeling of the IA bispectra established by Bakx *et al.* [44] and validated against simulations of dark matter halo shapes [45].

IV. MEASUREMENTS

A. Three-point correlation function and mass aperture statistic

The three-point correlation function has been previously measured on the DES-Y3 cosmic shear catalog by [1]. Their measurements were performed in two different sets: (i) one non-tomographic measurement over the whole catalog, and (ii) tomographic measurements of auto and cross-correlations between two newly defined redshift bins, created by merging Y3 source bins 1 and 2, and also merging bins 3 and 4.

We perform a new set of measurements of the cosmic shear 3PCF on DES Y3 data in order to complete a full tomographic measurement and prepare our $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ data vector for parameter inference. We use the METACALIBRATION weak lensing shape catalog described by [27]. The relation between the measured ellipticity (e_1, e_2) and the cosmic shear (γ_1, γ_2) is described by the response matrix \mathbf{R}_{γ} . The METACALIBRATION algorithm artificially applies an additional shear to the original images, creating four additional versions of the ellipticity catalog: one positively sheared and one negatively sheared version for each e_i component. Using finite differences between

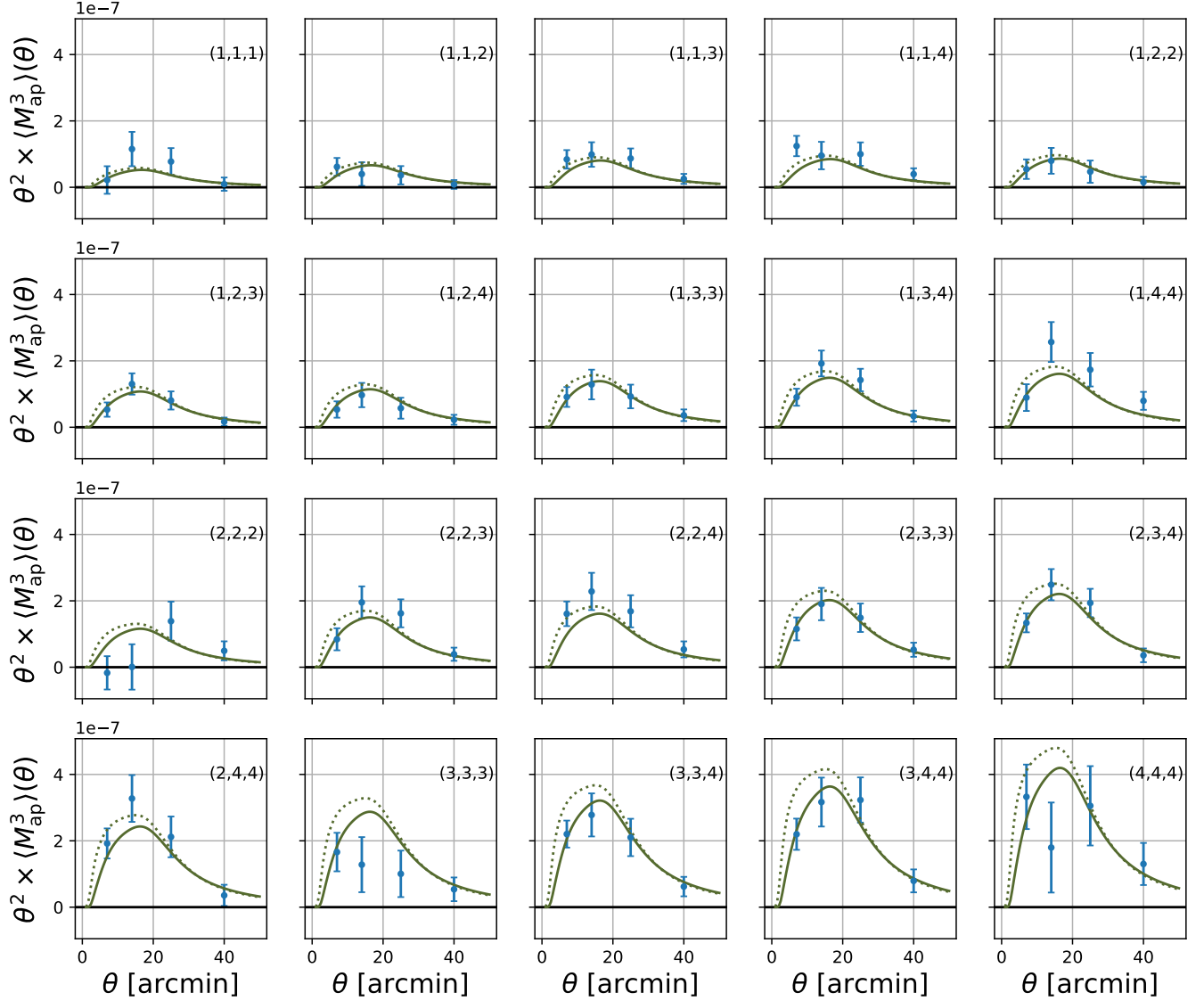


FIG. 1. The third moment of the mass aperture statistic Y3 data vector. The panels show the aperture mass statistic as a function of filter radii θ for different redshift-bin combinations (i, j, k) indicated on the upper right corner of each panel. The error bars are estimated from our simulation-based covariance. The theoretical prediction at our joint best fit cosmology, as described in Section VI, is shown by the continuous green line, which indicates the prediction with our 3PCF scale cut at 8 arcmin. The green dotted line shows the theoretical prediction with no scale cuts on the 3PCF. This is not expected to match the data (which does have scale cuts applied) but it shows that only a modest amount of the signal is lost by the scale cuts (difference between continuous and dotted lines).

these catalogs, the response matrix is estimated by

$$\mathbf{R}_{\gamma,ij} = \frac{e_i^+ - e_i^-}{\Delta\gamma_j}. \quad (18)$$

To the shear response matrix, we add the selection response \mathbf{R}_S matrix, as defined by [46], to account for selection effects. We use as our final response matrix $\mathbf{R} = \mathbf{R}_\gamma + \mathbf{R}_S$.

Following [1], we compute our three-point statistic using the unsheared ellipticity map from METACALI-

BRATION, masking it separately for each of the four DES Y3 source redshift bins. We subtract from the ellipticity component e_i of each galaxy its mean over the whole masked catalog, and divide it by the scalar $R = (\mathbf{R}_{11} + \mathbf{R}_{22})/2$, taking it as representative of the response matrix \mathbf{R} .

We perform the measurements of the three-point correlation function with **TreeCorr**[36], using the Multipole binning scheme. In this scheme, each triangle configuration is described by two side lengths and a set of mul-

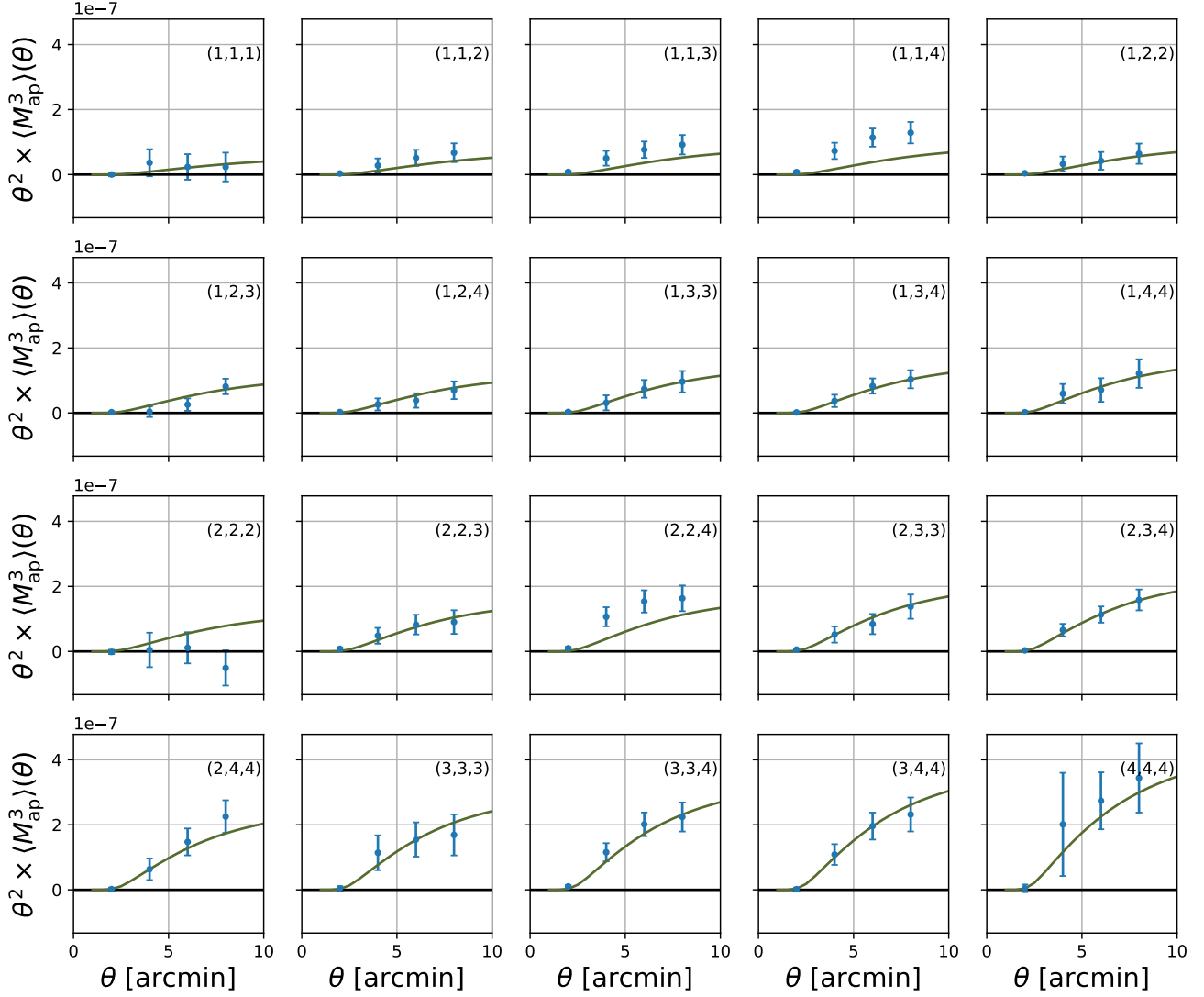


FIG. 2. Mass aperture statistic Y3 data, as in Figure 1, but with aperture filters at small scales. The green continuous line shows the theoretical prediction at best fit cosmology with our 3PCF scale cut at 8 arcmin. The dark-matter-only theoretical prediction is an adequate fit to the small scale data, with no systematic departure present across most of the redshift bin combinations.

tipoles of the opening angle. We take as the maximum multipole $\max_n = 100$. For the side binning, we take 20 bins logarithmically spaced between $\theta_{\min} = 0.5'$ and $\theta_{\max} = 80'$. The computational runtime of the measurement for each redshift bin combination is of the order of 3 hours with 128 CPUs at the Perlmutter system at the National Energy Research Scientific Computing Center (NERSC). The Multipole binning scheme, following the proposal by Porth *et al.* [35], allows for a measurement that scales linearly with the number of galaxies.

Next, we convert these measurements into the side-angle-side binning scheme, using 63 bins for the opening angle, which go from 0 to π with a linear spacing of $\phi = 0.05$.

Then, we perform a scale cut at the level of the 3PCF,

removing all triangle contributions with one or two sides smaller than $\theta = 8'$. This cut is motivated by considering the resolution of the simulations from which we compute our covariance matrix, and is designed so that we have a robust covariance across all scales.

Finally, these measurements are transformed into $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ data. The expression of $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ as an integral of the natural components Γ_i was developed by Jarvis *et al.* [36], and is implemented in **TreeCorr** through matrix multiplication of the binned 3PCF measurements.

Through this procedure, the mass aperture statistic is computed from the cosmic shear catalog, without the need of any convergence map reconstruction. There is another way to measure $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ from the shape catalog through reconstructing the convergence map, which,

however, suffers from masking effects on the convergence map and mass aperture field. Therefore, our approach of measuring $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ directly from the shape catalog, computing it as an integral of the 3PCF, naturally circumvents biases introduced by masked regions or irregular edges of survey footprint [1], and enables robust comparison of the theoretical prediction of $\langle \mathcal{M}_{\text{ap}}^3 \rangle$.

We use four equal-aperture filters, at $7'$, $14'$, $25'$, and $40'$. We introduce our smallest filter at $7'$ in order to avoid potential baryonic contamination. Our largest filter is chosen to be at $40'$, because we do not expect much gain in signal-to-noise ratio, and to avoid observational systematics at the large scales.

Our set of measurements is presented in Figure 1. The total $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ signal-to-noise, computed using the simulation-based covariance, is 12.7. The largest contributions towards the total S/N ratio are the bin combinations 234, 334, and 344, each with individual S/N ratios higher than 6. The lower redshift bins, however, still carry contributions, with the total $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ S/N being reduced to 10.5 if we remove all data that includes the first redshift bin.

We also measure the $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ with four additional filters at $2'$, $4'$, $6'$ and $8'$. We do not include these data points in our fits, but instead use them to further validate our methodology, by comparing them to the dark-matter-only theoretical prediction at our best-fit cosmology presented in Section VI. We show these measurements in Figure 2.

Finally, we also perform a non-tomographic measurement of the three-point correlation function over the whole catalog at the same scales of $\theta_{\text{min}} = 0.5'$ and $\theta_{\text{max}} = 80'$. We use this measurement to study the configuration dependence of the 3PCF. We select near-isosceles triangles, defined as those for which the d_2 angular scale bin equals that of d_3 . We then perform a weighted average over d_2 and d_3 , splitting our measurement between small scales ($d_2 \approx d_3 < 18'$) and larger scales ($18' < d_2 \approx d_3 < 80'$). We show our results in Figure 3. We take our error bars from jackknife sampling of the Y3 data, which we divide into 100 patches. We select this method instead of taking our covariance from simulations in order to avoid small scale biasing from simulation resolution. For the small scale selection, our S/N is 6.9, while for the larger scales, we find S/N = 11.1. We show the signal together with a reference theoretical expectation range, which we compute by averaging the non-tomographic signal of a set of CosmoGridV1 realizations. The amplitude of the signal is rescaled by the average ratio between the expected signals at CosmoGridV1 cosmology and at our $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ best fit cosmology. The upper and lower limits of our range are computed from cosmologies which are 1σ away from our best fit.

When we compare our measurements to those from Secco *et al.* [1], we notice that both our Y3 and CosmoGridV1 measurements remain shifted towards negative values as the opening angle approaches 180deg . In the case of the larger scales, we can see clearly that there is no

minimum point near such angles. The feature identified by Secco *et al.* [1] of a rise in the signal near $\phi = 180\text{deg}$, which doesn't match theoretical expectations, can be explained as the result of uneven binning when using the `ruv` binning option in `TreeCorr`. We remove this effect by performing our analysis with the `multipole` binning scheme, which adapts the multipole-based estimator developed by [35] into the tree structure.

The shape dependence of the 3PCF carries interesting features that can potentially serve as null tests for systematics. One such example is the zero crossing at an opening angle around $\phi \sim 120\text{deg}$. The expected sign of different components of the shear three-point function was studied by Zaldarriaga and Scoccimarro [47]. For isosceles triangles close to an equilateral configuration, we expect a tangential shear on the three vertices of the triangle. When the isosceles triangle approaches a flattened configuration, the presence of mass in the region inside the triangle induces almost always a radial shear on the obtuse angle vertex, and a tangential shear on the small angle vertices. This leads to a negative tangential γ_{ttt} .

Between these two limiting cases, equilateral and flattened isosceles triangles, the isosceles cosmic shear signal should always have a zero crossing. The angle of the zero-crossing is determined by the geometrical configuration of the mass and the triangle on the sky as described above, while the observational systematics, PSF systematics, does not follow this picture and can produce a shift in the position of zero-crossing. Therefore, the location of the zero-crossing could be used as a null test of the systematics. We leave this potential use of the zero-crossing as a null test for future research.

V. METHODOLOGY

A. Parameter inference pipeline

To infer cosmological and nuisance parameters from the DES Y3 data, we adopt a Bayesian framework in which the posterior distribution of the parameters \mathbf{p} given a data vector \mathbf{d} is proportional to the product of the likelihood and the prior distribution:

$$\mathcal{P}(\mathbf{p}|\mathbf{d}) \propto \mathcal{L}(\mathbf{d}|\mathbf{p})\Pi(\mathbf{p}). \quad (19)$$

We adopt the likelihood derived in Percival *et al.* [48], which naturally yields posterior credible intervals compatible with those from the frequentist approach. The log-likelihood is expressed in terms of the χ^2 statistic that measures the discrepancy between the observed data and the theoretical model prediction $\mathbf{t}(\mathbf{p})$. Thus, we have

$$\ln \mathcal{L}(\mathbf{d}|\mathbf{p}) = -\frac{m}{2} \ln \left(1 + \frac{\chi^2}{N_{\text{real}} - 1} \right) + \text{const}, \quad (20)$$

$$\chi^2 = [\mathbf{d} - \mathbf{t}(\mathbf{p})]^T \mathbf{C}^{-1} [\mathbf{d} - \mathbf{t}(\mathbf{p})], \quad (21)$$

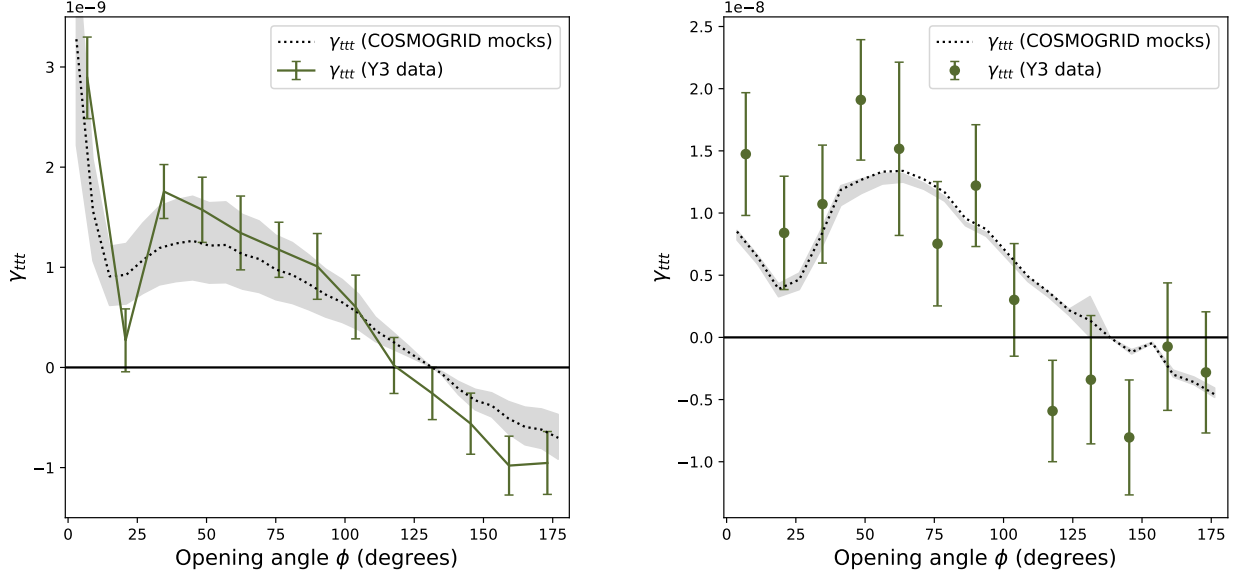


FIG. 3. Measurement of the isosceles three-point correlation function on DES Y3 data. We performed the measurement without redshift tomography, and present the results for scales below 18 arcmin (right panel) and between 18 and 80 arcmin (left panel). Both panels show the characteristic shape and zero crossing of the shear 3PCF. The error bars are estimated from the jackknife covariance. To provide a comparison with theoretical expectations, the dotted curve shows our simulation measurements, rescaled to account for the difference between cosmological parameters. The 1σ intervals around our $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ best fit cosmology is shown by the gray shaded region.

where $N_{\text{real}} = 796$ is the number of COSMOGRIDV1 realizations used to estimate the covariance matrix. We also define $N_{\mathbf{p}}$ as the number of free parameters and $N_{\mathbf{d}}$ as the length of the data vector. Then, we can write the m factor as

$$m = N_{\mathbf{p}} + 2 + \frac{N_{\text{real}} - 1 + f_D}{1 + f_D}, \quad (22)$$

where f_D is the Dodelson-Schneider factor [49], given by

$$f_D = \frac{(N_{\mathbf{d}} - N_{\mathbf{p}})(N_{\text{real}} - N_{\mathbf{d}} - 2)}{(N_{\text{real}} - N_{\mathbf{d}} - 4)(N_{\text{real}} - N_{\mathbf{d}} - 1)}. \quad (23)$$

Our fiducial data vector combines both two-point and three-point statistics: it includes the ξ_+ and ξ_- shear correlation functions as well as the third-order mass aperture statistic, $\langle \mathcal{M}_{\text{ap}}^3 \rangle$. For ξ_{\pm} , we use 20 angular bins logarithmically spaced between $2.5'$ and $250'$ for each redshift bin pair. We apply the fiducial Y3 scale cuts, described by Krause *et al.* [50], to mitigate the effects of baryonic physics, retaining 227 data points in total. We apply MOPED compression [51] on the ξ_{\pm} section of our data vector, following the procedure described in Paper I. We use the analytical ξ_{\pm} covariance matrix to determine the MOPED compression matrix, and then use the simulated covariance for cosmological inference. We found that this mitigates the risk for overconfidence that would be present if the compression matrix had been derived from a simulated covariance.

For the three-point part, we use all 80 mass aperture data points, spanning 20 redshift combinations and four aperture radii from $7'$ to $40'$. This choice of scales is described and shown to be robust against baryonic feedback in Paper I. In our simulated analysis, we find that through this set of filters the small scale baryonic suppression of the bispectrum is canceled out by the mid scale enhancement, generating a small net enhancement effect. A joint analysis of ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$, therefore, becomes even more robust to the effect of baryons.

We consider both Λ CDM and w CDM models in our inference process. For Λ CDM, our parameter set includes the six base parameters (Ω_m , S_8 , h_0 , Ω_b , m_{ν} , and n_s) and 10 nuisance parameters to account for the effects described in Section IIID. These nuisance parameters are four photo- z shift parameters Δz_i and four multiplicative shear biases m_i , corresponding to the four tomographic redshift bins, plus the two intrinsic alignment parameters A_1 and α_1 . For w CDM, we allow the dark energy equation of state parameter w_0 to vary, resulting in one additional free parameter. Our priors on the cosmological and nuisance parameters are the same as those of our simulated analysis and are shown in Paper I.

Our pipeline is implemented in COSMOSIS [52], which provides modular support for likelihood evaluation and sampling. We use its standard modules for modeling the two-point function and developed custom modules for computing the mass aperture three-point function, based

on the FASTNC framework introduced by Sugiyama *et al.* [20].

Posterior sampling is performed using the MULTINEST algorithm [53], integrated via the COSMOSIS interface. We perform our runs with `nlive` = 500, `efficiency` = 0.3, and `tolerance` = 0.1.

We measure our level of improvement on the Ω_m and S_8 constraints when adding $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ to ξ_{\pm} by comparing the figure-of-merit (FoM), which is defined by

$$\text{FoM} = \frac{1}{\sqrt{\text{Cov}(\Omega_m, S_8)}}, \quad (24)$$

where $\text{Cov}(\Omega_m, S_8)$ is the covariance of the posterior samples in the Ω_m and S_8 2D parameter space after marginalizing over all the other sampled parameters.

B. Blinding

It has been recognized that cosmological analyses are susceptible to the influence of biases from the one undertaking the analysis [55] [56]. The risk is that knowledge of results from other data sets can exert an influence over analysis choices. In order to avoid this risk, it is a common practice to perform data blinding while the pipeline is being tested. After consistency tests are performed on blinded data, the pipeline is then applied to the actual measured data.

We perform blinding at two different levels [57]. Firstly, we perform a catalog-level blinding. We follow Gatti *et al.* [27] and measure the two-point and three-point functions from a catalog with shifted ellipticities. Secondly, we also perform blinding at the level of the cosmological parameters. We hide the parameter values when displaying our posterior distributions and only view the relative values between our different data sets. We remove the blinding after ensuring that the χ^2 , the shift in cosmological parameter space, and the Ω_m - S_8 figure-of-merit improvement, all fall within their expected distribution ranges extracted from a set of 50 noisy realizations. These tests are detailed in Appendix B.

VI. RESULTS

A. Λ CDM Cosmological Constraints

We use our pipeline on the Y3 shear catalog and compare the constraints from the two-point function alone, with those including three-point data. For the nominal Λ CDM model, these results are presented in Figure 4. The improvement on the FoM on the Ω_m - S_8 plane is 111%. This value is slightly higher than the one forecast by our methodology paper. We verify the reasonable range of expected FoM improvements as part of our pre-unblinding tests, described in Appendix B.

Our cosmological constraints from the joint chain are

$$\xi_{\pm} + \langle \mathcal{M}_{\text{ap}}^3 \rangle: \begin{cases} \Omega_m &= 0.266^{+0.039}_{-0.040} \\ S_8 &= 0.780 \pm 0.015 \end{cases}. \quad (25)$$

We compare these results with those from our two-point analysis on Table I.

The consistency of the two-point and mass aperture data vectors can be verified by computing the shifts on parameter posteriors when using different parts of the data. Because $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ alone has little constraining power, we compare constraints from ξ_{\pm} directly with those from the joint analysis. Therefore, we compute the shift of the two-dimensional posterior peaks in the Ω_m - S_8 plane when moving from ξ_{\pm} to a joint analysis with $\langle \mathcal{M}_{\text{ap}}^3 \rangle$. We obtain a shift of around 0.5σ , which falls within the expected statistical scatter from noisy simulations, with a p-value of 0.78. (see Appendix B for blind tests).

We evaluate the goodness-of-fit of our joint analysis by computing the χ^2 at our maximum a-posteriori cosmology. We verify the χ^2 of the joint best fit to be of $\chi^2 = 87.3$. Our number of degrees of freedom equals $N_d - N_{\text{eff}}$, where N_d is our data vector length and N_{eff} is an effective value which is related to the number of free parameters and to constraints from our priors. Raveri and Hu [58] derive the expression for this value to be $N_{\text{eff}} = N_p - \text{tr}(C_{\Pi}^{-1}C_p)$, where C_{Π} and C_p are the parameter covariances of the prior and posterior distributions. We estimate $N_{\text{eff}} \approx 7$, which gives the number of degrees of freedom: $96 - 7 = 89$. Thus, our reduced χ^2 is 0.98, which indicates a good fit to the model.

A second goodness-of-fit test was also performed by running our pipeline on a set of 50 noisy realizations, generated by sampling around a theoretically computed data vector with our simulated covariance matrix. We compute the p-value, defined as the probability that the χ^2 from a random realization will exceed the value obtained with Y3 data. We find $p=0.14$. The distribution of χ^2 values is shown along with the Y3 value in Figure 5. The average χ^2 for a noisy mock simulation is 73.3. While this value is different than the number of degrees of freedom computed through the method by Raveri and Hu [58], both approaches show the statistical validity of our joint Y3 result. The differences between the results arise from the fact that the degrees-of-freedom approach assumes a Gaussian approximation for the posterior. The mock approach is free of this assumption, but its precision can be affected by the number of realizations.

We also visually confirm that our best fit matches well our ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ data points at all redshifts and scales, without any systematic departure from the fit. We show this result in Figures 1 (for $\langle \mathcal{M}_{\text{ap}}^3 \rangle$) and 14 (for ξ_{\pm}).

We also use our best fit cosmology to predict the ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ at scales excluded from the fit. This prediction is made with a dark-matter-only theoretical model, in order to verify if significant departures from the data would be identifiable at small scales. However, we still find adequate concordance both in $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ (Figure 2) and

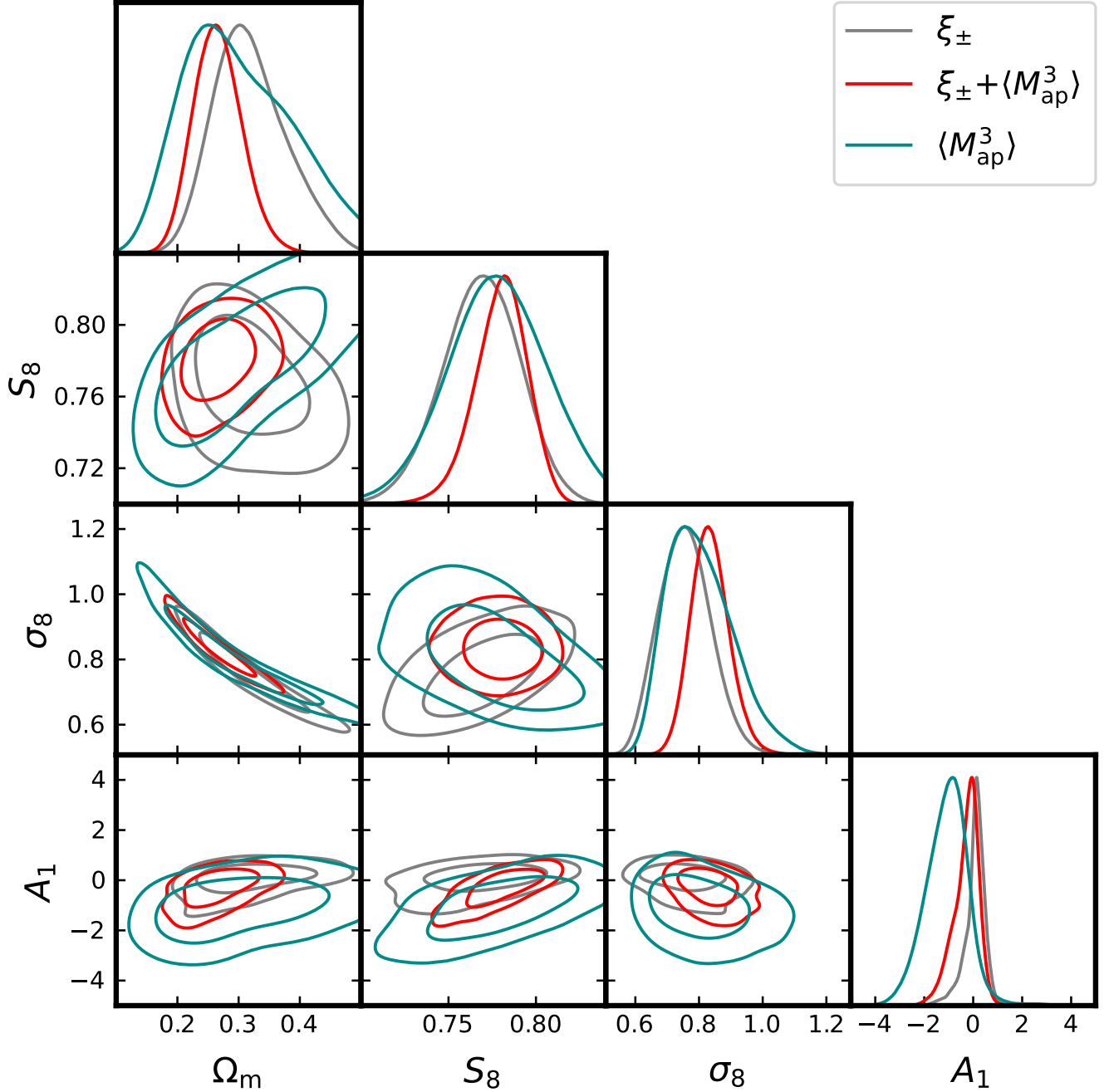


FIG. 4. Parameter constraints from DES-Y3 data from separate and joint analyses of ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ within Λ CDM. The cosmological parameters shift less than 1σ when moving from ξ_{\pm} to a joint 2PCF and 3PCF analysis. We find an improvement of 111% in the Ω_m - S_8 figure-of-merit. This improvement is driven in part by the difference between the degeneracy directions for ξ_{\pm} alone and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ alone. Here A_1 is the intrinsic alignment amplitude (see Eq. 17)

ξ_{\pm} (Appendix D). This serves as an independent validation of the robustness of the joint $\xi_{\pm} + \langle \mathcal{M}_{\text{ap}}^3 \rangle$ analysis against baryonic effects. In our previous methodology paper, we show that this robustness is achieved by selecting scales on $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ in which the net effect of baryons is one of a slight enhancement of the signal. The combination of these scales with the known suppression on the 2PCF

makes the joint analysis more robust against shifts on S_8 .

While in our methodology paper we had forecast a slight improvement on the intrinsic alignment amplitude parameter, we do not see it in Y3 data. The question remains whether the future development and validation of more complex IA modeling for three-point statistics will allow us to find a pattern of improvement when moving

TABLE I. Mean of parameter posteriors from 2PCF and 2PCF+ $\langle\mathcal{M}_{\text{ap}}^3\rangle$ analysis.

Parameter	ξ_{\pm}	$\xi_{\pm} + \langle\mathcal{M}_{\text{ap}}^3\rangle$	$\xi_{\pm} (w\text{CDM})$	$\xi_{\pm} + \langle\mathcal{M}_{\text{ap}}^3\rangle (w\text{CDM})$
Ω_m	$0.318^{+0.063}_{-0.059}$	$0.266^{+0.039}_{-0.040}$	$0.298^{+0.061}_{-0.059}$	$0.242^{+0.038}_{-0.037}$
S_8	0.769 ± 0.022	0.780 ± 0.015	0.743 ± 0.025	$0.749^{+0.027}_{-0.026}$
σ_8	$0.757^{+0.080}_{-0.084}$	$0.835^{+0.058}_{-0.060}$	0.756 ± 0.079	0.841 ± 0.057
w_0	-	-	$-1.42^{+0.32}_{-0.33}$	-1.39 ± 0.31

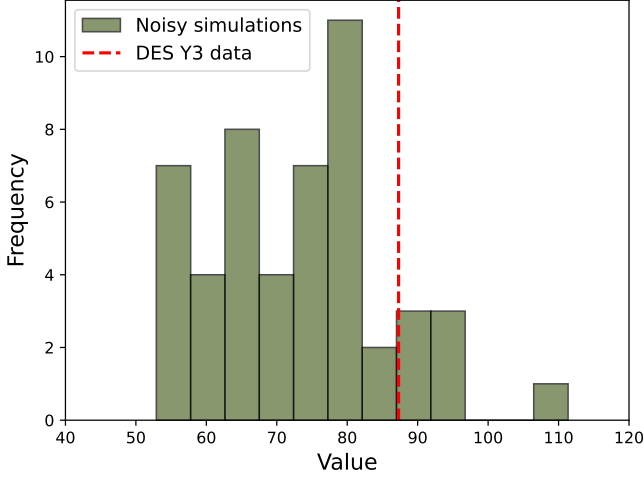


FIG. 5. The χ^2 distribution estimated from noisy mock simulations for the DES joint $\xi_{\pm} + \langle\mathcal{M}_{\text{ap}}^3\rangle$ analysis. The value from the analysis with Y3 data is represented by the dashed red line. Our data χ^2 is consistent with its expected distribution, with a probability of 14% of a random realization having higher χ^2 . Our distribution average is 73.3, which is sufficiently close to our effective number of degrees of freedom ($N = 89$).

from ξ_{\pm} to $\xi_{\pm} + \langle\mathcal{M}_{\text{ap}}^3\rangle$.

B. $w\text{CDM}$ Cosmological Constraints

We also perform a run of our pipeline with $w\text{CDM}$ modeling, and find an improvement of 66% on the Ω_m - S_8 FoM. Our improvement in the joint S_8 - w_0 constraint is of 22%. This is higher than the 5% improvement we see on the marginalized w_0 parameter. This difference is compatible with what we identified in our methodology paper, that the inclusion of mass aperture allows a degeneracy breaking in the S_8 - w_0 plane. We show our results in Figure 6, with the mean posteriors listed on Table I. When including $\langle\mathcal{M}_{\text{ap}}^3\rangle$ in the $w\text{CDM}$ analysis, our w_0 constraint remains compatible with $w_0 = -1$ at approximately 1σ .

While our analysis finds a better improvement on Ω_m - S_8 in ΛCDM rather than $w\text{CDM}$, this is not a pattern across all Y3 HOS analyses. The simulation-based in-

ference (SBI) by Gatti *et al.* [54] reports a gain of 70% with ΛCDM and 90% with $w\text{CDM}$ when adding third moments, scattering transforms and wavelet phase harmonics to the second moment of the convergence field. Another SBI result by Jeffrey *et al.* [17] combines power spectra, peak counts, and a direct map-level inference, finding for $w\text{CDM}$ improvements of 126% on the Ω_m - S_8 FoM and 148% on the Ω_m - w_0 FoM. A main difference between these analyses and the present work is that both of them require $-1 < w_0 < -1/3$, while we use wider priors and allow for $w_0 < -1$. Since most of our posterior distribution is situated on the $w_0 < -1$ regime, we are essentially probing a distinct part of the parameter space.

In order to directly compare the DES Y3 $\xi_{\pm} + \langle\mathcal{M}_{\text{ap}}^3\rangle$ $w\text{CDM}$ cosmological constraints with those from other higher-order statistics that require simulation-based inference, we perform an additional run of our pipeline with adapted priors to match those of Gatti *et al.* [54], Jeffrey *et al.* [17], and Prat *et al.* [18]. In this analysis, we restrict the dark energy equation of state parameter to $-1 < w_0 < -1/3$, and introduce Gaussian priors on h_0 , n_s , and $\Omega_b h_0^2$, as described in Table I of Jeffrey *et al.* [17]. We present, in Figure 7, the DES Y3 $\xi_{\pm} + \langle\mathcal{M}_{\text{ap}}^3\rangle$ constraints on Ω_m , S_8 , and w_0 , together with those from third moments, scattering transforms, and wavelet phase harmonics [54], persistent homology (Betti numbers) [18], and power spectrum, peaks, and map-level inference with CNNs [17]. We also include as a baseline the ξ_{\pm} result from Prat *et al.* [18], which adopts the same priors as those from the SBI analyses. All the constraints from these HOSs are consistent with each other and represent, individually, significant improvements relative to ξ_{\pm} alone.

C. Comparison with external CMB data

We compare our results with CMB constraints from the Planck Collaboration in order to investigate whether they affect the current tension between early and late universe probes of S_8 . Our results are shown in Figure 8. We quantify the tension in the S_8 - Ω_m plane through the method devised by Raveri and Hu [58]. We obtain, for the tension between DES Y3 ξ_{\pm} and Planck 2018, a significance of 2.3σ . The tension between DES Y3 $\langle\mathcal{M}_{\text{ap}}^3\rangle$

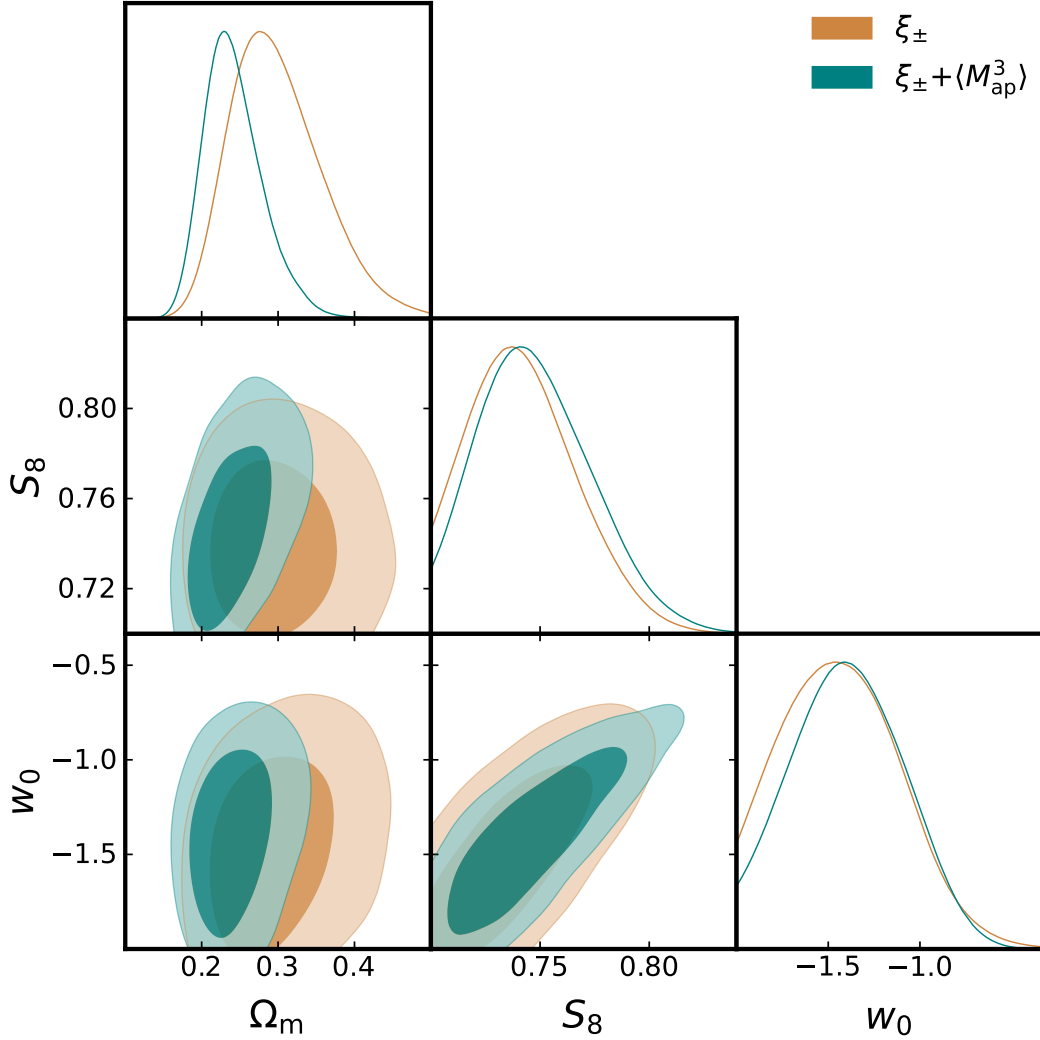


FIG. 6. Parameter constraints from DES Y3 data using ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$, as in Figure 4 but for the w CDM model. We find an improvement of 66% in the Ω_m - S_8 figure-of-merit. Our improvement on the w_0 parameter is only 5%, while the improvement on the joint w_0 - S_8 constraint is 22%.

alone and Planck is of 1.9σ . Finally, the tension between our joint constraint and the one from Planck is of 2.3σ , remaining the same as when we use the two-point only data set.

D. Uninformative priors on photo-z shift parameter

Our main cosmological analysis follows [29] and [2] in using informative priors on the photo-z shift parameters. Here we also run our pipeline removing the informative prior on Δz_4 and replacing it with a flat prior ranging from -0.3 to 0.3. This is motivated by the observation that the density of galaxies used for photo-z calibration of the fourth z-bin is small, which can lead to larger potential biases. We find that by using an uninformative

prior on Δz_4 , our pipeline finds its value to be significantly higher than that of the fiducial Y3 analysis, being 2σ above zero (Figure 9). Due to the degeneracy of the photo-z shifts with S_8 , we find a preference for a lower S_8 relative to the joint fiducial result. We obtain $S_8 = 0.770 \pm 0.015$. While the ξ_{\pm} result alone is degraded by removing the informative priors, the contours do not get larger for the joint analysis, which yields us a FoM improvement of 144% on the $\Omega_m - S_8$ plane. As a result of the shift in S_8 , the tension between the joint chain with uninformative Δz_4 priors and CMB results from Planck 2018 is slightly larger, with a significance of 2.6σ .

We also perform the same analysis procedure on a simulated data vector and obtain for the mean photo-z shifts results consistent (at 0.4σ) with their true values. Therefore, the observed shift on Y3 data should not be dismissed as resulting from projection effects alone.

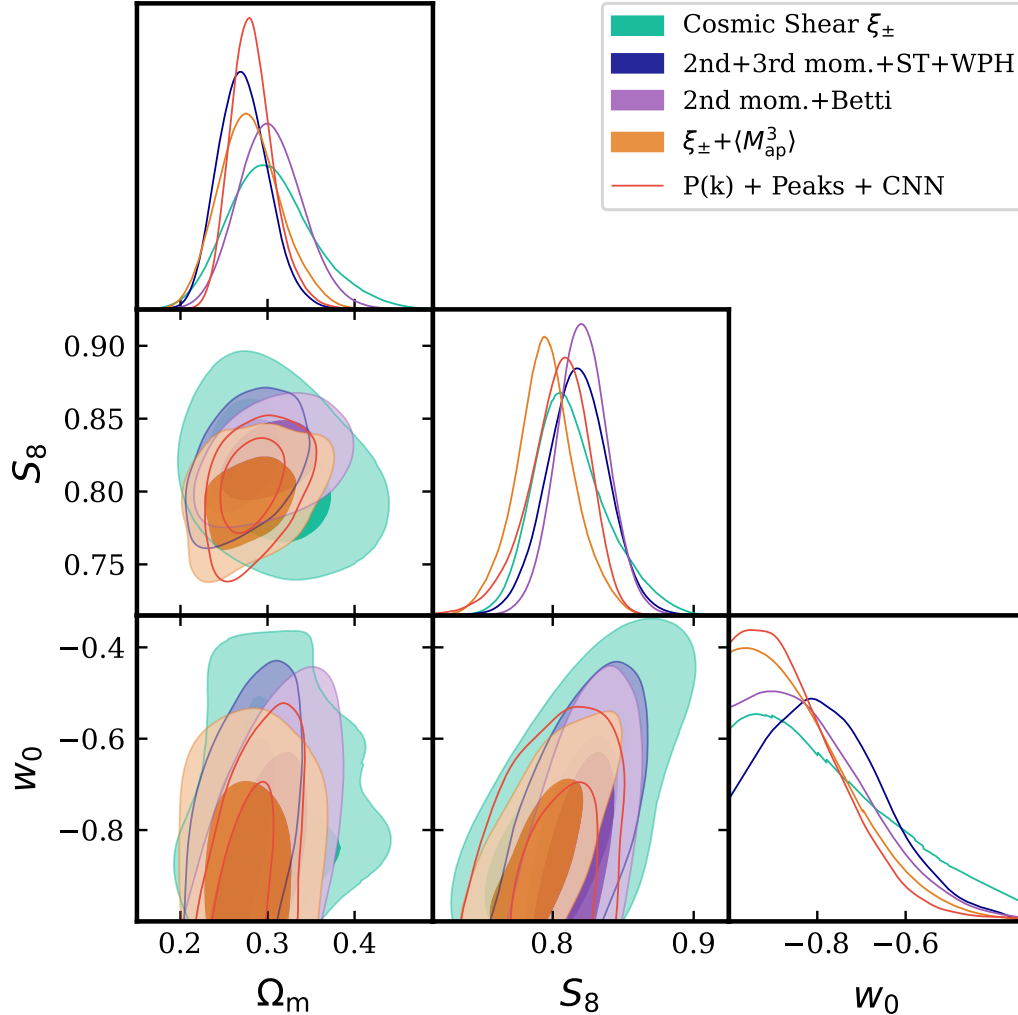


FIG. 7. Parameter constraints from DES Y3 data using different higher-order statistics. We compare constraints by Gatti *et al.* [54] from the second and third moment of the weak lensing mass map, scattering transforms, and wavelet phase harmonics, constraints by Prat *et al.* [18] from the second moment and Betti numbers, constraints by Jeffrey *et al.* [17] from power spectra, peak statistics, and CNN map-level inference, and the constraints from ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ presented in this work. We also include, as a baseline, the ξ_{\pm} result by Prat *et al.* [18]. We adapt the priors on our analysis to match those of the SBI-based statistics. All higher-order-statistics yield significant improvements relative to ξ_{\pm} alone and are consistent with each other.

VII. CONCLUSION

We measure the full configuration-dependent three-point correlation function on the DES Y3 shape catalog. We perform the first set of tomographic three-point measurements over all 20 combinations of the four fiducial source galaxy redshift bins. We compress these measurements into the mass aperture skewness statistic with essentially no loss of information. We also perform new measurements of the two-point correlation function on the shape catalog, correcting for an issue on the data vector of the DES fiducial analysis. We extend our measurements to small scales (down to $0.25'$) for validation purposes (see Appendix D for a test of baryonic feedback

effects).

We perform a Λ CDM cosmological analysis with two-point and three-point correlation functions measured on the DES Y3 shape catalog, through a joint $\xi_{\pm} - \langle \mathcal{M}_{\text{ap}}^3 \rangle$ data vector. We find $S_8 = 0.780 \pm 0.015$ and $\Omega_m = 0.266^{+0.039}_{-0.040}$. We obtain an improvement of 111% on the figure of merit on the Ω_m - S_8 plane. This level of improvement is driven by the slightly different degeneracy directions of the individual 2PCF and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ contours, as can be seen in Figure 4. This significant increase in cosmological information is somewhat surprising in view of other higher order statistics applied to data ([25]), but is consistent with that found by the KiDS collaboration [26] and by our simulated analysis. Much

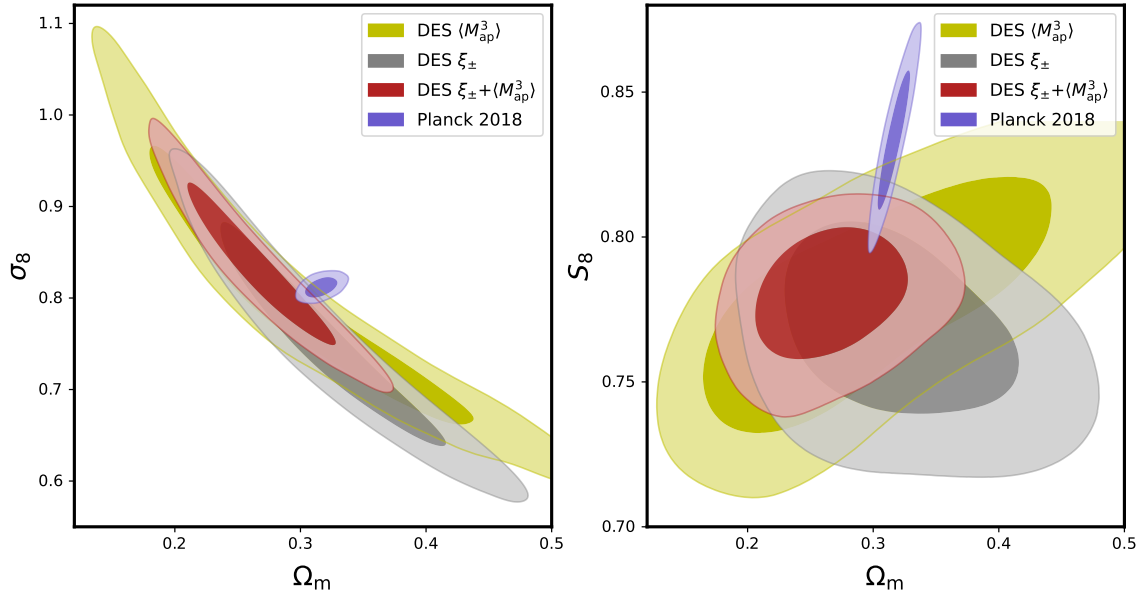


FIG. 8. Tension between DES Y3 results from ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ and CMB results from Planck 2018. We use their TT-TEEE+lowl/lowE results, which include temperature (TT), polarization E-modes (EE), the cross spectra (TE), and the low multipole values for both temperature and polarization. The $S_8 - \Omega_m$ tension between the DES Y3 shear two-point analysis and Planck data is of 2.3σ . While the tension is smaller for $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ (1.9σ) due to it being less informative, the tension between the joint $\xi_{\pm} - \langle \mathcal{M}_{\text{ap}}^3 \rangle$ DES Y3 constraint and Planck remains at 2.3σ .

of the gain comes from using the full configuration dependence of the three-point function. We also test the w CDM model and find a more modest 22% improvement in the joint $S_8 - w_0$ figure of merit. Our constraints from w CDM are $\Omega_m = 0.242^{+0.038}_{-0.037}$, $S_8 = 0.749^{+0.027}_{-0.026}$ and $w_0 = -1.39 \pm 0.31$.

We compare our results with primary CMB data and find that the tension between previous LSS results from DES and CMB results from Planck 2018 persists with addition of the new data set. We find a tension of 2.3σ between our joint $\xi_{\pm} + \langle \mathcal{M}_{\text{ap}}^3 \rangle$ results and Planck 2018. While our constraints were significantly tightened in parameter space, the persistence of the tension motivates further work towards DES Year 6 data and methods to use all the available information in the shear field.

We have found that the addition of $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ makes the analysis more robust to baryonic feedback [See Fig. 10 of 24]. This finding needs further study with a wider set of hydrodynamic simulations and variation in scale cuts. It opens the possibility of designing a joint analysis that improves constraints on both cosmological and baryonic feedback parameters.

We explored variations in the treatment of uncertainty due to photo- z estimation. Section VID shows the results of freeing up the prior on the mean redshift of the fourth bin, which has limited coverage in the calibrating spectroscopic redshifts. We find that the joint analysis can constrain the mean, with the posterior showing a shift

of about 0.05 towards higher redshift. This results in a slightly lower S_8 and increases the tension with Planck to 2.6σ . We also tested the effect of freeing up both the third and fourth redshift bin, with similar conclusions.

The assessment of the tension with Planck can also potentially be affected by the choice of non-linear power spectrum prescription. Anbajagane *et al.* [59] report a small shift towards higher S_8 (and lower tension) when using HMCode [60] instead of Halofit. A similar shift was noticed by Seco *et al.* [29], who attributed it to a projection effect of their free baryon parameters. We leave this topic to be explored in further study.

The methodology for marginalizing over systematics can be improved in future work. In particular, for modeling IA, we have restricted this study to the NLA model. While there are suggestions from previous DES, KiDS and HSC studies that the amplitude of this signal is modest (or consistent with zero), the TATT model is a more realistic model and also complex – it would lead to weaker cosmological constraints. Simulation-based models of IA are also emerging and would provide a useful cross-check [61]. At the level of precision of DES Y3, other systematics such as reduced shear or magnification are expected to be negligible, but these will need a more careful examination for Stage 4 surveys.

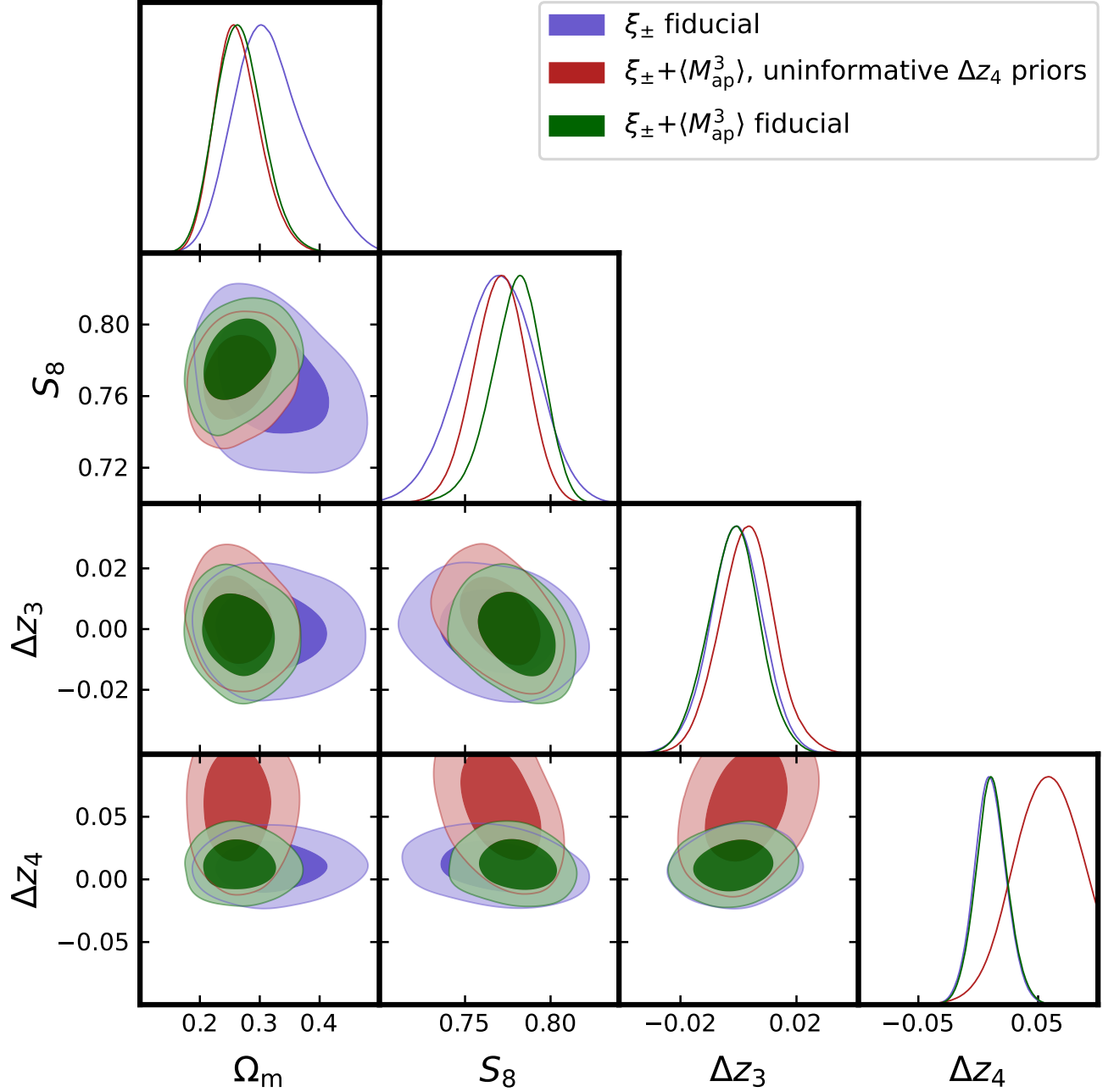


FIG. 9. Parameter constraints from ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ analysis using uninformative flat priors on the photo-z shift parameter Δz_4 . We find that by removing the informative prior, the joint ξ_{\pm} - $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ analysis calibrates Δz_4 to a value higher than the one assumed in the fiducial DES analyses, leading to a shift towards lower S_8 . We also include the third photo-z shift parameter Δz_3 to indicate the slight degeneracy between it and Δz_4 .

VIII. AUTHOR CONTRIBUTIONS

RCHG and SS developed the inference pipeline, performed the cosmological analyses, and composed the manuscript. BJ served as project advisor and contributed to the manuscript. MJ developed novel TREECORR functionalities that were used in this work.

DA generated the CosmogridV1 mocks used in this work. AH, GAM and SP served as collaboration internal reviewers. JM was the final publication reader. MG, ES, AAm, MB, MT, ACh, CDo, NM, ANA, IH, DG, GB, MJ, LFS, AF, TS, JMc, RPR, RCh, CC, SP, IT, JP, JEP, CS are credited for the development of the DES shape catalog; JMy, AAl, AAm, CS, SE, JD, JMc, DG,

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Appendix A: Comparison with other DES two-point analyses

For this work, we performed our own measurement and analysis of the two-point correlation function on DES Y3 data. Here we show the differences in the analysis choices between our pipeline and the fiducial Y3 pipeline.

Firstly, McCullough *et al.* [62] identified a mismatch between the tomographic binning file used in the Y3 fiducial analysis and an outdated version which was present on the shear catalog file and used to produce the two-point data vector. Following their procedure, we created a new two-point data vector making use of the same shear catalog but updating the tomographic binning to match the redshift distributions assumed by the pipeline (SOMPZ version 0.5).

Next, our analysis also differs from the Y3 fiducial contours in that we use NLA instead of TATT for intrinsic alignment modeling, in order to have a consistent modeling between ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$. This is what drives most of the difference between the contours. We do not use any additional shear ratio data vector, and we adapt our parameter space sampling to match with our emulator support range as described in Paper I. Finally, we make use of the Percival likelihood to better propagate the uncertainties of our simulated covariance into the final constraints (Eq 21). We show our contours alongside the DES Y3 fiducial ones on Figure 10, obtaining a small shift of 0.2σ on the mean S_8 value.

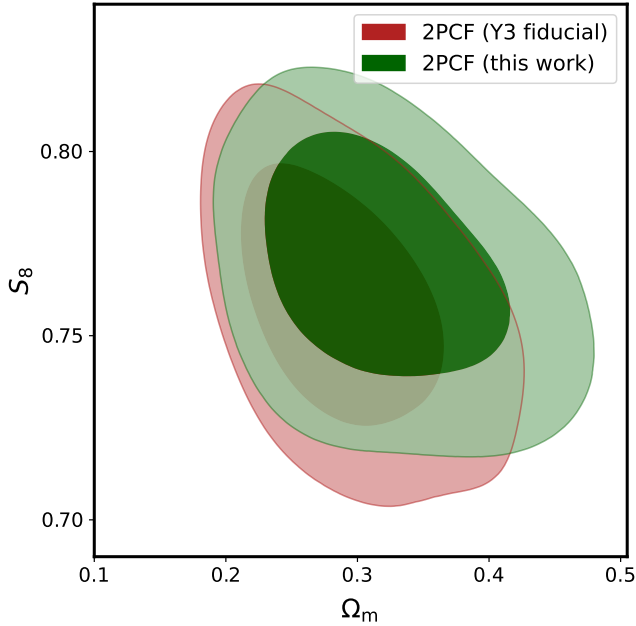


FIG. 10. Results from our two-point shear analysis along with the fiducial Y3 cosmic shear results. The differences between our chains include our correction of the tomographic binning as well as methodological differences such as the use of NLA for intrinsic alignments. The shift in S_8 between both analyses is of 0.2σ .

Appendix B: Blind analysis

Previous to running our chains on the actual Y3 data vector, we performed a blind analysis, by both making use of a shifted catalog, such as was done in previous Y3 analyses, and by hiding the parameter values on our plots and showing only shifts relative to the ξ_{\pm} mean posterior values. The only methodological change made after unblinding was the shift from a Gaussian likelihood to the one proposed by Percival *et al.* [48]. While this affects the shape of the posterior tail, our blind tests remain valid due to the use of posterior peaks and confidence intervals.

We generated 50 noisy realizations at CosmoGridV1 cosmology, using our measured covariance matrix to generate Gaussian noise. We ran our pipeline on each of the simulations in order to verify the distributions of χ^2 , of Ω_m - S_8 figure-of-merit, and of the shift in cosmological parameters when moving from a two-point only analysis to a joint analysis. Our χ^2 distribution is the same shown in Figure 5. The blind χ^2 value is 69.7, which gives a p-value of 0.58.

We also compute, for each one of our noisy simulations, the 2D parameter space distance between the peak of the Ω_m - S_8 posteriors from the two-point only and from the joint analysis. We show the distribution along with the Y3 inferred value in Figure 11. The shift found in data also lies within the expected distribution, with a p-value

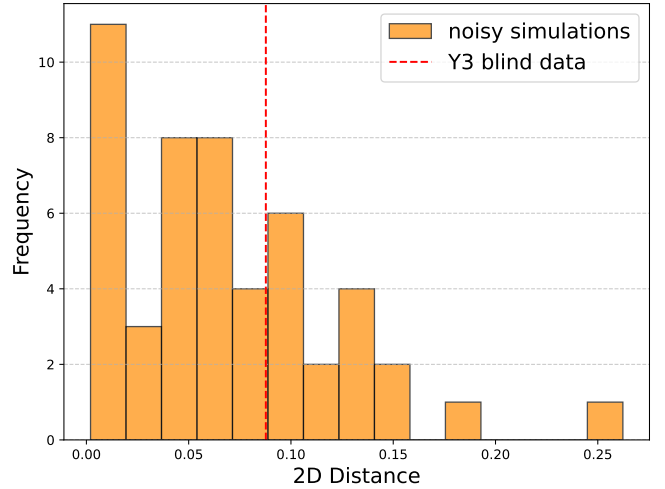


FIG. 11. Two-dimensional distance between the Ω_m - S_8 posterior means when shifting from two-point only to a joint analysis. The histogram shows the results for 50 different noise realizations, while the red dashed line indicates the shift obtained in blind Y3 data. We verify that the Y3 data shift falls within the expected distribution, with a p-value of 0.78.

of 0.32.

Finally, we verify the figure of merit on the Ω_m - S_8 plane for each of the noisy simulations in both the ξ_{\pm} scenario and the $\xi_{\pm} + \langle \mathcal{M}_{\text{ap}}^3 \rangle$. The probabilities of exceeding are 88% for ξ_{\pm} and 40% for the joint chain. When we analyze the distribution of the figure-of-merit ratios, we find an improvement range of between 24% to 170%. This indicates that the factor-of-two increase in constraining power when adding $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ to ξ_{\pm} can be significantly suppressed or boosted depending on the noise realization.

Appendix C: Full set of contours for the cosmological parameters

Here we present our full set of contours obtained with both Λ CDM and w CDM modeling. We do not include constraints on the sum of the neutrino masses because our $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ emulator assumes a fixed value of $\Sigma m_{\nu} = 0.06$. For Λ CDM, our results are shown in Figure 12. For w CDM, we present our results in Figure 13.

Appendix D: Small scale measurements of the 2PCF

Our cosmological analysis makes use of the same 2PCF measurements as done in the fiducial Y3 analyses, spanning from $2.5'$ to $250'$ in 20 log-spaced bins. Some of the smallest scales are removed to avoid biasing due to baryonic suppression.

We perform a new set of measurements at even smaller scales, not to be used in our fit, but to verify whether

a significant deviation from dark-matter-only theoretical predictions would be apparent from Y3 shear data alone. We introduce 10 new log-spaced bins between $0.25'$ and $2.5'$, and compute their error bars via jackknife sampling of the data. The fiducial Y3 measurements and our new small scale measurements are shown in Figures 14 and 15 along with a Halofit prediction using our joint best fit parameters. We find the joint cosmology to be still a

reasonable fit to the small scale data.

A similar analysis by Pandey *et al.* [63] found the Y3 2PCF small scale data points to be significantly below the Halofit theoretical prediction. This result, however, differs from ours in that their best fit parameters are obtained from a joint DES Y3 and ACT DR6 tSZ analysis with baryonic free parameters, which yields an S_8 value much higher than ours.

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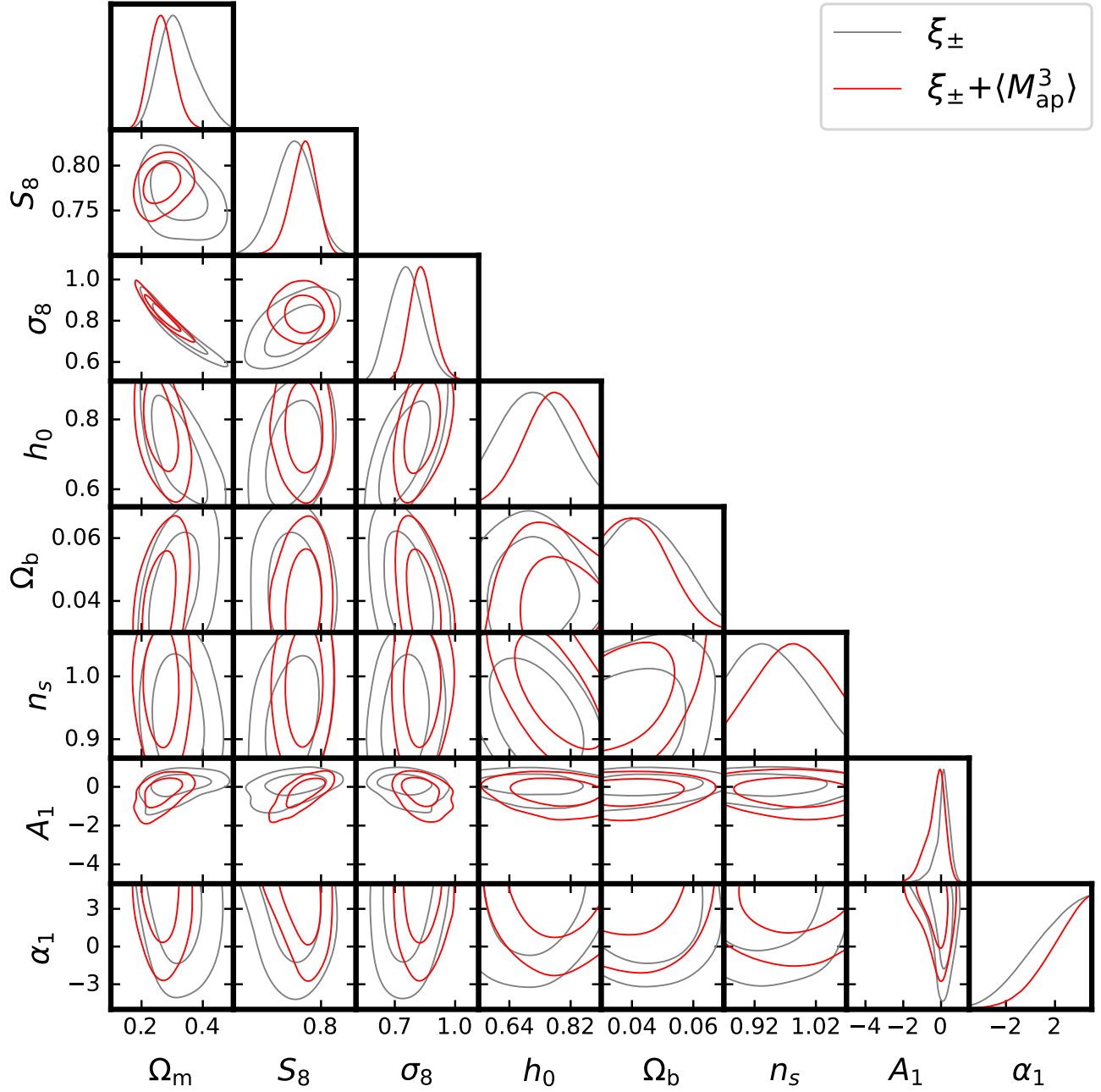


FIG. 12. Parameter constraints on the full set of cosmological parameters from ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ using Λ CDM modeling. The dark blue contours indicate results from the 2PCF alone, while the red contours indicate results from the 2PCF combined with $\langle \mathcal{M}_{\text{ap}}^3 \rangle$.

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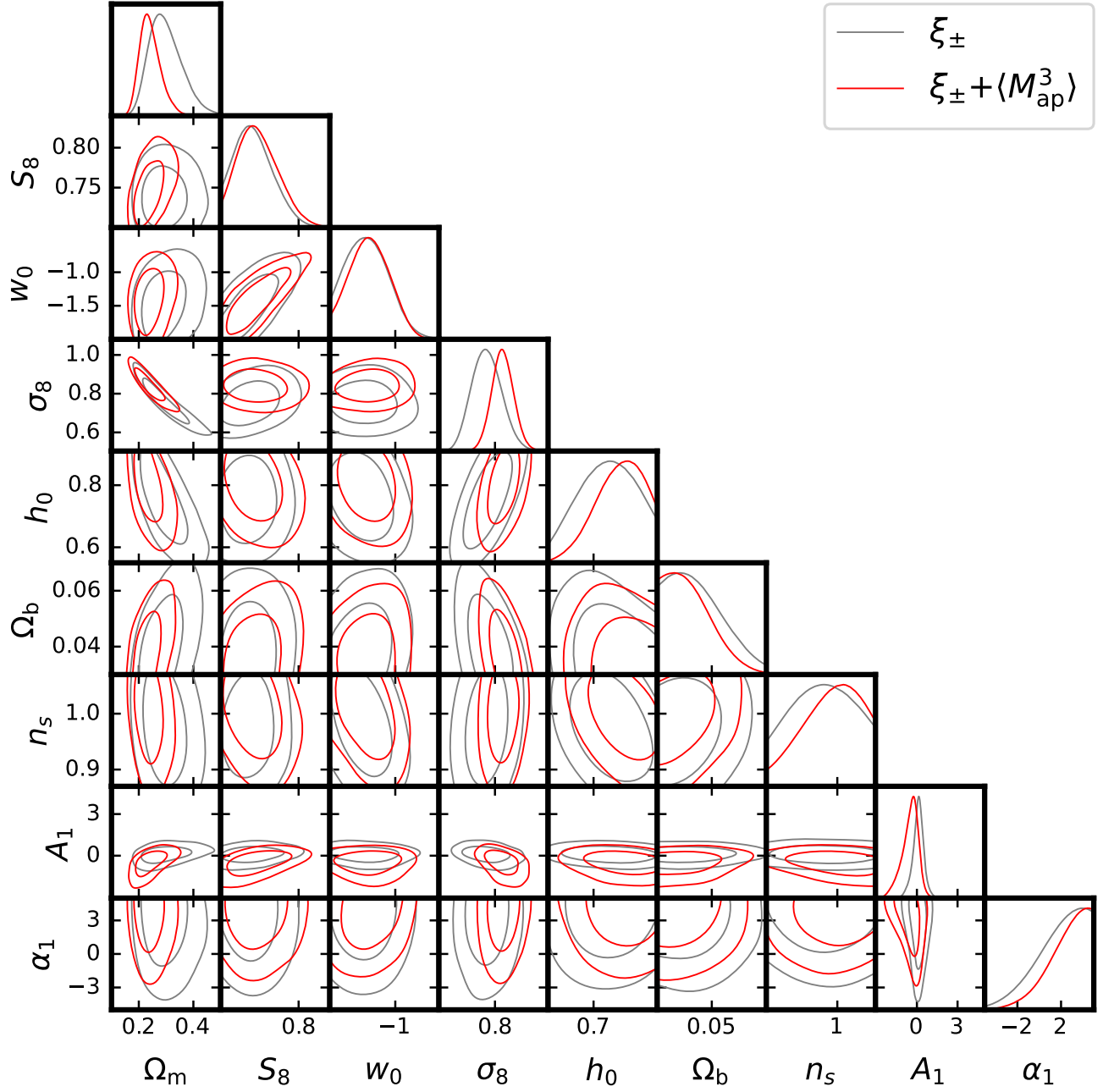


FIG. 13. Parameter constraints on the whole set of cosmological parameters from ξ_{\pm} and $\langle \mathcal{M}_{\text{ap}}^3 \rangle$ using w CDM modeling. The dark blue contours indicate results from the 2PCF alone, while the red contours indicate results from the 2PCF combined with the mass aperture statistic.

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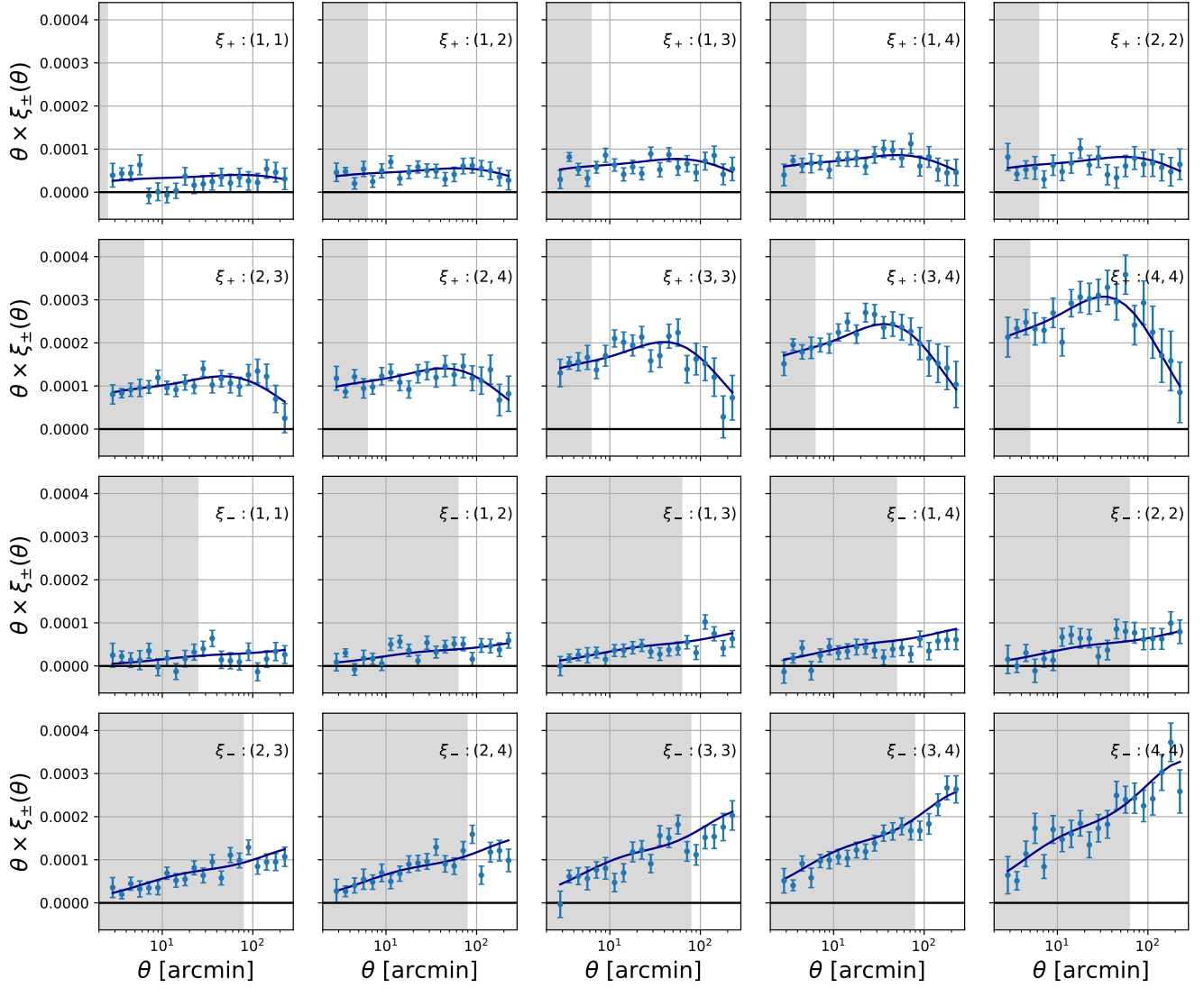


FIG. 14. 2PCF Y3 data along with the theoretical prediction at our joint best fit cosmology. The gray region corresponds to data points which were not used in the fit due to scale cuts. We verify that the joint dark matter-only theory model reasonably fits the excluded data points. A more detailed analysis is required in order to identify possible deviations in the data due to baryonic feedback and to quantify its level of agreement with the dark matter-only model.

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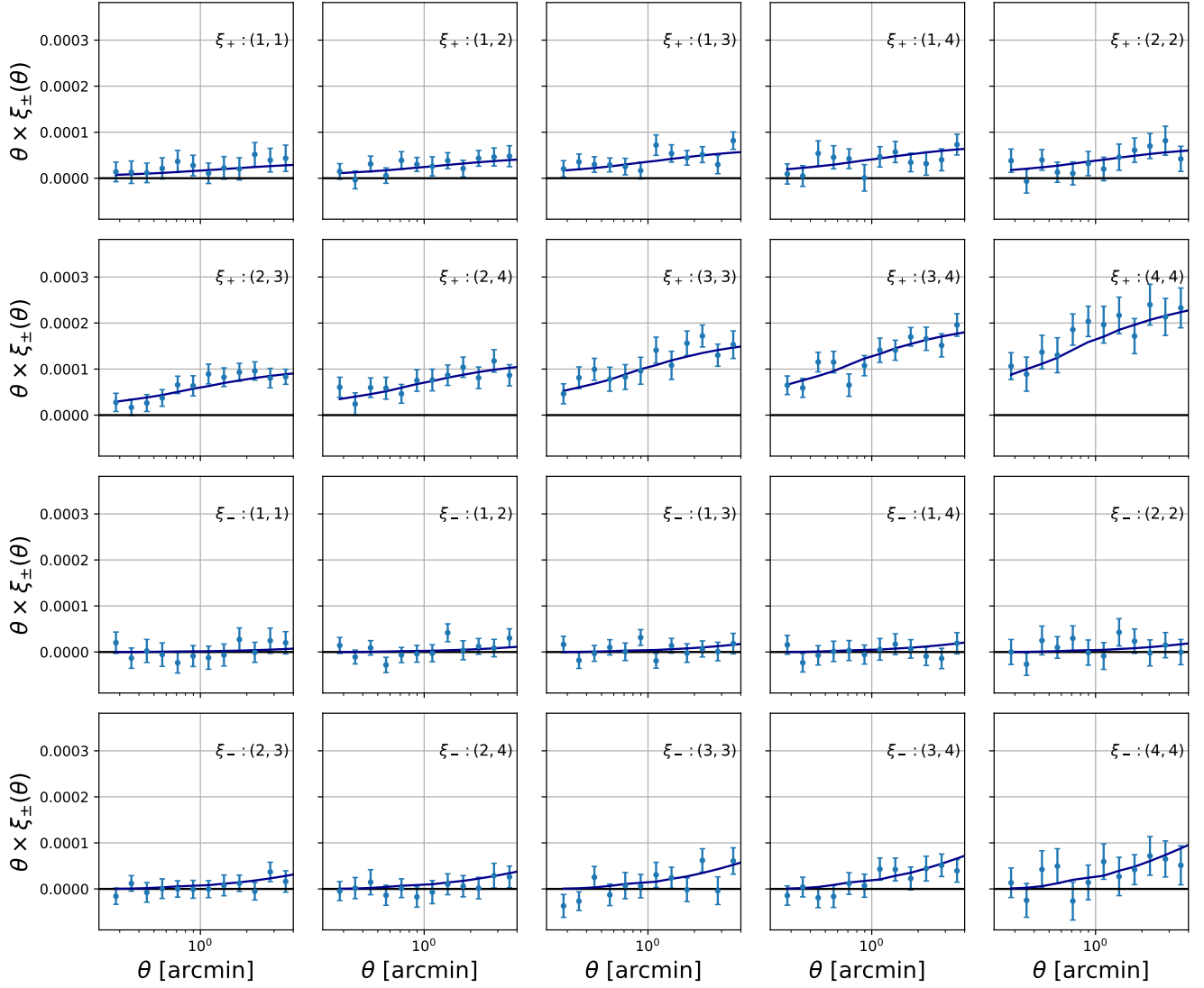


FIG. 15. 2PCF Y3 data at scales around 1 arcmin along with the theoretical dark matter-only prediction at joint best fit cosmology (computed with fiducial Y3 scale cuts). At our level of noise, the baryonic suppression of the signal is not clearly identifiable.

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